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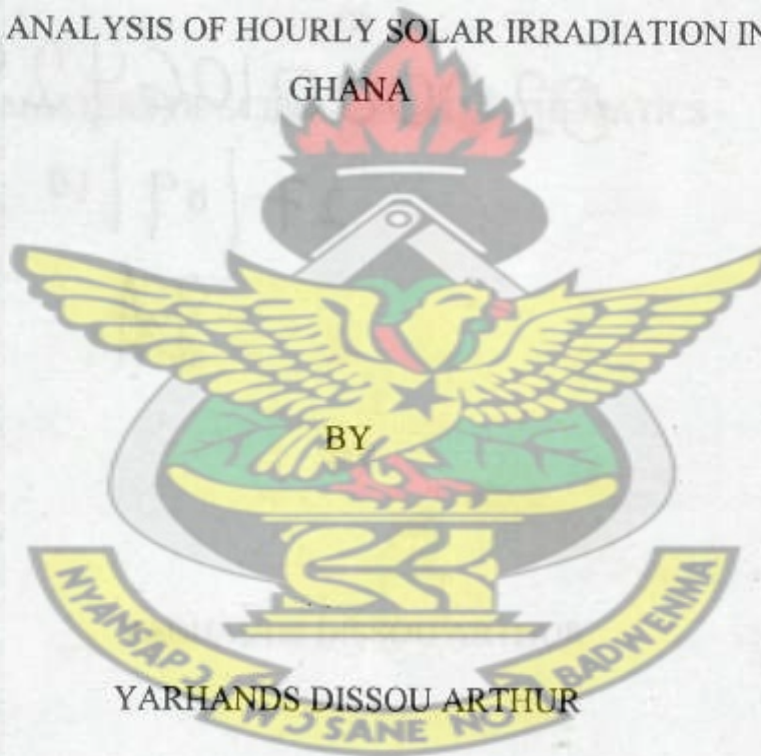
COLLEGE OF SCIENCE

FACULTY OF PHYSICAL SCIENCES

DEPARTMENT OF MATHEMATICS

KNUST

INFERENCEAL ANALYSIS OF HOURLY SOLAR IRRADIATION IN KUMASI,  
GHANA



YARHANDS DISSOU ARTHUR

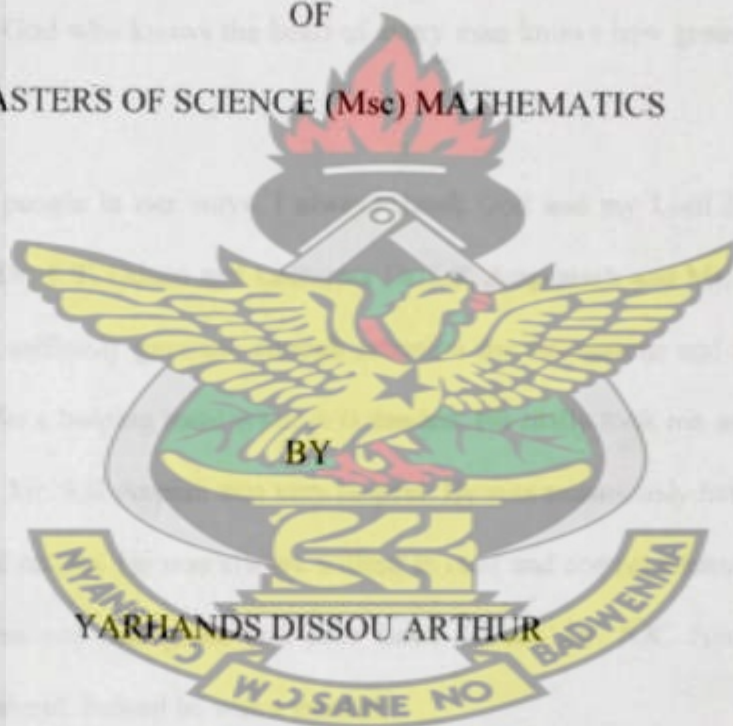
JULY 2009

**INFERENCEAL ANALYSIS OF HOURLY SOLAR IRRADIATION IN KUMASI,  
GHANA**

**A THESIS PRESENTED TO THE DEPARTMENT OF MATHEMATICS, KWAME  
NKRUMAH UNIVERSITY OF SCIENCE AND TECHNOLOGY, IN PARTIAL  
FULLFILLMENT OF THE REQUIREMENT FOR THE AWARD**

**OF**

**MASTERS OF SCIENCE (Msc) MATHEMATICS**



**BY**

**YARHANDS DISSOU ARTHUR**

**JULY 2009**

## ACKNOWLEDGEMENT

*But thanks be to God who gives us victory through our Lord Jesus Christ. Therefore, my brethren, be steadfast, immovable, always abounding in the work of the Lord, knowing that your labor is not in vain in the Lord. (1<sup>st</sup> Corinthians 15:57-58)*

I praise my God true His son Jesus Christ for his infinite blessings and mercies towards my life. May His name be glorified.

I want to thank my parents Mr. Charles Kwaku-vi and Miss Esther Botchwey for their parental support during my school period. I also want to thank Mr. and Mrs. Odei-Asiedu for their financial, spiritual and their moral support given to me during my period of study. Indeed words cannot express my gratitude but God who knows the heart of every man knows how grateful I am to Mr. and Mrs. Odei-Asiedu.

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And to you my special friends: Jemima Owusua Awenseba, Alexander Kwame Kwarting, Ebenerzer Kwaku Wilmot, Thywil Dzoveho and all friends of Yarhands, I say a big thank you all and God bless you where ever you are.

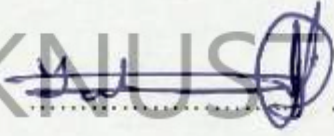
To God be the glory.



## DECLARATION

I hereby declare that this submission is my own work toward the award of the M.Sc. Mathematics, and that to the best of my knowledge, it contains no material previously published by another person nor material which had been accepted for the award of any other degree of the university, except where due acknowledgement had been made in the text.

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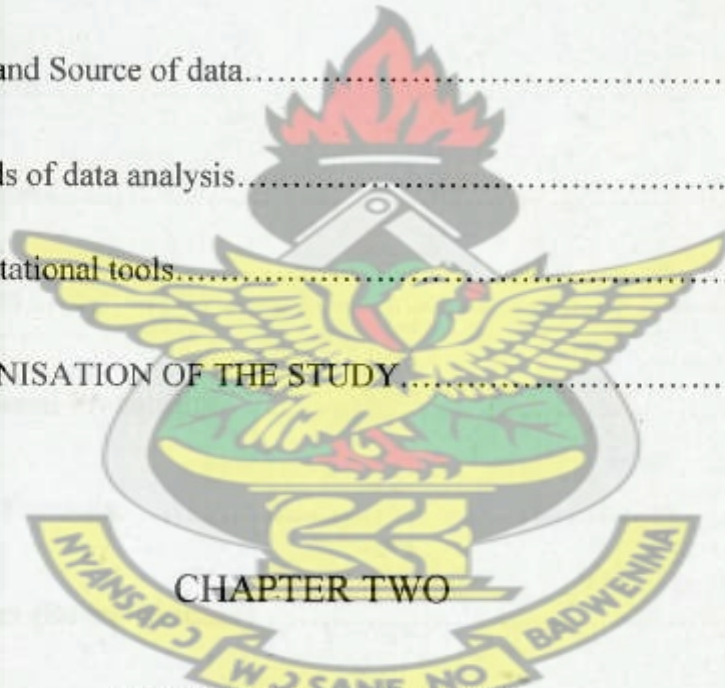
## Dedication

This thesis is dedicated to Mr. and Mrs. Odei-Asiedu and my mother.

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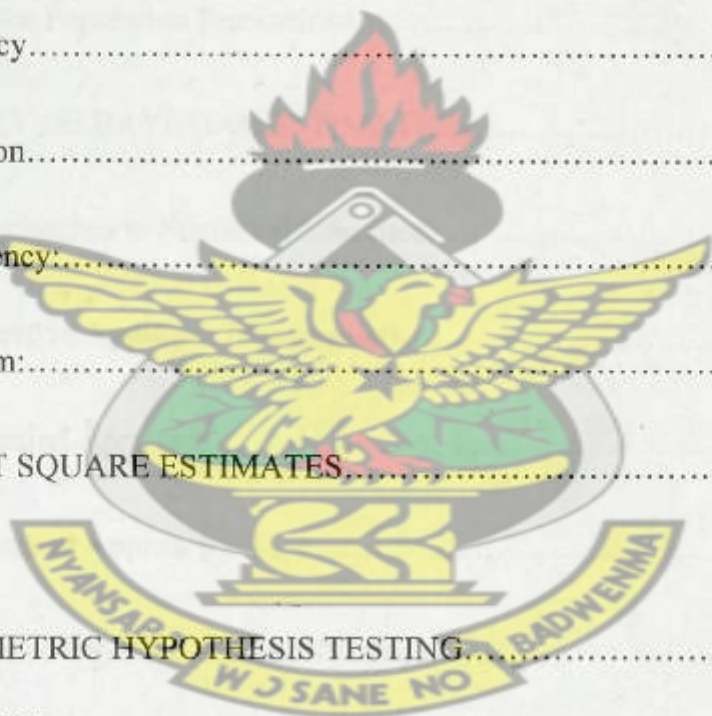
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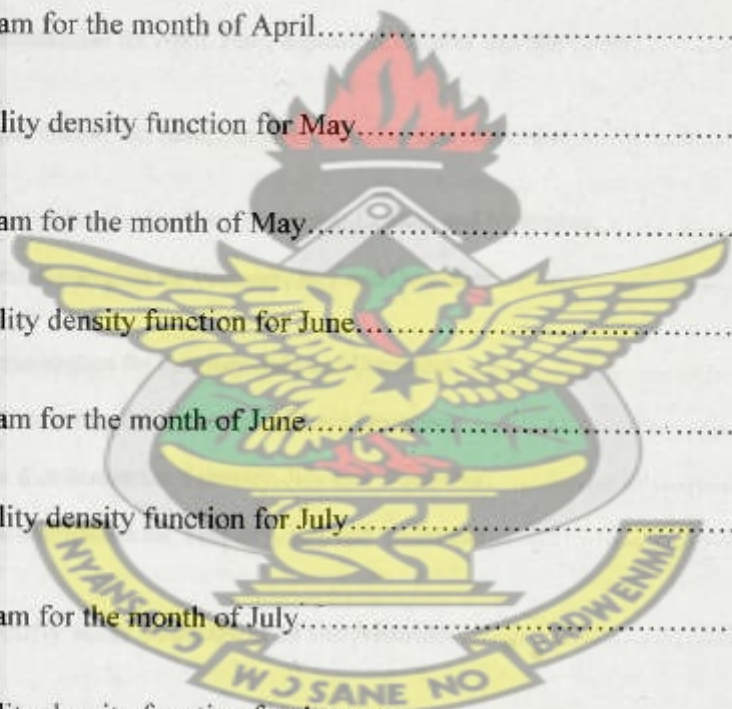


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## ABSTRACT

The inferential analysis of hourly solar irradiation for Kumasi –Ghana was conducted using 14 years of data of measured values by the KNUST SOLAR LABORATORY. The analysis was carried out to find out the probability distributions that best fit the data of a given month of the year. Further analysis was carried out to find out the cluster for the month of the year. From the analysis conducted, it was found that solar irradiation for January, March and May can be fitted with the lognormal probability distribution. The month of April can be fitted with the Exponential, Weibull, lognormal Geometric and Gamma distribution while the months of June to December can be fitted with Exponential, Weibull, Geometric and Gamma. From the hypothesis testing carried out it was discovered that the months of the year can be put into two clusters. The first cluster is from the month of January to October and the second cluster is the November and December. Further analysis was conducted to find out if there are significant difference between the two clusters of the year and it was found that the result of the test showed significant different between the clusters.

Bayesian analysis on high or low (i.e. respectively above or below a threshold of  $120kWhm^{-2}$  hourly solar irradiation) for each month with given prior beta distribution converged to posterior beta distribution after 20 iteration with average mean of 0.86. This shows that the on average the solar irradiation patterns in Kumasi tends to be high frequently. Also the prior variance of the various months of the year converged to the posterior tolerance level of 0.000001

Comparing the sunshine output of the two clusters indicated that cluster one have the highest hourly solar irradiation output than that of cluster two. The Bayesian analysis of the clusters also confirms that the first cluster have a higher possibility (0.8600) of solar irradiation output as compared with the second cluster (0.76000).

## CHAPTER ONE

### INTRODUCTION

#### 1.1 BACKGROUND

The mobilization of adequate national financial resources for the planning and development of the local solar energy resource depends on the availability of solar radiation data which could be used to evaluate available resources and to assess the probable long term performance of systems and hence their economic viability.

The solar radiation received at the earth's surface is subject to daily, seasonal and annual variations and hence many years of observation (perhaps at least 20 years) must be acquired in order to obtain a fairly accurate estimate of long term availability and distribution. However many locations in the developing countries do not have the facilities for continuous and accurate measurements of solar radiation and it is then necessary to use empirical methods which are based on easily measured meteorological parameters such as temperature, relative humidity, rainfall, cloudiness and duration of bright sunshine.

Many such formulae have been documented in the literature (Knight et al, 1991) although the most widely used correlation and perhaps the simplest, is the Angstrom (B) linear regression equation as modified by Page (1964) and others. This correlation relates the monthly average, daily global irradiation on the horizontal to the relative duration of sunshine, and it has been applied to a variety of climates including tropical locations. Except

for the recent work of Neba-Fabs et al., (1988) and of Exel(1978) nearly all the work done for locations in the West Africa sub-region and other tropical locations have sought to determine a single regression equation which could be used for all months and hence all seasons of the year. The results of Eze and Ododo (1988) and of other researchers however tend to indicate that the Angstrom-Page correlation coefficients depend on both the local climate and the season. Furthermore, it is anticipated that more accurate estimates of monthly average global irradiation would be obtained from correlations for particular months.

Liu and Jordan (1963) as well as Bendt et al (1981) conducted extensive statistical analysis of daily global irradiation on the horizontal particularly investigating possible variations of the frequency distribution with both location and season. Their results showed that frequency distribution of daily global irradiation on the horizontal for the monthly period corresponding to a specified value of a monthly mean clearness index, is almost independent of the location and the time of the year. Bendt *et al* (1981), moreover, went ahead and showed that the generalized cumulative distribution function may be obtained from a probability density function which assumed among others random daily insolation sequences.

In this study, we shall not dwell on the regression methods used to estimate the monthly global averages which, in any case, has already comprehensively been dealt with by Jackson et al (1990) but rather, taking advantage of the currently abundant data on solar irradiation data for Kumasi, we undertake the determination of the pertinent probability density curves based on randomly selected samples in respect of monthly or seasonal variations.

## 1.2 STATEMENT OF THE PROBLEM

There is the need to investigate the variability or otherwise of solar irradiation in terms of daily monthly seasonal or annual distributions. For example, one may need to determine or test whether there is a significant change from month to month or whether climatological conditions will affect the probability density functions. For purposes of applications knowledge of how universal are the probability distribution characterizations for a particular location is very important.

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## 1.3 OBJECTIVES

The major objectives of the study are as follows:

1. To determine the pertinent standard probability distributions functions in respect of the daily solar irradiation in Kumasi
2. To investigate the existence of significant differences in the pattern of solar irradiation within the year
3. To undertake a Bayesian analysis of the frequency of high solar irradiation.

## 1.4 JUSTIFICATION OF STUDY

The knowledge of the probability distribution function of solar irradiation will assist engineers to design systems based on months or seasons that have similar solar irradiation patterns.

## 1.5 METHODOLOGY

### 1.5.1 Nature and Source of data

Data on hourly solar irradiation in Kumasi was collected from the Solar Energy Laboratory of the Mechanical Engineering department of KNUST, Kumasi. The irradiation data which was measured in kilowatt hours per meter squared was collected by means of a pyranometer. The data consists of fourteen years of hourly solar irradiation data from 1995-2008.

### 1.5.2 Methods of data analysis

Analysis was conducted on the daily solar radiation data obtained in the past years to gain much insight of the data to constructively solve the problems as stated. The statistical methods adopted included

1. Descriptive statistics.
2. Curve fitting in respect of the various months of the year.
3. Inferential statistical analysis.
4. Bayesian analysis.

### 1.5.1 Computational tools

The various computational tools used included

1. Visual Basic
2. Microsoft Excel
3. INSTAT climatic software
4. SPSS

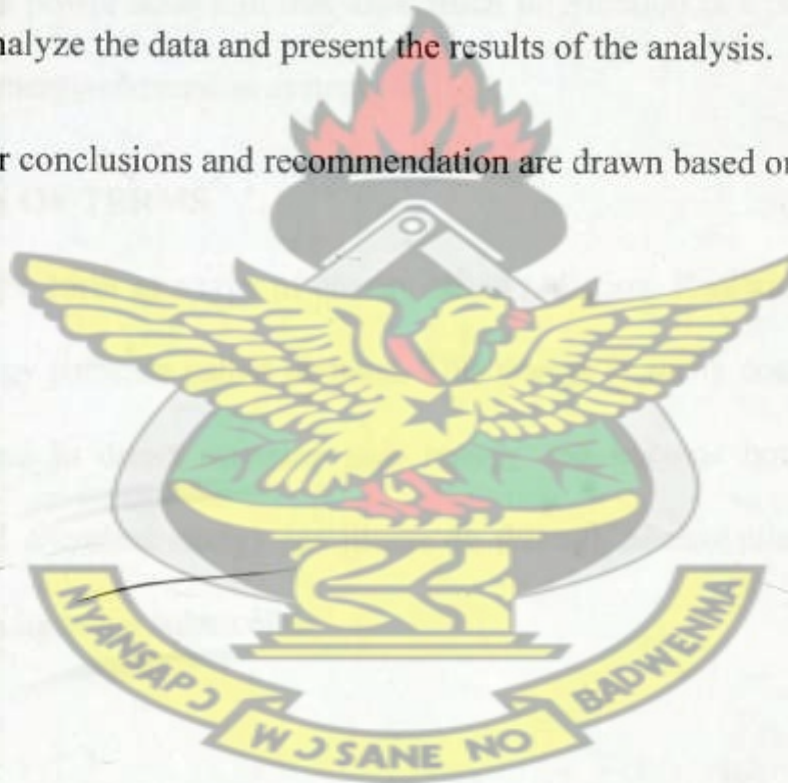
## 1.6 ORGANISATION OF THE STUDY

The thesis consists of four chapters .The first chapter comprises the introduction, background and methodology.

Chapter two seeks to review the relevant literature on the theory of statistics and probability as well as some relevant research work done on the statistical analysis of solar irradiation

In chapter three we analyze the data and present the results of the analysis.

Finally in chapter four conclusions and recommendation are drawn based on the findings.



## CHAPTER TWO

### LITERATURE REVIEW

Knowledge of the solar irradiation climate of an area is of paramount importance in assessing the potential use of a solar energy system, converted into either thermal or electrical energy, as a power source in that area. Such information is a prerequisite for the design of such solar energy conversion system.

#### 2.1 DEFINITION OF TERMS

Sunshine arrives on the earth as a type of energy called radiation. Radiation is composed of millions of high energy particles called photons. This energy is easily converted into heat energy (objects placed in direct sunshine gain energy and become hot). It can also be converted into stored chemical energy (as plants do through photosynthesis) or it can be converted into electricity using solar cells.

##### 2.1.1 Solar Irradiance

Solar irradiance refers to the solar radiation actually striking a surface, or the power received per unit area from the sun. This is measured in watts or kilowatts per square meter ( $W/m^2$  or  $kW/m^2$ ). If the solar module is facing the sun directly, irradiance will be much higher than if the module is at a large angle to the sun.

### 2.1.2 Insolation

Insolation (incident solar radiation) is a measure of the solar energy received on a specified area over a specified period of time. It is normally measured in kilowatt hours per square meter per day ( $\text{kWh}/\text{m}^2$  per day) or peak sun hours per day. It measures the amount of radiant energy collected at a site.

### 2.1.3 Pyranometer

It is the device used to record the sunlight data. It is made up of solar cell modules which harvest energy from the sun. The output of the solar cell modules depends on the amount of sunlight (or solar radiation) falling on them and it is affected by seasonal and daily solar radiation changes. It also changes depending on how cloudy or dusty the site is. It records two types of radiations: global average and diffused average.

The global average radiation is the hourly average irradiance of the direct solar energy reaching the earth's surface and the diffused average radiation is as a result of the direct solar energy being blocked by a black ring, clouds or dust before reaching the solar cell modules. However, this research work makes use of the global average radiation which is useful in the production of solar energy.

Clouds and dust absorb and scatter radiation, reducing the amount that reaches the ground.

On a sunny day most of the radiation is direct, but on a cloudy day, up to 100% of the radiation is diffuse.

## 2.2 REVIEW OF RELEVANT SOLAR IRRADIATION RESEARCH

Jackson *et al* studied Angstrom - Page type correlation between monthly average, daily global irradiation on the horizontal and the monthly average relative duration of sunshine has been derived for Kumasi, Ghana, using 20 years data of measured values by the Ghana Meteorological Services Department. The correlations were carried out at three levels: for each month of the year; for the dry, wet and harmattan seasons and for the full year. At the three levels, the maximum absolute errors in the estimated global irradiation are 0.59%, 4.24% and 8.34% respectively. The analyses were repeated after incorporating multiple reflections between the ground and sky and non-burning of the sunshine recorder chart when the solar elevation was below  $4^\circ$ . These latter considerations did not significantly affect their results.

The frequency distribution of insolation values for any location is essential for predicting long term performance of solar energy systems. Liu and Jordan (1963) first conducted statistical analysis of daily global irradiation on the horizontal for 27 locations. In approximately 5 years of data collection they produced cumulative distribution curves which were considered to have universal validity. Bendt *et al* (1981) conducted a study on the frequency distribution of daily insolation. In their study they observed that the data used by Liu and Jordan (1963) were limited both in location and duration of observation and they repeated the analysis using extensive data from 90 stations each for approximately 20 years of observation. They particularly investigated possible variations of the frequency distribution with both location and season. Their results truly confirmed Liu and Jordan's

observation, showing that frequency distribution of daily global irradiation on the horizontal for the monthly period corresponding to a specified value of a monthly mean clearness index, is almost independent of the location and the time of the year. Bendt *et al* (1981), moreover, went ahead and showed that the generalized cumulative distribution function may be obtained from a probability density function which assumed among others random daily insolation sequences.

An obvious consequence of the generalized cumulative distribution curves is that the maximum and the minimum values of the clearness index are also independent of location and season. Bendt *et al* (1981) suggested a constant value for the  $K_{min} = 0.05$  which corresponds to overcast sky conditions. Hollans and Huget published a report in 1983 titled 'a probability density function for the clearness index with application. In their paper they proposed that an expansion for the maximum value ( $k_{max}$ ) depends on only the monthly mean clearness index.

$$K_{max} = .6313k' - 11.9(k' - .75)^8$$

Saunier *et al* (1987) noted that the earlier analysis of the Liu and Jordan (1963) and also Bendt *et al* (1981) had used irradiation data from North America and they conducted similar analysis using 5 years of data for Bangkok, Thailand. Their results disagreed with the generalized CDC as proposed by Bendt *et al* (1981), and also proposed by Hollands and Hugets (1983) for the generalized CDC. Saunier *et al* (1987) observed that for other locations in Thailand as well as India, and they concluded that the generalized CDC may be invalid for tropical locations. However, satisfactory agreement was obtained when suitable values of  $K_{max}$  for Bangkok were used by the Bendt *et al* (1981). Consequently they

proposed an expression for  $K_{max}$ , based on observations in Thailand, to be used with the Bendt et al. CDC:

$$K_{max} = 0.362 + 0.59K' \dots\dots\dots (2)$$

They also suggested that the  $K_{max}$  defined by eqn(2) might be suitable for other tropical locations. Following the method used by Bendt *et al* (1981), Saunier *et al* (1987) also derived a higher order probability density of function for the daily global irradiation which provided better agreement with observation, as compared with the model by Bendt *et al* (1981). Again they suggested that their model may be valid for both tropical and temperate climates.

Saunier *et al* (1987) conclusion concerning the validity of the generalized CDC for tropical locations has been corroborated by the analysis of from other tropical locations. Feuillard *et al* (1989) noted that the CDC for tropical locations exhibits "a more pronounced S-shape" than that for temperate climates as a consequence of more pronounced peaks in the corresponding probability density functions. They also proposed a new probability density function and suggested that  $K_{max}$  be calculated from the annual correlation equation of daily clearness index with relative sunshine. This proposal is not in agreement with general observation, which clearly indicates that  $K_{max}$  varies and depends on the season.

The work of Olseth and Skartveit (1984) has also shown that the disagreement with the generalized CDC is not only confined to tropical climates. They observed that universal values of the maximum daily clearness index, as suggested by the generalized CDC are invalid; rather, the maximum daily clearness index is climatologically dependent.

Gordon and Reddy (1988) proposed an expression for the probability density function of daily global irradiation which they claim had universal validity. As noted by Knight *et al* (1991) this claim needs investigated further.

Akuffo and Hammond (1993) presented the results on statistical analysis of daily global irradiation for Kumasi, Ghana, also located in the tropics. Their study objective was to obtain the CDC and compare their findings with the generalized curves. Their work was to ascertain the validity of generalized curves for Ghana's conditions and consequently the applicability of available design procedures which are based on these curves. In addition, they compared their results with the model of Saunier *et al* and with the observations of IDeriah and Suleman (1989) for Ibadan, which is also located in the West Africa sub region and has similar agro-climatic conditions as pertains in Kumasi. The work by Akuffo and Hammond in 1993 demonstrated that Ibadan and Kumasi are all in the tropics though Kumasi had cloudier conditions. Again cumulative frequency curves for Kumasi differed from that of Ibadan in cases where there were discrepancies in  $K_{max}$  and  $K_{min}$ .

F. Youcef Ettoumi and *et al* (2002) published a paper on Statistical analysis of solar measurements in Algeria using beta distributions. In their paper a method of smoothing solar data by beta probability distributions was implemented. The method was used to process daily sunshine duration data recorded at thirty-three meteorological stations in Algeria for eleven year periods or more. Secondly the method has been applied to hourly global solar irradiation flux measured in Algiers during the 1987/89 period. For each location and each month of the year, beta probability density functions fitting the monthly frequency distributions of the daily sunshine duration measurements are obtained. Both the parameters

characterising the resulting beta distributions are then mapped, enabling them to build the frequency distributions of sunshine duration for every site in Algeria. In the case of solar radiation for Algiers, the recorded data have been processed following two different ways. The first one consists in sorting the hourly global solar irradiation data into eight typical classes of the daily clearness index. The second one is based on the repartition of these data per month. The results of the first classification show that for each class of daily clearness index, the hourly data under consideration are modeled by only one beta distribution. When using the second classification, linear combinations of two beta distributions are found to fit the monthly frequency distributions of the hourly solar radiation data.

## 2.3 REVIEW OF PROBABILITY THEORY

A Probability Space is a triple  $(\Omega, F, P)$  consisting of a set  $\Omega$  called the sample space, a  $\sigma$ -algebra  $F$  consisting of subsets of  $\Omega$  (these subsets are called events) and a measure  $P$  (called probability measure) on  $(\Omega, F)$  such that  $P(\Omega) = 1$  called the probability measure.

### 2.3.1 Random Variables

Let  $(\Omega, F, P)$  be a probability space and  $(Y, G)$  be a measurable space then the random variable  $X$  is defined as a measurable function  $X : \Omega \rightarrow Y$ . Where  $Y \subseteq \mathfrak{R}$  and  $G$  is a  $\sigma$ -algebra of  $Y$  consisting of subsets of  $Y$ .

A random variable  $X$  is said to be continuous if there exists a function  $f$ , called the probability density function or (probability distribution function (pdf)) of  $X$  and a function  $F$

called the cumulative distribution function (cdf), such that

$$F(x_2) - F(x_1) = \int_{x_1}^{x_2} f(x) dx = P\{x_1 \leq X \leq x_2\}$$

A random variable  $X$  will be called discrete if there exists a finite or countable set  $U$  of real numbers with  $U = \{x_1, x_2, \dots, x_n\}$  such that  $P\{U\} = \sum_{x_i \in U} P(X = x_i) = 1$ . The probability

distribution function (pdf) of a discrete random variable  $X$  and the cumulative distribution

function (cdf),  $F$  are given by  $f(x_i) = P(X = x_i)$  and  $F(x_n) = \sum_{k=0}^n f(x_k)$

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If  $X$  is a random variable such that  $f(x)$  is the probability distribution function then the expectation  $E$  and the variance  $Var$  of  $X$  are given as

$$E(X) = \begin{cases} \sum_x xf(x) & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} xf(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

$$Var(X) = \begin{cases} \sum_x (x - E(X))^2 f(x) & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} (x - E(X))^2 f(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

### 2.3.3 Properties of cdf of a Continuous Random Variable

1.  $F(X)$  is a monotonically non decreasing function.

2.  $\lim_{x \rightarrow -\infty} F(x) = 0$

3.  $\lim_{x \rightarrow \infty} F(x) = 1$

If  $F(X)$  is absolutely continuous if its derivative exists. Integrating the derivative gives back the (cdf) again. The random variable  $X$  is then said to have a probability density function (pdf).

$$f(x) = \frac{\partial F(x)}{\partial x}$$

For a set  $E \subseteq \mathfrak{R}$ , the probability of the random variable  $X$  being in  $E$  is

$$P(X \in E) = \int_{x \in E} dF(x) = \int_{x \in E} f(x) dx$$

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### 2.3.3 Partition and total probability rule

#### 2.3.3.1 Definition

A class of events  $H = \{H_1, H_2, \dots\}$  forms a partition of the sample space  $\Omega$  if the event excludes one another and one of them must occur. That means that

$$H_i \cap H_j = \phi \quad \forall i, j \text{ and}$$

$$H_1 \cup H_2 \cup \dots \cup \Omega$$

A partition is called finite or countable if it contains a finite or countable many infinite number of events.

#### 2.3.3.2 Definition

Given two partitions  $H = \{H_1, H_2, \dots\}$  and  $K = \{K_1, K_2, \dots\}$  we say that  $H$  is finer than  $K$  or  $H$  is refinement of  $K$  if for every  $H_i$  in  $H$  there exist an event  $K_j$  such that  $H_i \subset K_j$ .

### 2.3.3.3 Theorem

Let  $H = \{H_j, j = 1, 2, \dots\}$  be a positive finite or infinite partition and let  $A$  be an arbitrary event. Then

$$P(A) = P(A|H_1)P(H_1) + P(A|H_2)P(H_2) + \dots + P(A|H_n)P(H_n)$$

in the case of partition  $H$  into a finite number  $n$  of sets.

And

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$$P(A) = P(A|H_1)P(H_1) + P(A|H_2)P(H_2) + \dots$$

in the case of partition  $H$  into a countable number of set.

### 2.3.3.4 Conditional Probability

Let  $H = \{H_1, H_2, \dots\}$  be a positive partition of the sample space  $\Omega$  and let  $A, B$  be two event

with  $P(B) > 0$ , then  $P(A|B) = \sum_{j=1} P(A|B \cap H_j)P(H_j|B)$

**Proof**

Using the expression on the right hand side of the above equation it can be deduced that

$$\sum_{j=1} P(A|B \cap H_j)P(H_j|B) = \sum_{j=1} \frac{P(A \cap B \cap H_j)}{P(B \cap H_j)} \times \frac{P(B \cap H_j)}{P(B)}$$

$$\frac{1}{P(B)} \sum_{j=1} P(A \cap B \cap H_j) = \frac{1}{P(B)} P(A \cap B) \cap \bigcup_{j=1} H_j = \frac{P(A \cap B)}{P(B)} = P(A|B)$$

### 2.3.3.5 Bayes' Formula

We shall now consider a question opposite to the total probability formula. Given that the event occurred what is the probability of the event, of the partition ? The answer is contained in the theorem below.

### 2.3.3.6 Theorem (Bayes' formula)

Let  $H = \{H_1, H_2, \dots\}$  be a positive partition of  $\Omega$  and  $A$  be any event with  $P(A) > 0$  Then for any event  $H_k$  of the partition  $H$  we have

$$P(H_k|A) = \frac{P(A|H_k)P(H_k)}{P(A|H_1)P(H_1) + P(A|H_2)P(H_2) + \dots + P(A|H_n)P(H_n)}$$

## 2.4 EXAMPLES OF CONTINUOUS PROBABILITY DISTRIBUTIONS

### 2.4.1 The Uniform Distribution

The random variable  $X$  is uniformly distributed over the interval,  $[a, b]$  if it has probability density function defined as:

$$f(X) = \begin{cases} \frac{1}{a+b}, & a < X < b \\ 0 & \text{otherwise} \end{cases}$$

So that the mean and variance are:

$$E(X) = \frac{a+b}{2} \text{ and } Var(X) = \frac{(b-a)^2}{12}$$

The uniform distribution provides a simple probability model to describe a continuous random variable that can randomly assume any value between two points  $a$  and  $b$  ( $a < b$ ) on a line. The uniform probability density function has a rectangular shape over the interval  $[a, b]$  with height  $\frac{1}{b-a}$  so that the area under the density function equal to 1. Thus, the uniform distribution provides a good model for a continuous random variable whose values are uniformly distributed over an interval. For example:

1. If buses arrive at a given bus stop over 15 minutes and you arrive at the bus stop at a random time, the time you must wait for the next bus to arrive could be describe by the uniform distribution over the interval from 0 to 15.
2. In a milling operation, pieces of lumber less than 1 meter length are considered scrap. The distribution of the lengths of scrap lumber would have a uniform distribution on the interval from 0 to 1.

#### 2.4.2 The Exponential Distribution

The exponential probability distribution for the continuous random variable,  $X$  which represents an interval of time or space is defined as:

$$f(X) = \begin{cases} \theta e^{-\theta x} & x > 0 \\ 0 & \text{otherwise} \end{cases}$$

Where, the parameter  $\theta > 0$  is the mean number of events that occur in the given unit of time or space. The mean value of  $X$  (that is, the mean length of time or space between successive occurrences of the event and the variance are:

$$E(X) = \frac{1}{\theta} \quad \text{Var}(X) = \frac{1}{\theta^2}$$

Typical situations resulting in this random variable are:

1. The length of time until a machine or a component of it fails.
2. The length of time between arrivals at a car wash.
3. The length of time in a service line or queue.
4. The length of time between successive filing of claims in an insurance office.

From the above examples, the exponential distribution models situations in which the random variable represents waiting time or measurements of length of time between successive occurrences of an event. It plays an important role in Reliability Theory where we try to find the reliability of a system at time  $t$ .

### 2.4.3 The Normal Distribution.

The probability density function for the normal random variable,  $X$  which is simply called normal distribution is defined by:

$$f(X) = \begin{cases} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{X-\mu}{\sigma}\right)^2}, & -\infty \leq X \leq \infty \\ 0 & \text{otherwise} \end{cases}$$

$\sigma > 0$ ,  $-\infty < \mu < \infty$  and the mean and variance of the measurements,  $x$  are;

$$E(X) = \mu \text{ and } Var(X) = \sigma^2$$

If the random variable is modeled by the Normal distribution with mean  $\mu$  and variance,  $\sigma^2$  then it is simply denoted as  $x \sim N(\mu, \sigma^2)$ . The graph of a normal distribution is a bell-shaped smooth curve.

The Normal distribution is one of the most widely used probability distributions for modeling random experiments. It provides a good model for continuous random variables involving measurements such as time, heights/ weight of persons, marks scored in an examination, amount of rainfall, growth rate and many other scientific measurements.

The normal curve has the following desirable properties, accounting for the wide-spread applications of the Normal distribution.

1. The normal curve is symmetrical about its mean,  $\mu$ .
2. The mean, median and mode of  $x$  are the same.
3. The total area under the normal curve is equal to 1.
4. The probability distribution of the normal random variable,  $x$  is completely determined by its two parameters  $\mu$  and  $\sigma$ .
5. The curve is asymptotic to its horizontal axis,  $x$ .
6. The probabilities of the normal random variable,  $x$  are given by the areas under the curve.

#### 2.4.4 The Weibull Distribution

The continuous random variable  $X$  has a Weibull distribution, with parameters  $\alpha$  and  $\beta$  if its density function is given by

$$f(X; \alpha, \beta) = \begin{cases} \alpha \beta X^{\beta-1} e^{-\alpha X^\beta} & , X > 0 \\ 0, elsewhere & \end{cases}$$

where  $\alpha > 0$  and  $\beta > 0$ .

The mean and variance of the Weibull distribution are given by

$$\mu = \alpha^{-1/\beta} \Gamma\left(1 + \frac{1}{\beta}\right) \text{ and } \sigma^2 = \left\{ \Gamma\left(1 + \frac{2}{\beta}\right) - \left[ \Gamma\left(1 + \frac{1}{\beta}\right) \right]^2 \right\}$$

Modern technology has enabled us to design many complicated systems whose operation, or perhaps safety, depends on the reliability of various components. The Weibull distribution has been used extensively in recent times to deal with such problems. This is applied to reliability and life-testing problems such as time to failure or life length of a component, measured from specific time until it fails.

#### 2.4.5 The Gamma Distribution

The continuous random variable  $X$  has a gamma distribution, with parameter  $\alpha$  and  $\beta$ , if its density function is given by

$$f(X; \alpha, \beta) = \begin{cases} \frac{1}{\beta^\alpha \Gamma(\alpha)} X^{\alpha-1} e^{-x/\beta} & , X > 0 \\ 0, elsewhere & \end{cases}$$

Where,  $\alpha > 0$  and  $\beta > 0$ .

The gamma distribution for which  $\alpha = 1$  is the exponential distribution. The mean and variance of the gamma distribution are

$$\mu = \alpha\beta \quad \text{and} \quad \sigma^2 = \alpha\beta^2$$

The exponential thus turns out to be a special case of the gamma distribution. The exponential gamma distribution plays an important role in both queuing theory and reliability problems. Time between arrivals at service facilities, and time to failure of component parts and electrical systems, often is nicely modeled by the exponential distribution. The relationship between the gamma and exponential allows the gamma to be involved in similar types of problems

#### 2.4.6 Lognormal Distribution

The continuous random variable  $X$  has a lognormal distribution if the random variable  $Y = \ln(X)$  has normal distribution with mean  $\mu$  and standard deviation  $\sigma$ . The resulting density function of  $X$  is

$$f(X; \mu, \sigma) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma X} e^{-\frac{1}{2\sigma^2}[\ln(X) - \mu]^2}, & X > 0 \\ 0, & \text{elsewhere, } X < 0 \end{cases}$$

The mean and variance of the lognormal distribution are

$$\mu = e^{\mu + \frac{\sigma^2}{2}} \quad \text{and} \quad \sigma^2 = e^{2\mu + 2\sigma^2} (e^{\sigma^2} - 1)$$

The lognormal distribution is used for a wide variety of applications. The distribution applies in cases where a natural log transformation results in a normal distribution. The cumulative distribution function is quite simple due to its relationship to the normal distribution

#### 2.4.7 Beta Distribution

The family of distribution most commonly used to model researchers' uncertainty about the unknown probability  $p$  of some event is the family of beta distributions, defined as follows:

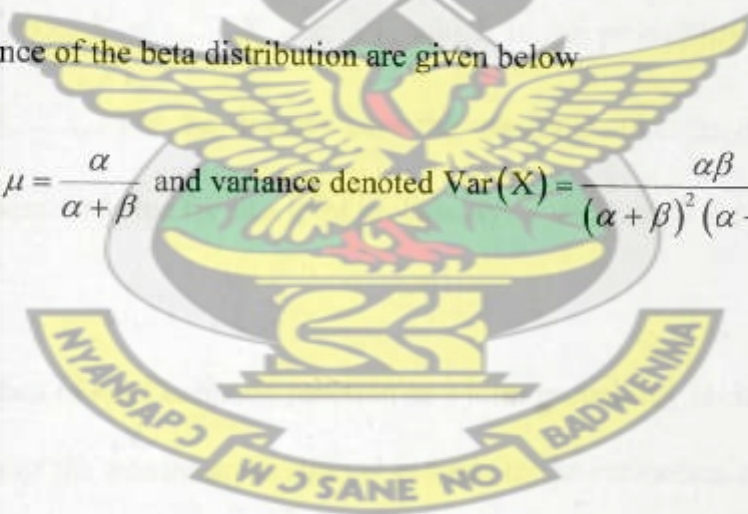
A random variable  $X$  is said to have a beta distribution with parameters  $\alpha > 0$   $\beta > 0$  if the

density of  $X$  is  $f(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$  for  $0 \leq x \leq 1$  and

$f(x) = 0$  outside the interval  $[0,1]$

The mean and the variance of the beta distribution are given below

Mean denoted  $E(X) = \mu = \frac{\alpha}{\alpha + \beta}$  and variance denoted  $\text{Var}(X) = \frac{\alpha\beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$



## 2.5.0 METHODS OF PARAMETER ESTIMATION

### 2.5.1 The Maximum Likelihood Estimator

The Maximum likelihood estimator represents one of the most important approaches to estimation in all statistical inference. The method of maximum likelihood is that for which the likelihood function is maximized. The likelihood function is best illustrated through the use of a discrete distribution and a single parameter.

Let  $x_1, x_2, \dots, x_n$  denote independent random variables taken from a discrete probability distribution represented by  $f(x, \theta)$ , where  $\theta$  is a single parameter of the distribution.

$L(x_1, x_2, \dots, x_n; \theta) = f(x_1, x_2, \dots, x_n; \theta) = f(x_1, \theta) f(x_2, \theta) \dots f(x_n, \theta)$  is the joint distribution of the random variables. The quantity  $L(x_1, x_2, \dots, x_n; \theta)$ , the likelihood of the sample, is the joint probability:  $P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n; \theta)$ , which is the probability of obtaining the sample values  $x_1, x_2, \dots, x_n$ . For the discrete case the maximum likelihood estimator is one that results in a maximum value for this joint probability or maximizes the likelihood of the sample.

Although the interpretation of the likelihood function as a joint probability is confined to the discrete case, the notion of the maximum likelihood extends to the estimation of parameters of a continuous distribution. The maximum likelihood estimation is as follows:

Given independent observation  $x_1, x_2, \dots, x_n$  from a probability density function (continuous) or probability mass function (discrete case)  $f(x, \theta)$ , the maximum likelihood estimator  $\hat{\theta}$ ,

is that which maximizes the likelihood function

$$L(x_1, x_2, \dots, x_n; \theta) = f(x_1, \theta) f(x_2, \theta) \dots f(x_n, \theta).$$

$$L(\theta) = f(x_1, \theta) \cdot f(x_2, \theta) \dots f(x_n, \theta) = \prod_{i=1}^n f(x_i, \theta), \theta \in \Omega$$

$$\frac{\partial \ln L(\theta)}{\partial \theta_1} = 0, \frac{\partial \ln L(\theta)}{\partial \theta_2} = 0, \dots, \frac{\partial \ln L(\theta)}{\partial \theta_k} = 0.$$

Solving the resulting system of equation will give

$$\hat{\theta} = h(x_1, x_2, \dots, x_n)$$

### 2.5.1.1 Properties of Maximum Likelihood Estimation

#### 2.5.1.2 Unbiasedness

Let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with probability distribution

$f(X, \theta)$ . An estimator  $\hat{\theta} = h(x_1, x_2, \dots, x_n)$  is said to be unbiased for  $\theta$  if  $E(\hat{\theta}) = \theta$

#### 2.5.1.3 Efficiency

Let  $\hat{\theta}_1$  and  $\hat{\theta}_2$  be two unbiased estimators for  $\theta$  with variance  $Var(\hat{\theta}_1)$  and  $Var(\hat{\theta}_2)$

respectively,  $\hat{\theta}_1$  is said to be relatively more efficient than  $\hat{\theta}_2$  if  $Var(\hat{\theta}_1) < Var(\hat{\theta}_2)$

#### 2.5.1.4 Definition

An unbiased estimator,  $\hat{\theta} = h(x_1, x_2, \dots, x_n)$  is said to be efficient if and only if

$$\text{Var}(\hat{\theta}) = \frac{1}{nE\left[\left(\frac{\partial \ln f(X, \theta)}{\partial \theta}\right)^2\right]}, \text{ Thus the efficient estimator for } \theta \text{ is the minimum variance}$$

unbiased estimator whose variance attains the Crammer-Rao Lower bound.

#### 2.5.1.5 Consistency:

Let  $\hat{\theta} = h(x_1, x_2, \dots, x_n)$  be an estimate of the parameter  $\theta$ . The estimator  $\hat{\theta}$  is said to be consistent for  $\theta$  if  $\lim_{n \rightarrow \infty} P\{|\theta - \hat{\theta}| > \varepsilon\} = 0$  OR  $\lim_{n \rightarrow \infty} P\{|\theta - \hat{\theta}| \leq \varepsilon\} = 1 \forall \theta$  and  $\varepsilon > 0$

#### 2.5.1.6 Theorem:

Let  $\hat{\theta} = h(x_1, x_2, \dots, x_n)$  be an estimate of the parameter  $\theta$  based on sample of size  $n$ . if

$$\lim_{n \rightarrow \infty} E(\hat{\theta}) = \theta \text{ and } \lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}) = 0 \text{ then } \hat{\theta} \text{ is a consistent estimator for } \theta$$

## 2.6 LEAST SQUARE ESTIMATES

For any given set of data, our problem or main objective is to draw the particular straight line that best reflects the linear trend indicated by the points. In other words, the problem is simply that of determining appropriate constants  $a$  and  $b$  associated with the straight line.

The principle involved in obtaining "the line of best fit" is called the method of least

squares. Suppose  $\hat{y} = a + bx$  is the line that best fits the data consisting of  $n$  pairs of values  $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$ , this is called the prediction equation.

For instance  $\hat{y}_i = a + bx_i$  represents the predicted value of dependent variable when the value of the independent variable is  $x_i$ .  $y_i$  is the corresponding observed  $y$ -value in the data. The quantities  $y_1 - \hat{y}_1, y_2 - \hat{y}_2, \dots, y_n - \hat{y}_n$ , that is  $y_1 - (a + bx_1), y_2 - (a + bx_2), \dots, y_n - (a + bx_n)$ , give deviations of the observed  $y$ -values and each is called the residual or error denoted  $e$ . Thus

$$e_i = y_i - \hat{y}_i.$$

The principle of the method of least squares is stipulated as follows. Of all possible straight lines that can be drawn on a scatter diagram, choose as the line of best fit the one for which the sum of the squares of the deviations of the observed  $y$ -values from the predicted  $y$ -values is a minimum. That is determining the constants  $a$  and  $b$  in such a way that

$$\sum e_i^2 = \sum [y_i - (a + bx_i)]^2 \text{ is the minimum.}$$

The straight line obtained using this criterion is called the least squares regression line. The line of  $a$  and  $b$  that determine the line of best fit can be obtained from the following formulas. The least squares formulas for the slope of best fit is given by

$$b = \frac{n \sum x_i y_i - \sum x_i \sum y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

And its  $y$ -intercept by

$$a = \bar{y} - b\bar{x}$$

Where  $n$  is the number of pairs of observations in the data,  $\bar{y}$  is the mean of the observed  $y$ -values and  $\bar{x}$  is the mean of the  $x$ -values.

## 2.7 PARAMETRIC HYPOTHESIS TESTING

In some practical problem of statistical inference we may be required to take decision concerning the parameters of the population instead of finding estimates for them. The following are some situations requiring such decisions:

1. A health personnel may claim that a drug is effective in 90% cases if it is administered.
2. The mean life span of a type of an electric bulb is at least 8,000 hours.
3. An accused, in a criminal trial, is always assumed to be innocent until proven otherwise.
4. An educationist may claim that two methods of teaching are equally effective.
5. An educational programme will result in improved communication between parents and children.
6. A medical researcher may hypothesize that a new drug is more effective than another one in curing a disease.

The above statements can be subjected to statistical verification using the sample observations.

Thus by means of hypothesis testing we are able to determine whether or not the statements are consistent with available data or evidence.

### 2.7.1 Definitions

1. A hypothesis is a statement, assertion or conjecture about the nature of one or more situations (or populations) to be studied.
2. Hypothesis Testing is a statistical procedure that uses a random sample data to determine whether a statement about a population should be rejected or not. Hypothesis testing involving population parameters are called Parametric or Classical tests.

### 2.7.2 Types of Hypothesis

In testing for the validity of a hypothesis we usually propose two types of hypotheses, namely;

1. Null hypothesis, denoted  $H_0$ , which is the tentative statement assumed to be true.
2. Alternative hypothesis denoted  $H_1$ , which contradicts the null hypothesis. It is accepted only when sufficient evidence exists to establish its truth.

### 2.7.3 Formulation of $H_0$ and $H_1$ .

When we wish to establish the validity of a statement about a population using the evidence obtained from a random sample data, the negative of the statement is what we take as the null hypothesis. The statement itself constitutes the alternative hypothesis. In some

applications, it is not obvious how  $H_0$  and  $H_1$  should be formulated. The following guidelines for developing hypotheses of three types of situations are suggested:

1. Testing Research Hypothesis: This is formulated as alternative hypothesis.
2. Testing the validity of a claim: This generally corresponds to the “innocent until proven guilty” analogy. The claim made is chosen as the null hypothesis while the challenge to the claim is taken as the alternative hypothesis.
3. Testing in decision-making situations: This occurs when the decision-maker must choose between two courses of action, one associated with  $H_0$  and the other with  $H_1$ . If example, the decision involves the population parameter,  $\theta$  we should have the two hypotheses formulated as:  $H_0: \theta = \theta_0$  against  $H_1: \theta \neq \theta_0$ , where  $\theta_0$  is a particular value of  $\theta$  and an instant action is taken if  $H_0$  is rejected.

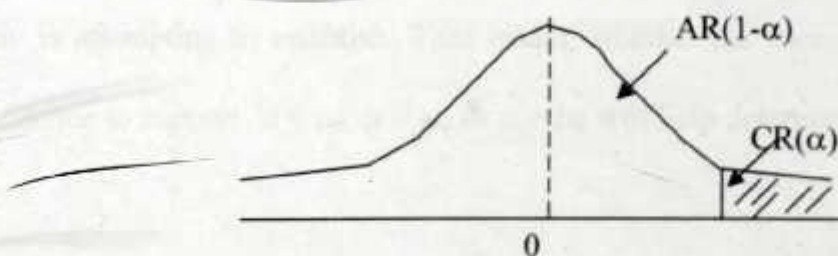
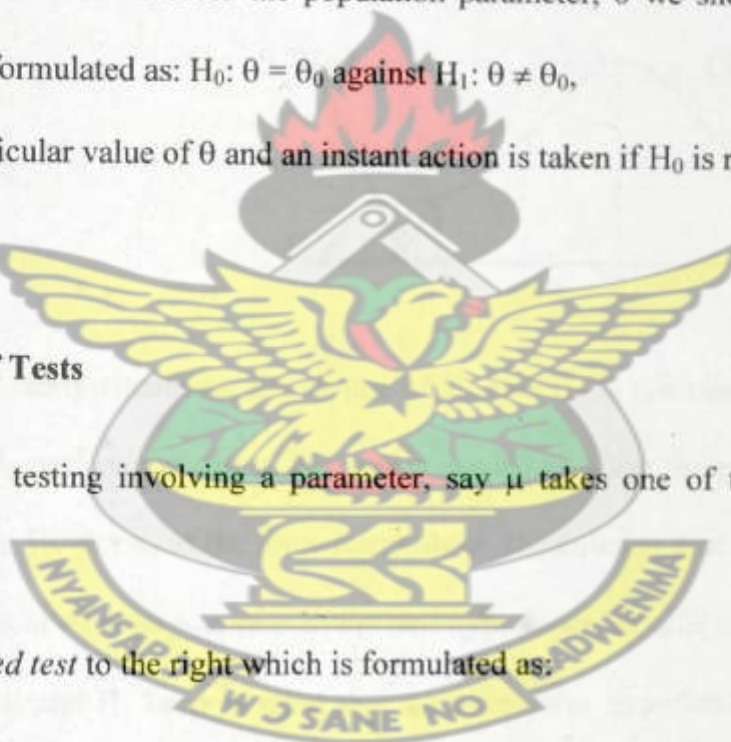
#### 2.7.4 Forms of Tests

In general a hypothesis testing involving a parameter, say  $\mu$  takes one of the following forms:

- (i) *One-Tailed test to the right* which is formulated as:

$$H_0: \mu \leq \mu_0$$

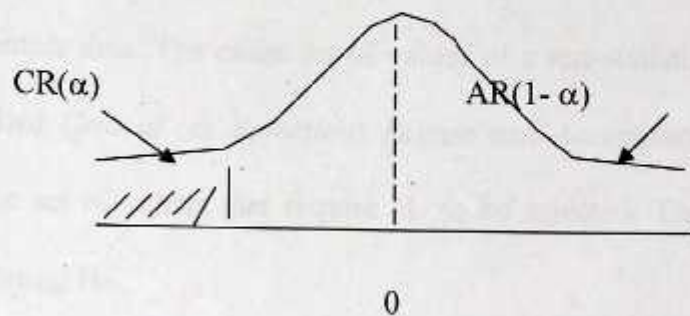
$$H_1: \mu > \mu_0$$



(ii) *One-tailed test to the left*, formulated as:

$$H_0: \mu \geq \mu_0$$

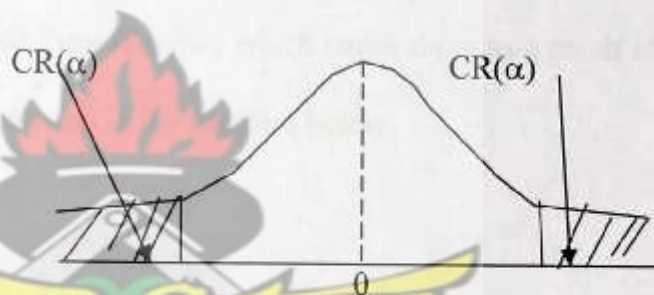
$$H_1: \mu < \mu_0$$



(iii) *Two-tailed Test*, which takes the form of choice between two courses of action.

$$H_0: \mu = \mu_0$$

$$H_1: \mu \neq \mu_0$$



(iv) *Remark:* In many situations, the choice of  $H_0$  and  $H_1$  is not obvious. In such cases, judgment on the part of the user is needed to select the proper form of  $H_0$  and  $H_1$ . However, as the three forms show, the equality part of expression (either  $\geq$ ,  $\leq$  or  $=$ ) always appear in the null hypothesis. In selecting the proper form of  $H_0$  and  $H_1$  keep in mind that the alternative hypothesis is what the sampling study is attempting to establish. Thus asking whether the user is looking for evidence to support  $\mu < \mu_0$ ,  $\mu > \mu_0$  or  $\mu \neq \mu_0$  will help determine  $H_1$ .

### 2.7.5 Test-Statistic

The test-statistic is a decision rule which leads to a rejection or acceptance of  $H_0$ . Its value is obtained from the randomly selected sample data. The entire set of values of a test-statistic is divided into two sets or regions called *Critical (or Rejection) Region* and *Acceptance Region*. The critical region contains the set of values that require  $H_0$  to be rejected. The acceptance region contains values supporting  $H_0$ .

### 2.7.6 Errors in Hypothesis Testing

When  $H_0$  is tested against  $H_1$  using a randomly selected sample data, two possible errors may be committed. These are the *Types I* and *Type II errors* which come about as a result of the decisions which are taken. These are illustrated in the diagram below.

Decision	Actual Situation	
	$H_0$ is true	$H_0$ is false ( $H_1$ is true)
Accept $H_0$	Correct decision ( $1 - \alpha$ )	Type II error ( $\beta$ )
Reject $H_0$	Type I error ( $\alpha$ )	Correct decision ( $1 - \beta$ )

- (i) Type I error is committed when the null hypothesis,  $H_0$  is rejected when in actual situation it is true. The probability of committing this error is  $\alpha$ , which is also referred to *level of significance* and indicates the size of the critical region.
- (ii) Type II error is committed when  $H_0$  is accepted when in actual situation it is false. The probability of committing this error is  $\beta$ .

### 2.7.6 P -Value

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The *p-value* is the smallest level of significance for which the observed data would call for rejection of  $H_0$  in favour of  $H_1$ . The *p-value* gives additional insight into the strength of the decision taken. A very small *p-value*, such as 0.0001, indicates that there virtually no likelihood that  $H_0$  is true. On the other hand, a high *p-value* such as 0.20 means that  $H_0$  is not rejected and there is little likelihood that it is false. The *p-value* is often referred to as the *observed level of significance*. For a given level of significance,  $\alpha$ , the null hypothesis,  $H_0$

1. is rejected if  $p\text{-value} \leq \alpha$ .
2. fails to be rejected if  $p\text{-value} > \alpha$ .

### 2.7.7 Elements of Statistical Tests

We note from the above discussion that a typical statistical test of hypothesis involves the following elements:

1. The Null and Alternative Hypotheses
2. The Test-Statistic
3. Critical or Rejected Region

### 2.7.8 Tests for the Population Means and Proportions

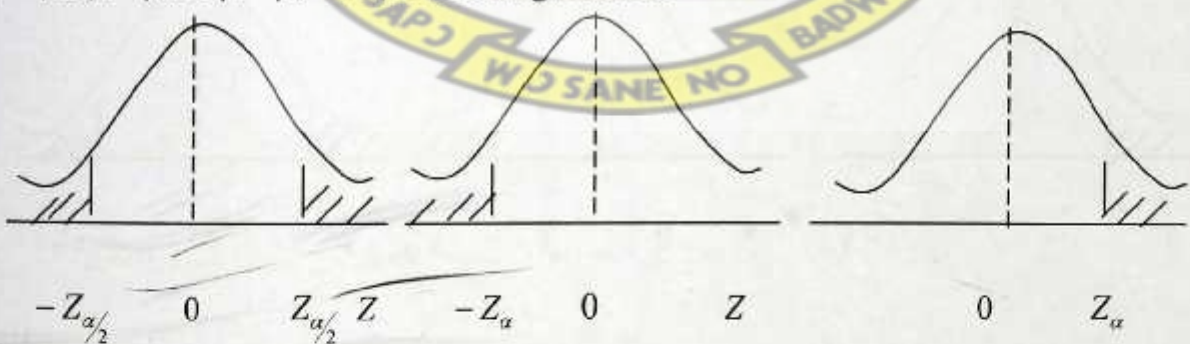
We shall discuss the tests for the population mean,  $\mu$  and the difference between two population means,  $(\mu_1 - \mu_2)$  taking into consideration for each situation the size of the sample(s). In addition, tests for population  $P$  and difference between two proportions,  $(P_1 - P_2)$  shall be discussed.

### 2.7.9 Testing for the Population Mean, $\mu$

(a) Case I: When the sample size,  $n$  is large.

(i) The hypotheses are  $H_0: \mu = \mu_0, \mu \geq \mu_0$  or  $\mu \leq \mu_0$

$H_1: \mu \neq \mu_0, \mu < \mu_0$  or  $\mu > \mu_0$  at  $\alpha$  level of significance.



$H_1: \mu \neq \mu_0$

$H_1: \mu < \mu_0$

$H_1: \mu > \mu_0$

- (ii) The test-statistic is 
$$Z = \frac{\bar{x} - \mu_0}{\sigma\sqrt{n}},$$

where  $\sigma$  is replaced by its estimate  $s$  if it is unknown.

- (iii) The decision rule is to reject  $H_0$  if the test-statistic ( $Z$ ) falls in the critical region. That is,  $H_0$  is rejected if  $|Z| \geq Z_{\alpha/2}$ , for two-tailed test; or  $Z \leq -Z_\alpha$  or  $Z \geq Z_\alpha$ , for one-tailed test.

- (iv) Assumptions: The sample observations are randomly and independently selected from a normally distributed population. The sample size,  $n \geq 30$ .

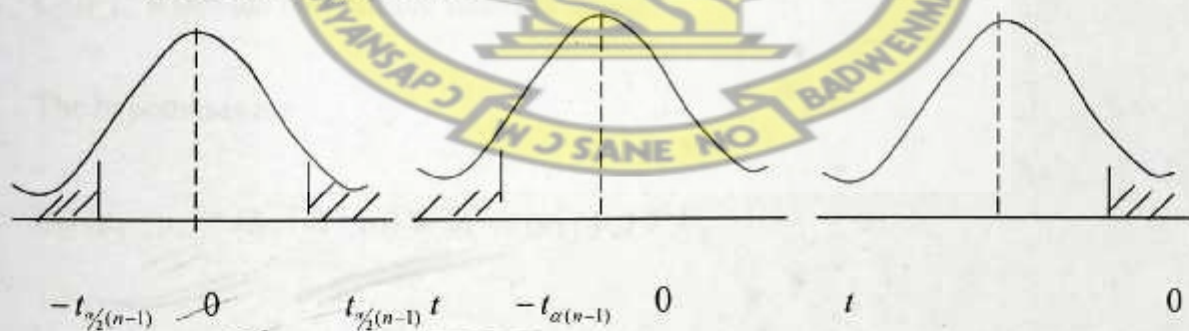
- (b) Case II: When the sample size,  $n$  is small

- (i) The hypotheses are

$$H_0: \mu = \mu_0, \mu \geq \mu_0 \text{ or } \mu \leq \mu_0$$

$$H_1: \mu \neq \mu_0, \mu < \mu_0 \text{ or } \mu > \mu_0$$

at  $\alpha$ -level of significance



$$t_{\alpha(n-1)}$$

$$H_1: \mu \neq \mu_0$$

$$H_1: \mu < \mu_0$$

$$H_1: \mu > \mu_0$$

(ii) The test-statistic is: 
$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$$

which has  $t$ -distribution with degrees of freedom,  $(n-1)$ .

(iii) The decision rule is to reject  $H_0$  if the test-statistic ( $t$ ) falls in the critical region. That is,  $H_0$  is rejected if  $|t| \geq t_{\alpha/2, (n-1)}$ , for two-tailed test; or

$t \leq -t_{\alpha, (n-1)}$  or  $t \geq t_{\alpha, (n-1)}$ , for one-tailed test.

(iv) Assumptions: Same as large sample size except that  $n < 30$  and  $\sigma^2$  is unknown.

### 2.7.10 Testing For The Difference Between Two Population Means, $(\mu_1 - \mu_2)$

(a) Case I: When the two sample sizes are large.

(i) The hypotheses are

$$H_0: (\mu_1 - \mu_2) = D_0, (\mu_1 - \mu_2) \geq D_0 \text{ or } (\mu_1 - \mu_2) \leq D_0$$

$$H_1: (\mu_1 - \mu_2) \neq D_0, (\mu_1 - \mu_2) < D_0 \text{ or } (\mu_1 - \mu_2) > D_0$$

at  $\alpha$  - level of significance, where  $D_0$  is the difference between the two means.

(ii) The test-statistic is: 
$$Z = \frac{(\bar{x}_1 - \bar{x}_2) - D_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

where  $\sigma_1^2$  and  $\sigma_2^2$  are replaced by  $s_1^2$  and  $s_2^2$  respectively if they are unknown.

(iii) The decision rule is to reject  $H_0$  if the test-statistic ( $Z$ ) falls in the critical region. That is,  $H_0$  is rejected if  $|z| \geq Z_{\alpha/2}$ , for two-tailed test; or if  $Z \leq -Z_\alpha$  or  $Z \geq Z_\alpha$ , for one-tailed test.

(b) Case II: When the same sizes are small.

(i) The hypotheses are

$$H_0: (\mu_1 - \mu_2) = D_0, (\mu_1 - \mu_2) \geq D_0 \text{ or } (\mu_1 - \mu_2) \leq D_0$$

$$H_1: (\mu_1 - \mu_2) \neq D_0, (\mu_1 - \mu_2) < D_0 \text{ or } (\mu_1 - \mu_2) > D_0$$

at  $\alpha$ -level of significance.

(ii) The test-statistic is: 
$$t = \frac{(\bar{x}_1 - \bar{x}_2) - D_0}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

which has t-distribution with degrees of freedom,  $(n_1 + n_2 - 2)$  and

$$s_p^2 = \frac{(n_1 - 1) s_1^2 + (n_2 - 1) s_2^2}{(n_1 + n_2 - 2)}$$

which is the estimate of a common variance assumed for the two populations

(iii) The decision rule is to reject  $H_0$  if the test-statistic ( $t$ ) falls in the critical region. That is,

$H_0$

is rejected if  $|t| \geq t_{\alpha/2, (n_1+n_2-2)}$ , for two-tailed test; or  $t \leq -t_{\alpha, (n_1+n_2-2)}$  or

$t \geq t_{\alpha, (n_1+n_2-2)}$  for

one-tailed test.

(iii) The assumptions are that the samples are randomly and independently selected from two normally distributed populations whose variances are the same. The sample sizes,  $n_1 < 30$  and  $n_2 < 30$ .

### 2.7.11 Testing for Population Proportions

(a) Test for the population proportion,  $P$ :

(i) The hypotheses are

$$H_0: P = p_0, P \geq p_0 \text{ or } P \leq p_0$$

$$H_1: P \neq p_0, P < p_0 \text{ or } P > p_0$$

at  $\alpha$ -level of significance

(ii) Test-statistic is: 
$$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$

where  $\hat{p} = \frac{x}{n}$  and  $n > 30$ .

(iii) The decision rule is to reject  $H_0$  if the test-statistic ( $Z$ ) falls in the critical region. That is,  $H_0$  is rejected if  $|Z| \geq Z_{\alpha/2}$ , for two-tailed test; or  $Z \leq -Z_{\alpha}$  or  $Z \geq Z_{\alpha}$ , for one-tailed test.

(b) Test for the difference between two Populations Proportions,  $(P_1 - P_2)$ ;

(i) The hypotheses are

$$H_0: (P_1 - P_2) = D_0, (P_1 - P_2) \geq D_0 \text{ or } (P_1 - P_2) \leq D_0$$

$$H_1: (P_1 - P_2) \neq D_0, (P_1 - P_2) < D_0 \text{ or } (P_1 - P_2) > D_0$$

at  $\alpha$ -level of significance

(ii) The test-statistics is: 
$$Z = \frac{(\hat{p}_1 - \hat{p}_2) - D_0}{\sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}}$$

If however,  $P_1 = P_2 = P$ , then  $D_0 = 0$  and the test-statistic becomes,

$$Z = \frac{(\hat{p}_1 - \hat{p}_2) - 0}{\sqrt{\frac{p(1-p)}{n_1} + \frac{p(1-p)}{n_2}}} = \frac{(\hat{p}_1 - \hat{p}_2)}{\sqrt{\bar{p}(1-\bar{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

where  $\bar{p} = \frac{x_1 + x_2}{n_1 + n_2} = \frac{n_1\hat{p}_1 + n_2\hat{p}_2}{n_1 + n_2}$ , the estimate for the common

proportion  $P$ ,  $n_1 > 30$  and  $n_2 > 30$ .

(iii) The decision rule is to reject  $H_0$  if the test-statistic ( $Z$ ) falls in the

critical region. That is,  $H_0$  is rejected if  $|Z| \geq Z_{\alpha/2}$ , for two-tailed test; or  $Z \leq -Z_\alpha$  or

$Z \geq Z_\alpha$ , for one-tailed test.

## 2.8 HISTORY OF BAYESIAN ESTIMATION

Though several key concepts were developed earlier—such as mathematical expectation (Huygens 1657), significance testing (Arbuthnot 1711), and the approximation of the binomial by the normal distribution (de Moivre 1718), many of the earliest statistical methods were developed in the latter part of the nineteenth century. For example, the concept of linear regression first appeared in the work of Galton (1889), building on the concept of least squares introduced by Legendre (1805) and arguably also by Gauss (who claimed that he had been using the method since 1795 (Gauss 1809)). Galton (1888) also introduced the concept of statistical correlation, which was later developed by Edgeworth (1893), Yule (1897) and Pearson (1896), who also began the development of goodness-of-fit measures (e.g. Pearson 1900) at around the turn of the century. The field really took off in the 1920s and 1930s when Fisher (1922, 1925) developed the notion of likelihood for

general estimation; Neyman & Pearson (1933) developed the basis for frequentist (often termed classical) hypothesis testing; and Yates & Cochran (1938) established the principles for the analysis of variance. Since then, the field has continued to grow and statistical inference now plays a key role in nearly every area of scientific research. Bayesian methods date to the original paper by the Reverend Thomas Bayes, read to the Royal Statistical Society in 1763 (the paper was in fact read by Richard Price several years after Bayes' death). The area generated interest from Laplace (1774), Gauss (1809) and Pearson (Pearson & Filon 1898) amongst others and dominated statistical thinking throughout the nineteenth century. The Bayesian approach fell out of favour at the beginning of the twentieth century due, in part, to the domination of the field by staunch opponents such as Neyman and Fisher, who held philosophical objections to the subjectivity of the Bayesian approach. Nonetheless, Bayesian stalwarts such as Jeffreys (1939), Savage (1954), Lindley (1965) and de Finetti (1970) kept the Bayesian flame alive by continuing to advocate Bayesian methods as remedies for certain deficiencies in the frequentist approach. It was not until the late 1980s that the Bayesian approach began to re-emerge, motivated both by rapid recent developments in computing and by the growing desire to describe increasingly complex scientific phenomena that older sampling theories were ill equipped to address. Since then the Bayesian approach has begun to dominate statistical research once more with new computational tools providing a far more flexible framework for statistical inference matching exactly the increasing complexity of scientific research as we move into the twenty first century.

### 2.8.1 Two Approaches to Statistical Inference

The two competing approaches to statistical inference are perhaps best explained in the context of a simple example. Suppose that we wish to toss a single coin  $X$ . Our statistical model places a probability, say  $\theta$  on it coming down heads ( $X = 1$ ) and a probability,  $1 - \theta$  of it coming down tails. ( $X = 0$ ) This is a statistical model parameterized by the unknown probability  $\theta$  which simply says that  $\text{Prob}(X = 1) = \theta$  and  $\text{Prob}(X = 0) = 1 - \theta$ . In fact, we can write down a general formula which covers both cases, i.e.

$$P(X = x) = \theta^x (1 - \theta)^{1-x} = f(x|\theta) \text{ for } x = 0 \text{ and } x = 1$$

Given a whole series of  $n$  coin tosses  $X_1, \dots, X_n$  we can write down a joint probability for all  $n$  outcomes:  $f(x_1, x_2, \dots, x_n | \theta)$ , which is read aloud as 'the joint probability distribution,  $f$ , of  $x_1$  up to  $x_n$  given the value of  $\theta$  and explicitly states the dependence of the series of observed coin tosses on the value of the parameter  $\theta$ . This joint probability distribution then forms the basis for statistical inference, whichever inferential approach is used.

### 2.8.2 The Objective Approach (Frequentist)

Statistical inference is essentially an inversion problem. The probability of an event is the proportion of times the event would occur if the experiment were repeated many times under identical conditions.

Given a parametric model and associated parameter values, it is possible to predict potential outcomes. For example, given the probability distribution above and the value of  $\theta$  it would

be possible to simulate or predict the outcomes of the coin-tossing experiment. However, in practice, it is the outcome of the experiment that is observed and the values of the parameters that are unknown. For example, we might observe three coin tosses each of which were heads and wish to infer from these observations what the value of  $\theta$  (i.e. the probability of a head) might be. In practical terms, this means that the frequentist approach to inference involves the maximization of a function of the model parameters. In some simple cases, this maximization can be performed analytically: by taking derivatives of the likelihood function, for example. However, for more complex examples numerical optimization procedures are required, which treat the likelihood function as a multi-dimensional surface where the height of the surface at any vector point  $\theta$  is given by the likelihood at that point. The frequentist approach to this problem relies on the strong repeated sampling principle and assesses the performance of our statistical estimation procedure for a single sample of data on the basis of the expected long-run performance given a hypothetical series of datasets collected under identical conditions.

### 2.8.3 The Bayesian Approach

The Bayesian approach to statistical inference is based upon Bayes' theorem (Bayes 1763). Consider the problem of finding a point estimate of the parameter  $\theta$  for the population with distribution  $f(x|\theta)$  given  $\theta$ . Denote  $\pi(\theta)$  the prior distribution of  $\theta$ .

Suppose that a random sample of size  $n$ , denoted by  $X = (x_1, x_2, \dots, x_n)$  is observed. For continuous distribution the bayes' theorem is defined as follows  $\pi(\theta|x) \propto f(x|\theta)\pi(\theta)$

Definition: The distribution of  $\theta$ , given data  $x$ , which is called the posterior distribution is given by

$$\pi(\theta | x) = \frac{f(x|\theta)\pi(\theta)}{g(x)} \quad \text{Where } g(x) \text{ is the marginal distribution of } x.$$

The marginal distribution in the above definition can be calculated using the following formula

$$g(x) = \begin{cases} \sum_{\theta} f(x|\theta)\pi(\theta), & \theta \text{ is discrete} \\ \int_{-\infty}^{\infty} f(x|\theta)\pi(\theta), & \theta \text{ is continuous} \end{cases}$$

#### 2.8.4 Differences in Approach

As in the frequentist case above, the Bayesian approach ends up with a function of the model parameters given the observed data, but this time our inference is in the form of a probability distribution for  $\theta$  rather than a simple point estimate. This illustrates one of the fundamental differences between the beliefs of the Bayesian and frequentist statisticians. The frequentist statistician believes that there exists a fixed true value for the model parameters and tries to estimate this value by maximizing the likelihood. The Bayesian believes that the unknown parameters have a fixed but unknown distribution which represents their beliefs about those parameters having observed the data. Note that the Bayesian may also believe that a true value exists, but since this can never be known with absolute certainty they prefer to think instead in terms of a distribution which reflects this

uncertainty. As their level of information increases, the width of this distribution decreases and, in the limit

### 2.8.5 Principles of Bayesian Statistical Analysis

There are four basic principles of Bayesian statistical analysis

1. Specify an objective probability model of the trials in terms of some unknown parameters.
2. Form subjective beliefs about the unknown parameters. These original beliefs are called the prior probabilities.
3. Update beliefs in light of the information contained in the sample by applying Bayes' Rule. The updated beliefs are called the posterior probabilities.
4. Base decisions on the updated beliefs (i.e., the posterior probabilities).

Remember that the prior distribution shows your beliefs about different parameter values before seeing any data. The likelihood essentially shows what the observed data tells about how probable different parameter values are. The posterior probability combines the information in these two distributions, and shows your updated beliefs about parameter values after having seen the data.

In frequentist statistics, the goal is to obtain good point estimates of parameter values.

In Bayesian statistics, the goal is instead to obtain a probability distribution over all possible parameter values. This is the posterior probability distribution. It shows how uncertain we

are about the parameter value, and can be used as the basis for asking many different questions.

### 2.8.6 Bayesian Prior, Posteriors and Estimators

If  $Y_1, Y_2, \dots, Y_n$  denote the random variables associated with a sample of size  $n$ . Let the notation  $L(y_1, y_2, \dots, y_n | \theta)$  denote the likelihood of the sample. In the discrete case, this function is defined to be the joint probability  $P(Y_1 = y_1, Y_2 = y_2, \dots, Y_n = y_n)$  and in the continuous case, it is the joint density of  $Y_1, Y_2, \dots, Y_n$  evaluated at  $y_1, y_2, \dots, y_n$ . The parameter  $\theta$  is included among the argument of  $L(y_1, y_2, \dots, y_n | \theta)$  to denote this function depends explicitly on the value of some parameter  $\theta$ .

In Bayesian approach, the unknown parameter  $\theta$  is viewed to be a random variable with a probability distribution called the prior distribution of  $\theta$ . This prior distribution of  $\theta$  is specified before any data are collected and provides a theoretical description of information about  $\theta$  that was available before any data were obtained. In our initial discussion we will assume that the parameter  $\theta$  has a continuous distribution with density  $g(\theta)$  that has no unknown parameters. Using the likelihood of the data and the prior on  $\theta$ , it follows that the joint likelihood  $Y_1, Y_2, \dots, Y_n, \theta$  is

$f(y_1, y_2, \dots, y_n, \theta) = L(y_1, y_2, \dots, y_n | \theta) \times g(\theta)$  And that the marginal density or mass function of

$Y_1, Y_2, \dots, Y_n$  is  $m(y_1, y_2, \dots, y_n, \theta) = \int_{-\infty}^{\infty} L(y_1, y_2, \dots, y_n | \theta) \times g(\theta) d\theta$  also the posterior density

$$\text{function of } \theta | y_1, y_2, \dots, y_n \text{ is } g^*(\theta | y_1, y_2, \dots, y_n) = \frac{L(y_1, y_2, \dots, y_n | \theta) \times g(\theta)}{\int_{-\infty}^{\infty} L(y_1, y_2, \dots, y_n | \theta) \times g(\theta) d\theta}$$

Let  $Y_1, Y_2, \dots, Y_n$  denote a random sample from a Bernoulli distribution of  $n$  observed morning

where  $P(Y_i = 1) = p$  and  $P(Y_i = 0) = 1 - p$  and assume that the prior distribution for  $p$  is

$\text{beta}(\alpha, \beta)$

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## 2.9 BETA AS CONJUGATE PRIOR

What is the posterior distribution for  $p$  where  $p$  is the probability of sunshine?

Since the Bernoulli probability function can be written as

$P(y_i | p) = p^{y_i} (1 - p)^{1 - y_i}$ ,  $y_i = 0, 1$ , the likelihood  $L(y_1, y_2, \dots, y_n | p)$  is

$$\begin{aligned} L(y_1, y_2, \dots, y_n | p) &= P(y_1, y_2, \dots, y_n | p) \\ &= p^{y_1} (1 - p)^{1 - y_1} \times p^{y_2} (1 - p)^{1 - y_2} \times \dots \times p^{y_n} (1 - p)^{1 - y_n} \\ &= p^{\sum y_i} (1 - p)^{n - \sum y_i}, \quad y_i = 0, 1 \text{ and } 0 < p < 1 \end{aligned}$$

The joint probability distribution is

$$\begin{aligned} f(y_1, y_2, \dots, y_n, p) &= L(y_1, y_2, \dots, y_n | p) \times g(p) \\ &= p^{\sum y_i} (1-p)^{n-\sum y_i} \times \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1} \\ &= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\sum y_i + \alpha - 1} (1-p)^{n - \sum y_i + \beta - 1} \end{aligned}$$

And the marginal density function is

$$\begin{aligned} m(y_1, y_2, \dots, y_n) &= \int_0^1 \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\sum y_i + \alpha - 1} (1-p)^{n - \sum y_i + \beta - 1} dp \\ &= \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \times \frac{\Gamma(\sum y_i + \alpha)\Gamma(n - \sum y_i + \beta)}{\Gamma(n + \alpha + \beta)} \end{aligned}$$

The posterior density of  $p$  is given by

$$g^*(\theta | y_1, y_2, \dots, y_n) = \frac{L(y_1, y_2, \dots, y_n | \theta) \times g(\theta)}{\int_{-\infty}^{\infty} L(y_1, y_2, \dots, y_n | \theta) \times g(\theta) d\theta}$$

Hence

$$\begin{aligned} g^*(p | y_1, y_2, \dots, y_n) &= \frac{\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} p^{\sum y_i + \alpha - 1} (1-p)^{n - \sum y_i + \beta - 1}}{\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \times \frac{\Gamma(\sum y_i + \alpha)\Gamma(n - \sum y_i + \beta)}{\Gamma(n + \alpha + \beta)}} \\ &= \frac{\Gamma(n + \alpha + \beta)}{\Gamma(\sum y_i + \alpha)\Gamma(n - \sum y_i + \beta)} \times p^{\sum y_i + \alpha - 1} (1-p)^{n - \sum y_i + \beta - 1}, \quad 0 < p < 1 \end{aligned}$$

The posterior beta density function is with the parameters

$$\alpha^* = \sum y_i + \alpha \text{ and } \beta^* = n - \sum y_i + \beta$$

## CHAPTER THREE

### 3.1 ANALYSIS OF DATA

#### 3.1.1 Nature and Source of Data

The solar radiation data was obtained from the Department of Mechanical Engineering KNUST. It comprised 14 years (1995-2008) of daily global irradiation, compiled from hourly measurements at the KNUST Solar laboratory. The observations were made with Bellan distillation pyranometers. These radiometers are robust and simple and are therefore appropriate for the prevailing conditions in many developing countries. In spite of these observations, the data for Kumasi is consistent with measurements from locations in the sub region with similar climatic and vegetation characteristics

#### 3.1.2 Data Quality Information

A lot of data is usually lost due to malfunctioning of the Pyranometer sensor at certain periods. The data from 2005 – 2007 for instance were lost, because during that period the calibrator of the Pyranometer was undergoing repairs. Moreover, there are few missing data points in the 1995 and 1996 data. In 1997, some of the sunlight measurements were negatives so these data may not be reliable since their true accuracy is unknown. It is therefore highly probable that useful information and details might have been lost in compiling data, which could affect results of the analysis.

### 3.1.2 Scope of Analysis

Sunlight has a lot of uses, especially in Africa and for that matter Ghana where it is in abundance. It is useful in hybrid lighting, drying clothes and bricks, etc. This research work however, focuses mainly on fitting a probability distribution function for the month of the year and also cluster the various months of the year.

Moreover analysis will also include Bayesian analysis of the daily solar irradiation in Kumasi-Ghana.

### 3.1.3 Compilation of Data

Due to the large size of the secondary data, and also the fact that total amount of energy available changes considerably from month to month, the data was grouped into months ranging from 1995 – 2008. With the help of the data analysis tool in excel, one thousand five hundred random samples were selected from the population for each month for the from 1995-2008 available data. (Consisting of over 7,000 data points) and analyzed as follows. A detailed presentation of the primary data is provided in the appendix of this document.

### 3.1.4 Analysis of Data

The analysis adopted both descriptive and inferential methods. The descriptive methods used are frequency and relative frequency distribution tables, probability density curves, and

histograms. Hypothesis testing, least squares estimation, and maximum likelihood estimation are the only inferential statistics used.

### 3.1.5 Frequency Distribution

Since the sampled data were too many to gain an insight into, and impossible for one to convey much information about its characteristics, it became necessary to organize and reduce it into meaningful forms as follows:

The data for each month was organized into groups, called classes or categories using the Sturges (approximation) rule, which states that:

The number of classes,  $K = 1 + 3.322 \log_{10} n$ . and class width,  $C = \frac{\text{Range}}{K}$

Where  $n$  is the total number of observations or frequencies and

Range = (maximum – minimum) of the observations.

### 3.1.7 Probability Distributions

To make inferences with respect to the population distribution of each month, some standard continuous probability distributions were selected based on their shape parameters to fit the data for each month. The distributions are the exponential, Weibull, Gamma, Lognormal and Beta, Least square and maximum likelihood estimation procedures were then applied to estimate the parameter for each distribution from the data. The probability density function and their respective histogram for the various months are shown below.

### 3.1.9 Mean Squared Errors

The mean squared error of the monthly data and each distribution selected were computed to

come up with the best candidate fitting the data as follows: 
$$MSE = \frac{\sum_{i=1}^n (rf_i - f(x_i))^2}{n}$$

Where  $rf_i$  = each relative frequency for the month

$f(x_i)$  = values computed using the probability distribution under consideration.

$x_i$  = the class mark(x), i.e. the hourly solar irradiance for the month under consideration.

## 3.2 PRESENTATION OF RESULTS

### 3.2.1 Histogram and Probability Density Function

A random sample of size one thousand five hundred was selected from each month of the year and probability density function and histogram in respect of hourly solar irradiance for the various months of the year was produced.

#### 3.2.1.1 Probability Density Function and Histogram for January

The histogram and the probability density function for the month of January are presented in the Fig 3.1 and Fig 3.1a below respectively.

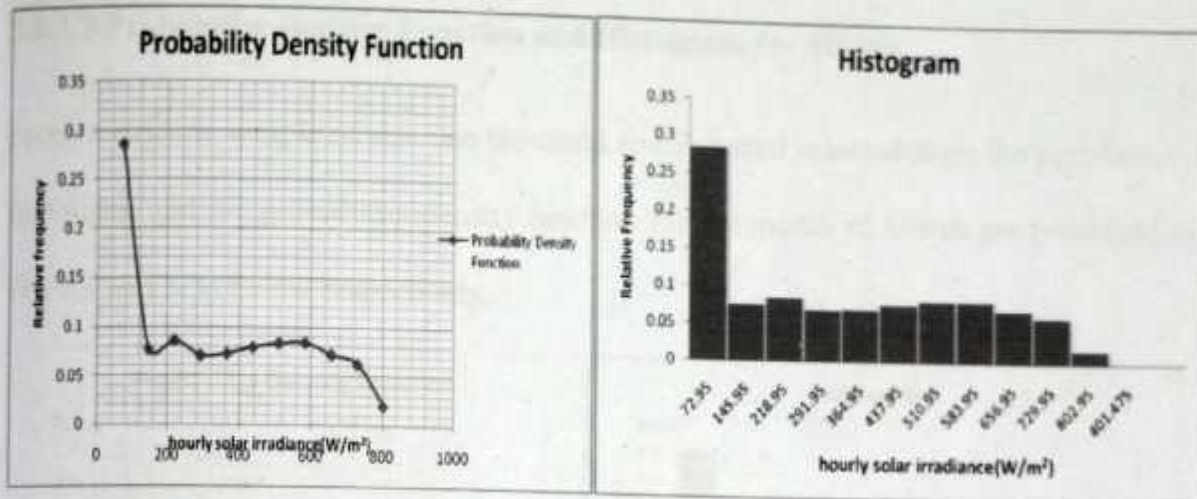


Fig 3.1:Probability density function for Jan

Fig 3.1a:Histogram for the month of Jan

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### 3.2.1.2 Probability Density Function and Histogram for February

From a random sample of size one thousand five hundred was selected from the population .The histogram and the probability density function for the month of February are presented in the Fig 3.2 and Fig 3.2a respectively.

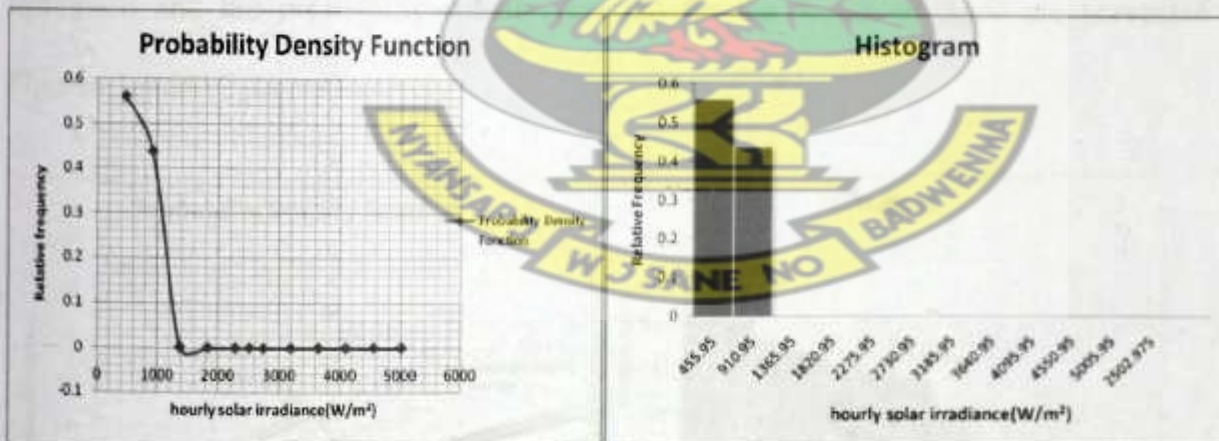


Fig 3.2:Probability density function for Feb

Fig 3.1a:Histogram for the month of Feb

### 3.2.1.3 Probability Density Function and Histogram for March

From a random sample of size one thousand five hundred selected from the population .The histogram and the probability density function for the month of March are presented in the Fig 3.3 and 3.3a below respectively

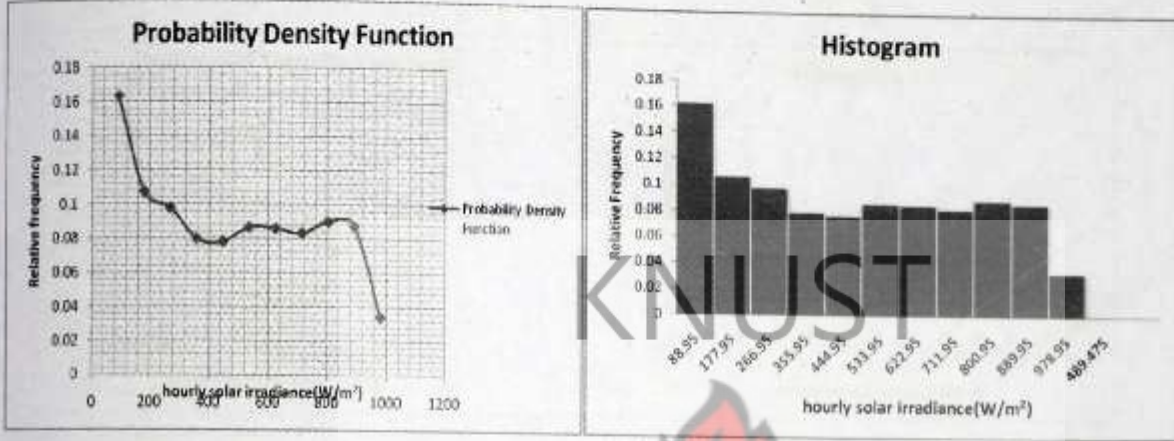


Fig 3.3:Probability density function for march

Fig 3.3a:Histogram for the month of march

### 3.2.1.4 Probability Density Function and Histogram for April

From a random sample of size one thousand five hundred selected from the population .The histogram and the probability density function for the month of April are presented in Figures 3.4 and 3.4a respectively.

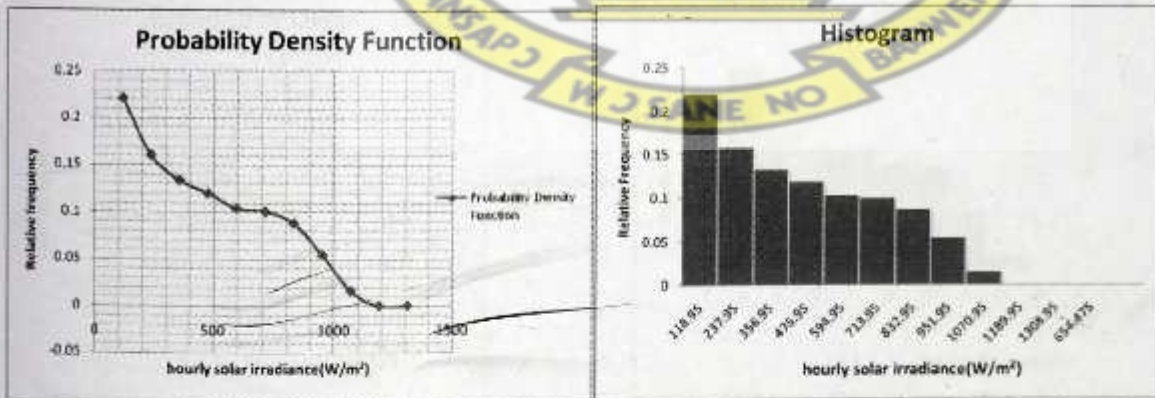


Fig 3.1: Probability density function for April

Fig 3.11: Histogram for the month of April

### 3.2.1.5 Probability Density Function and Histogram for May

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function for the month of May were obtained and are presented in Figures 3.5 and 3.5a respectively.

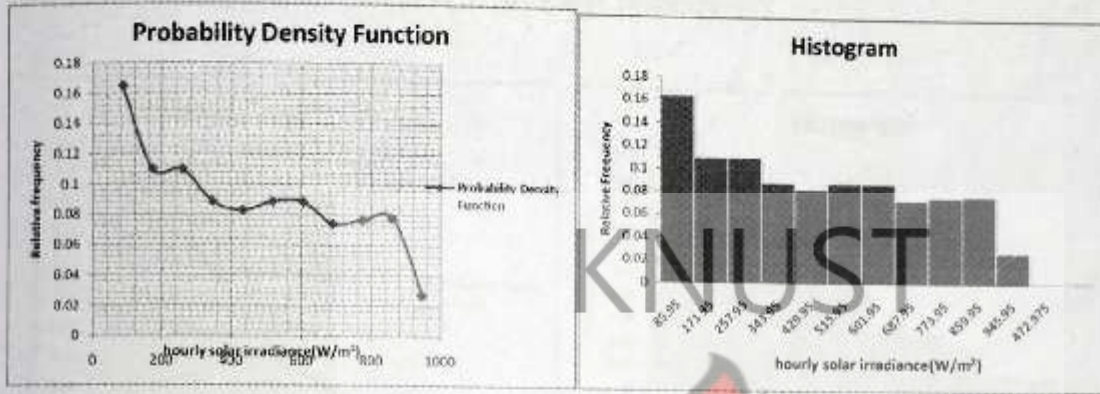


Fig 3.5: Probability density function for May

Fig 3.5a: Histogram for the month of May

### 3.2.1.6 Probability Density Function and Histogram for June

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function for the month of June were obtained and are here presented in Figures 3.6 and 3.6a below respectively.

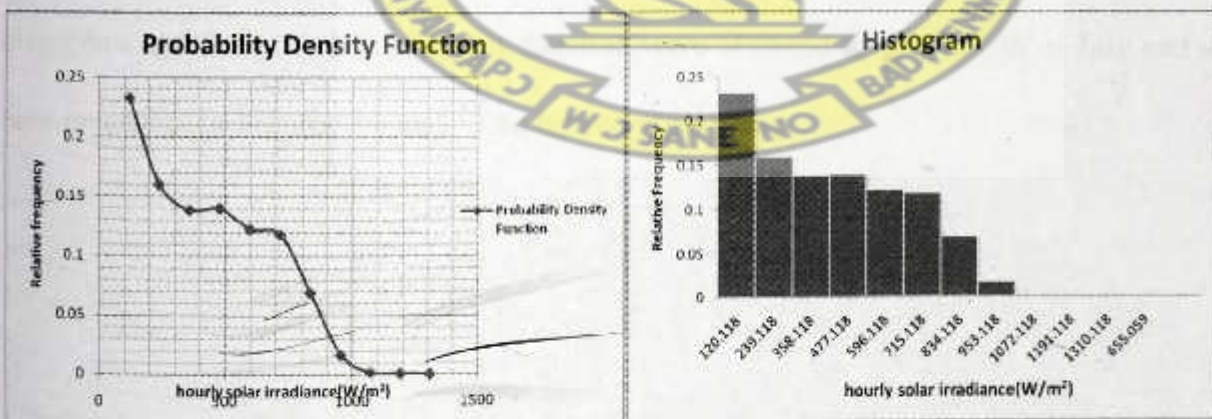


Fig 3.6: Probability density function for June

Fig 3.6a: Histogram for the month of June

### 3.2.1.7 Probability Density Function and Histogram for July

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function were obtained for the month of July and are here presented in Figures 3.7 and 3.7a below respectively

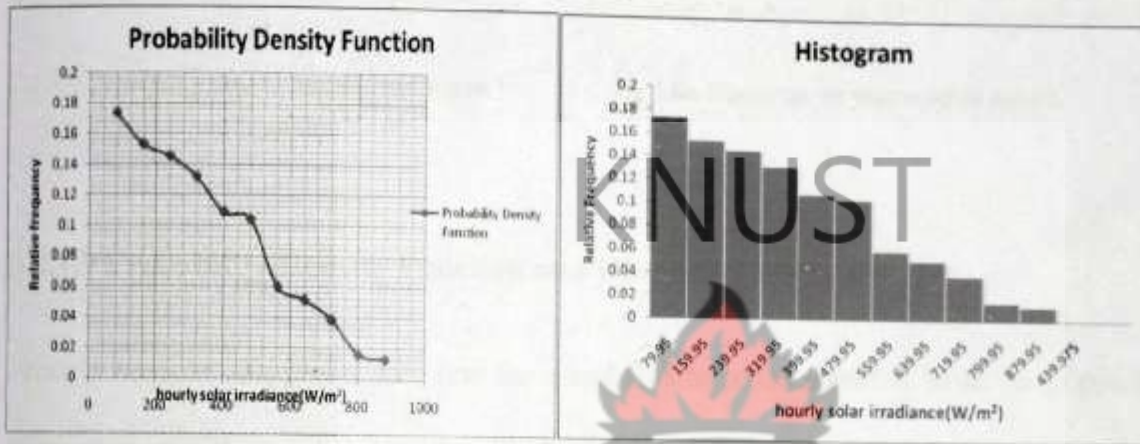


Fig 3.7: Probability density function for July

Fig 3.7a: Histogram for the month of July

### 3.2.1.8 Probability Density Function and Histogram for August

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function were obtained for the month of July and are here presented in Figures 3.8 and 3.8a below respectively.

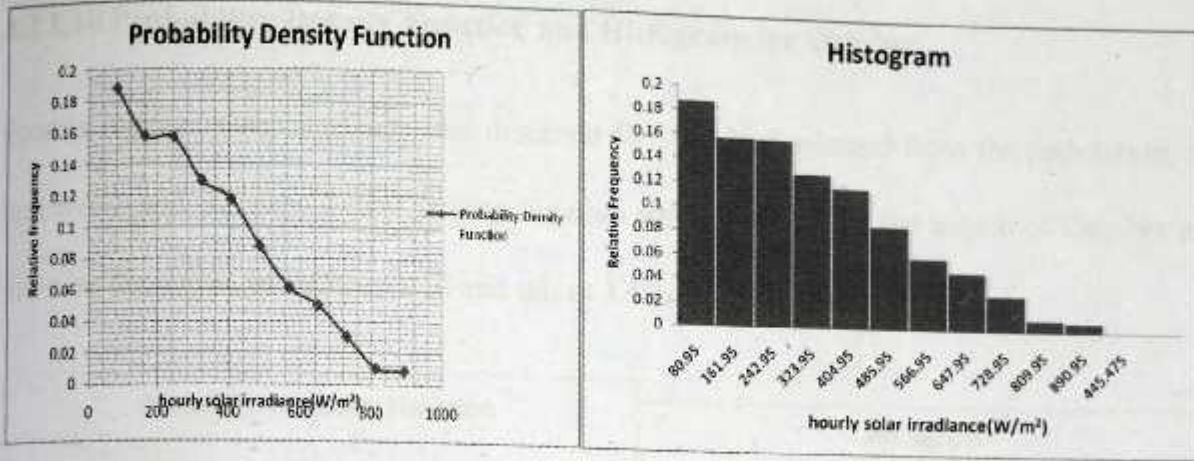


Fig 3.8: Probability density function for August

Fig 3.8a: Histogram for the month of August

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## 3.2.1.9 Probability Density Function and Histogram for September

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function were obtained for the month of September and are here presented in Figures 3.9 and 3.9a below respectively.

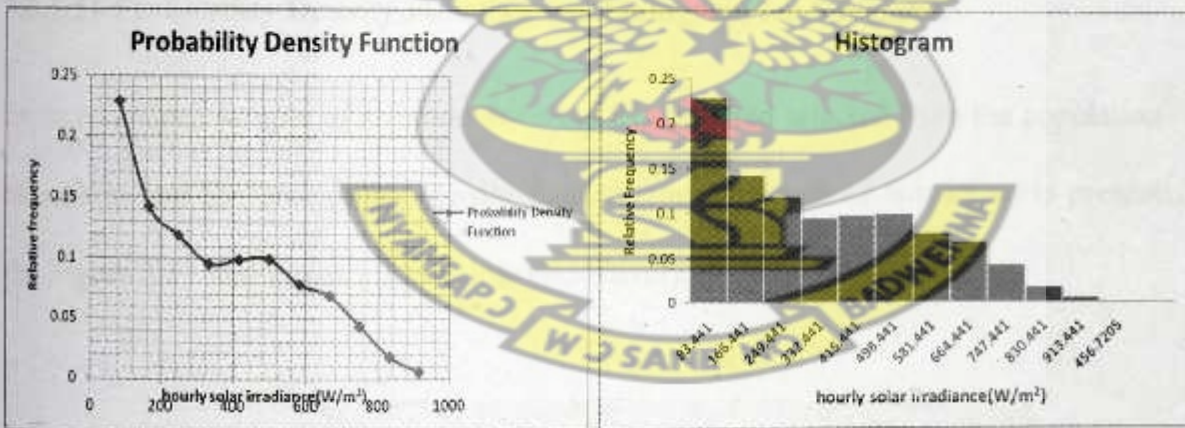


Fig 3.9: Probability density function for September

Fig 3.9a: Histogram for the month of September

### 3.2.1.10 Probability Density Function and Histogram for October

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function were obtained for the month of October and are here presented in Figures 3.10 and tables 3.10a below respectively

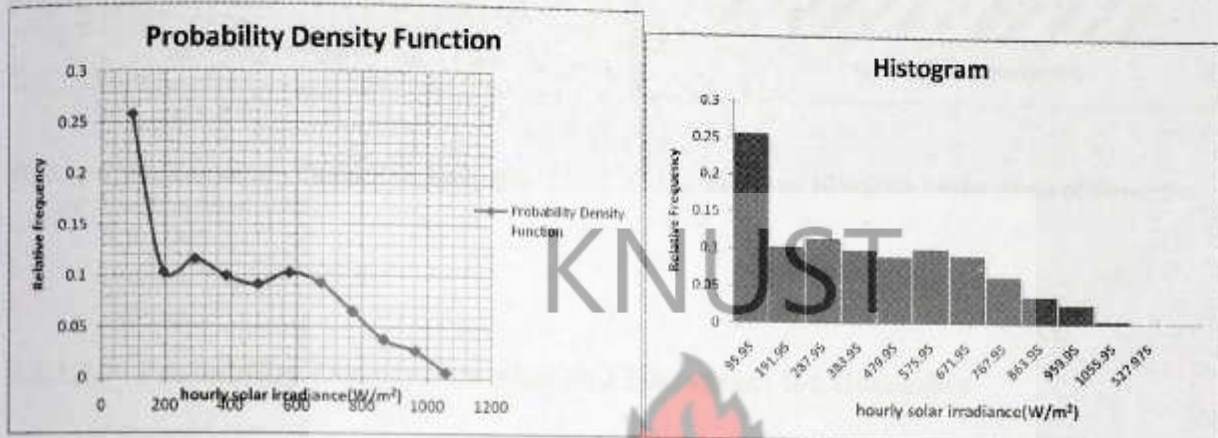


Fig 3.10:Probability density function for October

Fig 3.10a:Histogram for the month of October

### 3.2.1.11 Probability Density Function and Histogram for November

From a random sample of size one thousand five hundred selected from the population .The histogram and the probability density function for the month of November is presented in the Figure 3.11 and tables 3.11a below respectively.

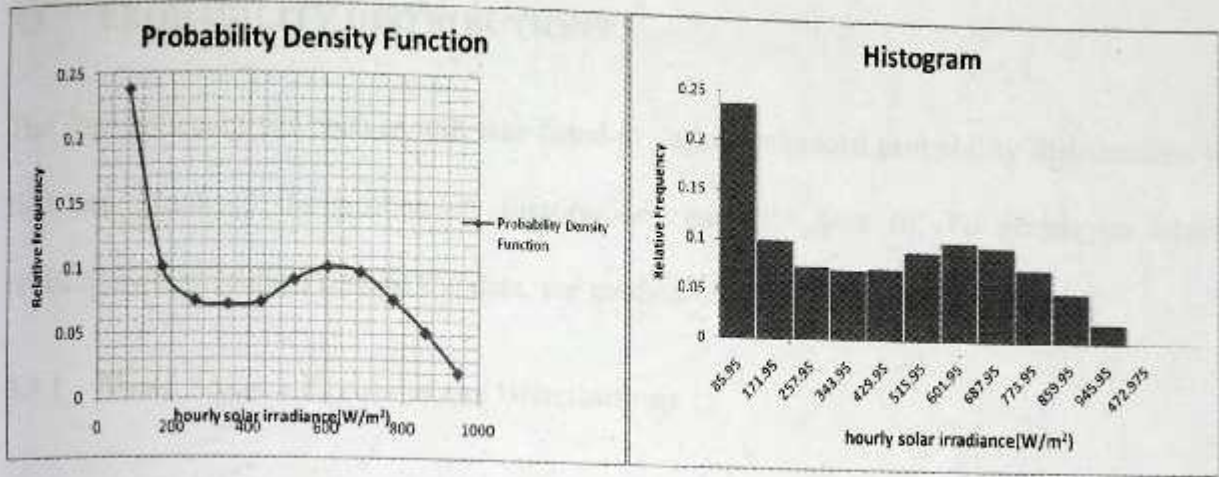


Fig 3.11: Probability density function for November

Fig 3.11a: Histogram for the month of November

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### 3.2.1.12 Probability Density Function and Histogram for December

From a random sample of size one thousand five hundred selected from the population, the histogram and the probability density function were obtained for the month of December and are here presented in Figures 3.12 and 3.12a respectively.

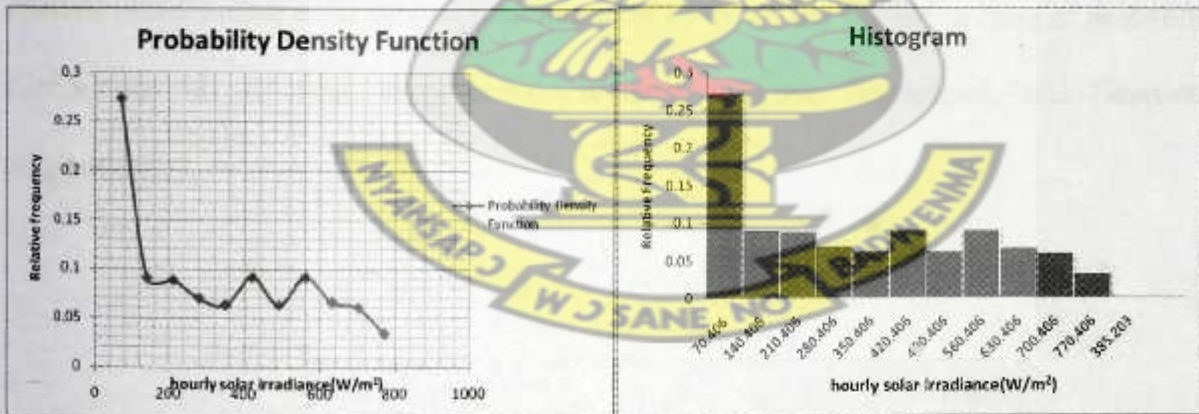


Fig 3.12: Probability density function for Dec

Fig 3.12a: Histogram for the month of December

### 3.3 PROBABILITY DISTRIBUTIONS

The data collected for each month was fitted to various standard probability distributions to find out which of the probability distributions gave the best fit. To decide on which probability distribution best fit the data, the method of mean square error was used.

#### 3.3.1 Mean Square Errors of the Distributions

The standard distribution with the smallest mean square error with respect to the data was considered the best candidate distribution for the particular month under consideration.

Due to the large number of distributions dealt with, the remaining tables and charts have been omitted here and rather presented in the appendix.

Tables 3.1 and 3.2 below show the mean square errors for various standard probability distributions for the months of January and February respectively. We find that the lognormal distribution is the best fitted distribution for the month of January since it has the smallest mean square error while in the month of February the following class of probability distributions all fit best: Exponential, Weibull, Gamma, Lognormal, and Geometric distributions.

### 3.3.1.1 Mean Square Errors of the Distributions for Jan and Feb

Table 3.1: Distribution Mean Square Error for Jan

DISTRIBUTION	MEAN SQUARE ERROR
<i>Exponential Distribution</i>	0.000003
<i>Weibull Distribution</i>	0.000003
<i>Gamma Distribution</i>	0.000003
<i>Lognormal Distribution</i>	2.54456E-06
<i>Beta Distribution</i>	2.005E+191
<i>Geometric Distribution</i>	0.000003

Table 3.2: Distribution Mean Square Error for Feb

DISTRIBUTION	MEAN SQUARE ERROR
<i>Exponential Distribution</i>	3.7037E-08
<i>Weibull Distribution</i>	3.7037E-08
<i>Gamma Distribution</i>	3.7037E-08
<i>Lognormal Distribution</i>	3.59166E-08
<i>Beta Distribution</i>	2.88371E+34
<i>Geometric Distribution</i>	3.7037E-08

### 3.3.1.2 Mean Square Errors of the Distributions for March and May

Tables 3.3 and 3.4 below show the mean square errors for various standard probability distributions for March and May respectively. The log normal distribution can be seen to be the best fitted distribution for the month of March and May since it has the smallest mean square error in both cases.

Table 3.3: Distribution Mean Square Error for March

DISTRIBUTION	MEAN SQUARE ERROR
<i>Exponential Distribution</i>	0.000003
<i>Weibull Distribution</i>	0.000003
<i>Gamma Distribution</i>	0.000003
<i>Lognormal Distribution</i>	5.61688E-06
<i>Beta Distribution</i>	1.79165E+60
<i>Geometric Distribution</i>	0.000003

Table 3.4: Distribution Mean Square Error for May

DISTRIBUTION	MEAN SQUARE ERROR
<i>Exponential Distribution</i>	2.5037E-05
<i>Weibull Distribution</i>	2.5037E-05
<i>Gamma Distribution</i>	2.5037E-05
<i>Lognormal Distribution</i>	4.98008E-07
<i>Beta Distribution</i>	1.29291E+91
<i>Geometric Distribution</i>	2.5037E-05

### 3.3.1.3 Mean Square Errors of the Distributions for April

The table 3.5 below shows the mean square errors for the various standard probability distributions under consideration for the month of April. The Weibull, exponential, gamma and the geometric distribution as can be seen below is the best fitted distribution for the month of April since it has the smallest mean square error.

Table 3.5: Distribution Mean Square Error for April

DISTRIBUTION	MEAN SQUARE ERROR
<i>Exponential Distribution</i>	3.7037E-08
<i>Weibull Distribution</i>	3.7037E-08
<i>Gamma Distribution</i>	3.7037E-08
<i>Lognormal Distribution</i>	7.10438E-06
<i>Beta Distribution</i>	4.0449E+248
<i>Geometric Distribution</i>	3.7037E-08

### 3.3.1.4 Mean Square Errors of the Distributions for June and Sept

Tables 3.6 and 3.7 below show the mean square errors for the various standard probability distributions under consideration for June and September respectively. The Exponential, Gamma, Weibull and Geometric distributions as can be seen below were the best fitted distributions for both the months of June and September since they all have the smallest mean square error

Table 3.6: Distribution Mean Square Error for June Table3.7: Distribution Mean Square Error for Sept

DISTRIBUTION	MEAN SQUARE ERROR	DISTRIBUTION	MEAN SQUARE ERROR
Exponential Distribution	3.7037E-08	Exponential Distribution	9.25926E-07
Weibull Distribution	3.7037E-08	Weibull Distribution	9.25926E-07
Gamma Distribution	3.7037E-08	Gamma Distribution	9.25926E-07
Lognormal Distribution	6.79067E-06	Lognormal Distribution	7.11702E-06
Beta Distribution	1.0862E+237	Beta Distribution	8.0732E+235
Geometric Distribution	3.7037E-08	Geometric Distribution	9.25926E-07

### 3.3.1.5 Mean Square Errors of the distributions for July

Table 3.8 below shows the mean square errors for the various standard probability distributions under consideration for July. The Exponential, Gamma, Weibull and Geometric distributions were the best fitted distributions for both the months of June and September since they each have the smallest mean square error.

Table 3.8: Distribution Mean Square Error for July

DISTRIBUTION	MEAN SQUARE ERROR
Exponential Distribution	1.33333E-06
Weibull Distribution	1.33333E-06
Gamma Distribution	1.33333E-06
Lognormal Distribution	7.5403E-06
Beta Distribution	1.2338E+233
Geometric Distribution	1.33333E-06

### 3.3.1.6 Mean Square Errors of the Distributions for August

Table 3.9 below shows the mean square errors for the various probability distributions under consideration for August best fitted distribution for the month of August since it has the smallest mean square error.

Table 3.9: Distribution Mean Square Error for August

<i>DISTRIBUTION</i>	<i>MEAN SQUARE ERROR</i>
<i>Exponential Distribution</i>	9.25926E-07
<i>Weibull Distribution</i>	9.25926E-07
<i>Gamma Distribution</i>	9.25926E-07
<i>Lognormal Distribution</i>	7.49618E-06
<i>Beta Distribution</i>	5.7038E+234
<i>Geometric Distribution</i>	9.25926E-07

### 3.3.1.7 Mean Square Errors of the Distributions for Oct and Nov

Tables 3.10 and 3.11 below show the mean square errors for various standard probability distributions for October and November respectively. The Exponential, Gamma, Weibull and Geometric distributions were the best fitted distributions for both the months of October and November since they all have the smallest mean square error.

Table 3.10: Distribution Mean Square Error for October

DISTRIBUTION	MEAN SQUARE ERROR
Exponential Distribution	3.7037E-08
Weibull Distribution	5.7201E-128
Gamma Distribution	3.7037E-08
Lognormal Distribution	6.37252E-07
Beta Distribution	1.13913E+81
Geometric Distribution	3.7037E-08

Table 3.11: Distribution Mean Square Error for

DISTRIBUTION	MEAN SQUARE ERROR
Exponential Distribution	3.33333E-07
Weibull Distribution	3.33333E-07
Gamma Distribution	3.33333E-07
Lognormal Distribution	7.33666E-06
Beta Distribution	1.8199E+191
Geometric Distribution	3.33333E-07

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### 3.3.1.8 Mean Square Errors of the Distributions for December

Table 3.12 below shows the mean square errors for the various standard probability distributions under consideration for December. The Exponential, Gamma, Weibull Lognormal and Geometric distributions were the best fitted distributions for the month of December since they all have the smallest mean square error.

Table 3.12: Distribution Mean Square Error for December

DISTRIBUTION	MEAN SQUARE ERROR
Exponential Distribution	2.37037E-06
Weibull Distribution	2.37037E-06
Gamma Distribution	2.37037E-06
Lognormal Distribution	4.14557E-06
Beta Distribution	1.0932E+178
Geometric Distribution	2.37037E-06

### 3.3.2 Monthly Probability Distributions

The randomly selected solar irradiation data was fitted into the various probability distributions and the resulting graphs of their probability distributions are presented for the various months of January march and May.

As discussed in section in section 3.3 the lognormal probability distribution can best fit the month of January march and May. Fig 3.16 below gives the graph of the lognormal distribution for the month of January march and May solar irradiation.

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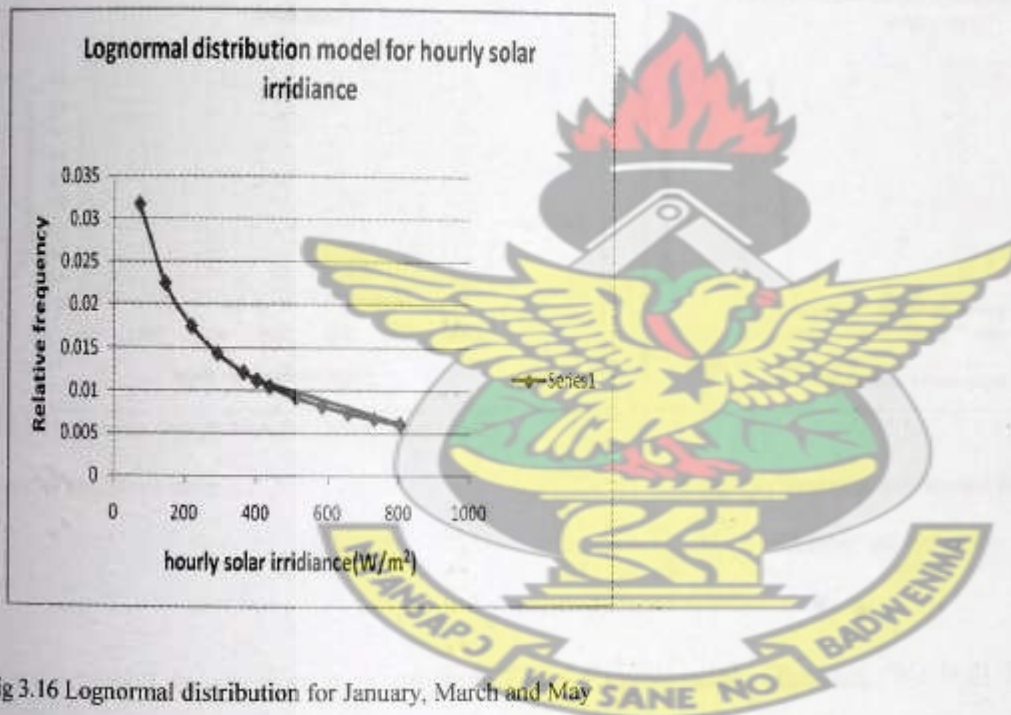


Fig 3.16 Lognormal distribution for January, March and May

As discussed in section in section 3.3 the Exponential, Gamma, Weibull Lognormal and Geometric distributions can best fit the month of April June September October and November . Fig 3.16 to fig 3.21 below gives the graph of the above mentioned probability distribution for the month of April, June, September, October and November

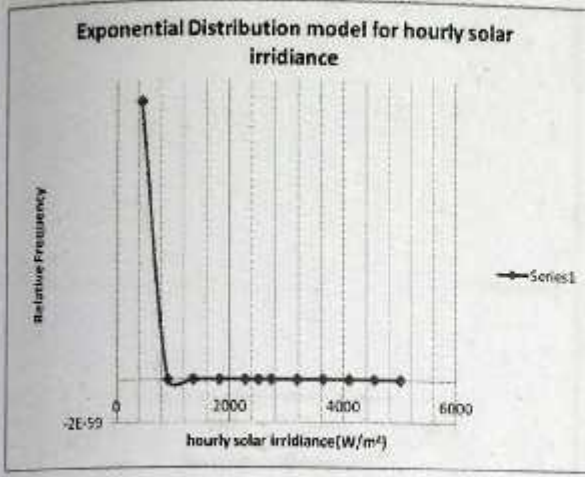


Fig 3.17 Exponential distribution for April, June, September, October and November

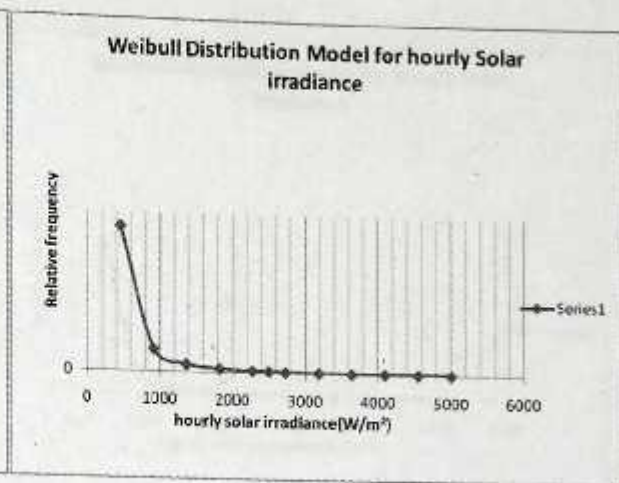


Fig 3.18 Weibull distribution for April, June, October and November

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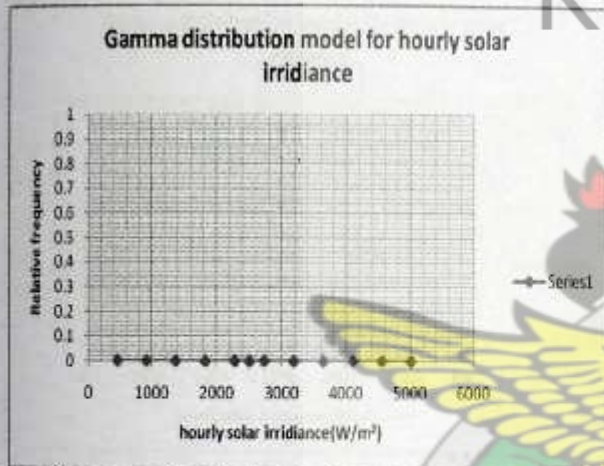


Fig 3.19 Gamma distribution for June, September, October and November

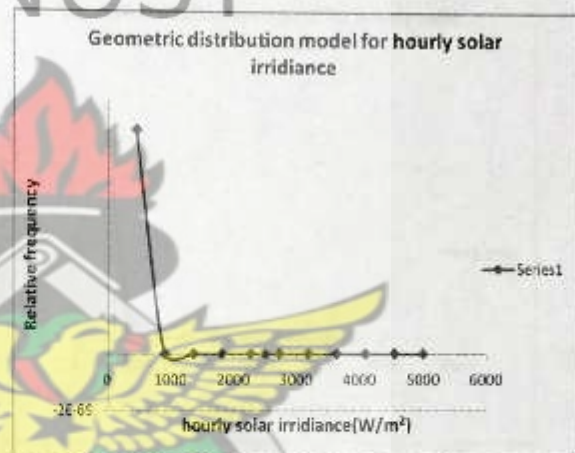


Fig 3.20 Geometric distribution for June, September, October and November

As discussed in section in section 3.3 the Exponential, Gamma, Weibull Lognormal and Geometric distributions can best fit the month of February, July and December. Fig 3.21 to fig 3.25 below gives the graph of the above mentioned probability distribution for the month of February, July and December

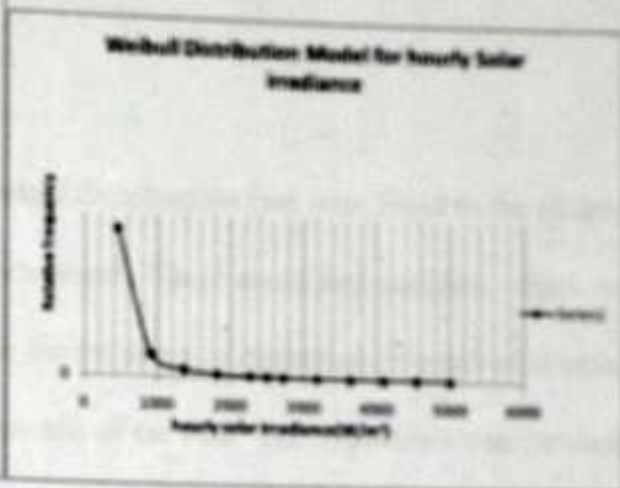
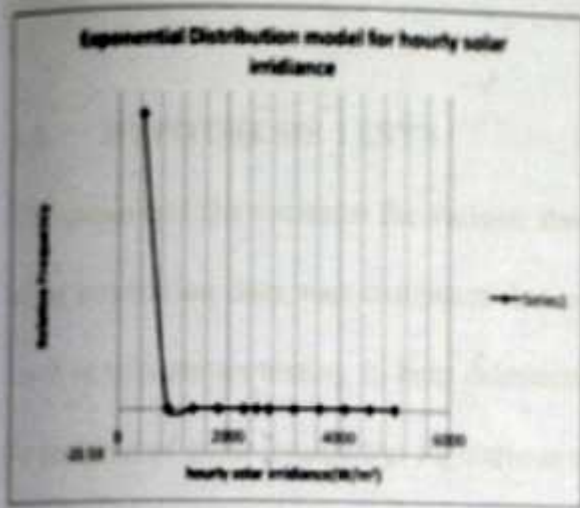


Fig 3.21 Exponential distribution for February, July and December, Fig 3.22 Weibull distribution for February, July and

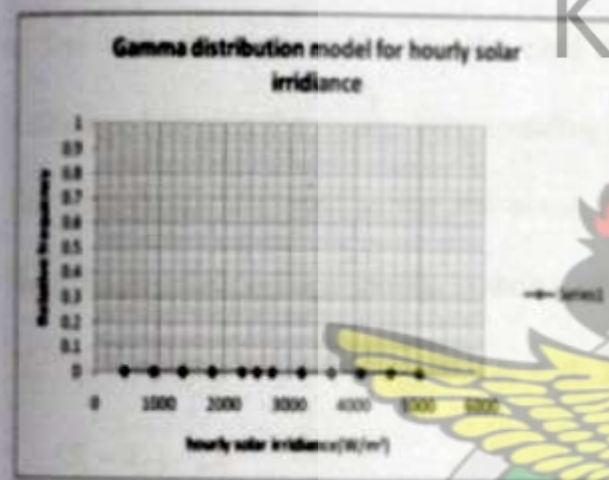


Fig 3.23 Gamma distribution for, February, July and December, Fig 3.24 Geometric distribution for, February, July and December,

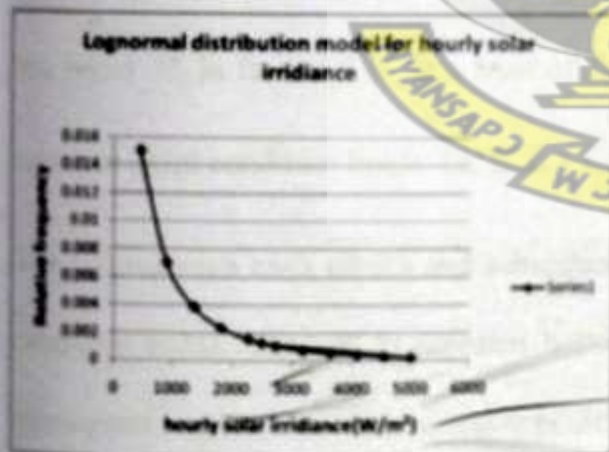


Fig 3.25 Geometric distribution for, February, July and December,

### 3.5 HYPOTHESIS TESTS

The mean and the variance the various standard distributions that were fitted to the observed solar irradiation data was computed for each month. These mean and variance values were used in hypothesis testing to help determine the existence of clustering in terms of similarity in patterns of solar irradiation for various months of the year. The hypothesis was conducted to test the difference between the population means of the various months of the year solar irradiation based on the sample selected. January being the first month of the year was used to test with the rest of the months until there was a rejection (existence of significant difference between the means corresponding to the months).

The hypothesis test used in the test about the difference between two populations means. Here the population corresponds to the month of the year. Testing:

$H_0 : \mu_1 = \mu_2$  or  $\mu_1 - \mu_2 = 0$        $H_a : \mu_1 \neq \mu_2$  With the test statistic:

$$z = \frac{(\bar{x}_1 - \bar{x}_2) - D_0}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

We reject  $H_0$  in favor of  $H_a$  at a level of significance  $\alpha=0.05$  if and only if the appropriate rejection point condition holds. I.e. if  $Z > Z_{\alpha/2}$  or  $Z < -Z_{\alpha/2}$

The test between each month and subsequent months is performed until the rejection point condition holds. The test is repeated between the last month used in the latter test and subsequent months until the rejection point condition holds again. This procedure grouped the months into clusters, shown in the tables below.

### 3.5.1 Testing of Hypothesis between Mean Solar Irradiation of January and February

The month of January was used to test against the month of February, it was noticed that there is no significant difference between mean solar output for January and February. Thus January and February belong to the same cluster. The Table 3.20 below shows the results of the test.

Table 3.20: Test of Hypothesis between Mean Solar Irradiation of January and February

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	FEB	1500	675	50625

#### RESULTS

Z =	-0.448966081	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-592.2117801
Decision:	Accept $H_0$ ( since $Z > -1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.2 Testing of Hypothesis between Mean Solar Irradiation of January and March

The month of January was tested with March and it was found that there is no significant difference between the two months which indicate that January and March can also be in the same cluster. The table 3.21 below shows the result of the test.

Table3.21: Test of Hypothesis between Mean Solar Irradiation of January and March

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	MARCH	1500	37.14205451	200752.5614

**RESULTS**

Z =	0.008801078	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	45.64616544
Decision:	Accept Ho ( since $Z < 1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.3 Testing of Hypothesis between Mean Solar Irradiation of January and April

The month of January was tested with April and it was found that there is no significant difference between the two months which indicate that January and April can also be in the same cluster. The table 3.22 below shows the result of the test.

Table3.22: Test of Hypothesis between Mean Solar Irradiation of January and April

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	APRIL	1500	36.42145334	148625.8932

**RESULTS**

Z =	0.012233982	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	46.36670661
Decision:	Accept Ho ( since $Z < 1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.4 Testing of Hypothesis between Mean Solar Irradiation of January and May

The month of January was tested with May and it was found that there is no significant difference between the two months which indicate that January and May can also be in the same cluster. The table 3.23 below shows the result of the test.

Table 3.23: Test of Hypothesis between Mean Solar Irradiation of January and May

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82,78821995	6853.889362
Sample 2	MAY	1500	84,30471912	169874.2951

#### RESULTS

Z =	0.011044822	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	48.48350083
Decision:	Accept Ho ( since $Z < 1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.5 Testing of Hypothesis between Mean Solar Irradiation of January and June

The month of January was tested with June and it was found that there is no significant difference between the two months which indicate that January and June can also be in the same cluster. The table 3.24 below shows the result of the test.

Table 3.24: Test of Hypothesis between Mean Solar Irradiation of January and June

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	JUNE	1500	404	1616

**RESULTS**

Z =	-1.766656266	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-321.2117801
Decision:	Accept Ho ( since Z > -1.96)		
Conclusion:	There is no difference between the population means		

### 3.5.6 Testing of Hypothesis between Mean Solar Irradiation of January and July

The month of January was tested with July and it was found that there is no significant difference between the two months which indicate that January and July can also be in the same cluster. The table 3.25 below shows the result of the test.

Table 3.25: Test of Hypothesis between Mean Solar Irradiation of January and February

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	JULY	1500	308	616

**RESULTS**

Z =	-1.267513639	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-225.2117801
Decision:	Accept Ho ( since Z > -1.96)		
Conclusion:	There is no difference between the population means		

### 3.5.7 Testing of Hypothesis between Mean Solar Irradiation of January and August

The month of January was tested with August and it was found that there is no significant difference between the two months which indicate that January and August can also be in the same cluster. The table 3.26 below shows the result of the test.

Table 3.26: Test of Hypothesis between Mean Solar Irradiation of January and August

	Month	sample size (n)	Mean ( $\bar{x}$ )	Variance ( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	AUG	1500	312	936

#### RESULTS

Z =	-1.283314174	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-229.2117801
Decision:	Accept Ho ( since $Z > -1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.8 Testing of Hypothesis between Mean Solar Irradiation of January and September

The month of January was tested with September and it was found that there is no significant difference between the two months which indicate that January and September can also be in the same cluster. The table 3.27 below shows the result of the test.

Table3.27: Test of Hypothesis between Mean Solar Irradiation of January and September

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	SEPT	1500	306	612

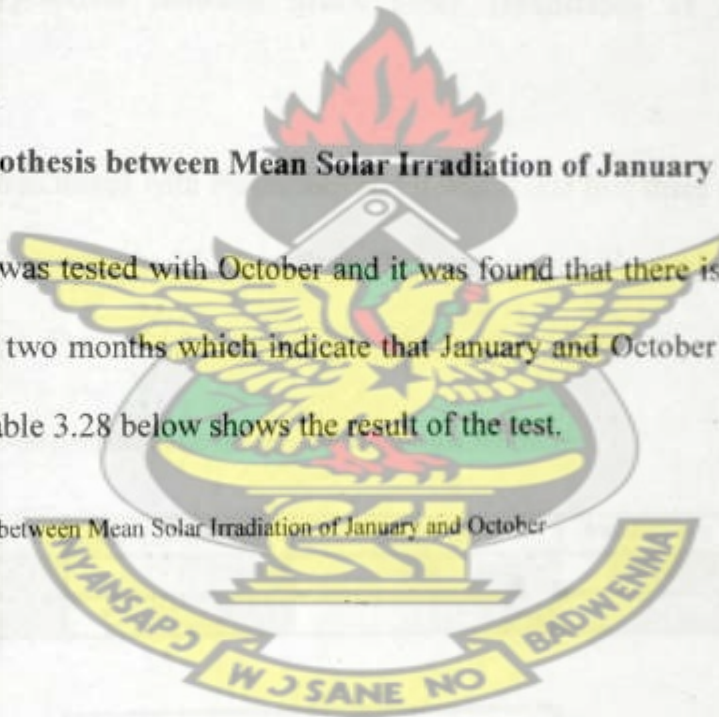
**RESULTS**

Z =	-1.256322605	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-223.2117801
Decision:	Accept Ho ( since Z > -1.96)		
Conclusion:	There is no difference between the population means		

### 3.5.9 Testing of Hypothesis between Mean Solar Irradiation of January and October

The month of January was tested with October and it was found that there is no significant difference between the two months which indicate that January and October can also be in the same cluster. The table 3.28 below shows the result of the test.

Table3.28: Test of Hypothesis between Mean Solar Irradiation of January and October



	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	OCT	1500	241	868

### RESULTS

Z =	-0.886935958	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-158.2117801
Decision:	Accept Ho ( since Z > -1.96)		
Conclusion:	There is no difference between the population means		

# KNUST

## 3.5.10 Testing of Hypothesis between Mean Solar Irradiation of January and November

The month of January was tested with November and it was found that there is a significant difference between the two months which indicate that January and November cannot be in the same the same cluster. The table 3.29 below shows the result of the test.

Table 3.29: Test of Hypothesis between Mean Solar Irradiation of January and November

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	JAN	1500	82.78821995	6853.889362
Sample 2	NOV	1500	494	988

### RESULTS

Z =	-2.299895326	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-411.2117801
Decision:	Reject Ho in favour of Ha ( since Z < -1.96)		
Conclusion:	There is a difference between the population means		

### 3.5.11 Testing of Hypothesis between Mean Solar Irradiation of November and December

Since the test for January and November shows a significant difference, we will now test November with December to check the result. It was found that there is no significant difference between the solar irradiation for November and December and these two months can also form another cluster. Table 3.31 below shows the test results.

Table 3.30: Test of Hypothesis between Mean Solar Irradiation of November and December

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	NOV	1500	494	988
Sample 2	DEC	1500	512	2048

#### RESULTS

Z =	-0.30658722	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	-18
Decision:	Accept Ho ( since $Z > -1.96$ )		
Conclusion:	There is no difference between the population means		

### 3.5.12 Testing of Hypothesis between Mean Solar Irradiation of November and January

The month of November was tested with January and it was found that there is a significant difference between the two months which confirms the fact that January and November cannot be in the same cluster. Table 3.31 below shows the result of t

Table 3.31: Test of Hypothesis between Mean Solar Irradiation of November and January

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	NOV	1500	494	988
Sample 2	JAN	1500	82.78821995	6853.889362

### RESULTS

Z =	2.299895326	Hypothesized difference (Do)	0
$ Z_{\alpha/2}  =$	1.96	Mean difference	411.2117801
<b>Decision:</b>	Reject $H_0$ in favour of $H_a$ ( since $Z > 1.96$ )		
<b>Conclusion:</b>	There is a difference between the population means		

#### 3.5.4 Testing of Hypothesis between Mean Solar Irradiation of the Clusters.

The result of the series of tests indicate that there are two clusters in the year in respect of similar patterns of solar irradiation, namely, one from January to October and the other from November to December.

Further hypothesis testing analysis was conducted on these two clusters of the year and it was also found that there is a significant difference between the mean outputs of solar irradiation for the two clusters of the year. The result of the cluster analysis is shown in Table 3.33 below.

Table 3.32: Test of Hypothesis between Mean Solar Irradiation of Jan-Oct and Nov-Dec

**INPUTS**

**Level of Significance** 0.05

	Month	sample size (n)	Mean( $\bar{x}$ )	Variance( $s^2$ )
Sample 1	jan-oct	1500	428	856
Sample 2	nov-dec	1500	326	652

**RESULTS**

<b>Z =</b>	3.671313969	<b>Hypothesized difference (Do)</b>	0
<b> Z<sub>α/2</sub>  =</b>	1.96	<b>Mean difference</b>	102
<b>Decision:</b>	Reject Ho in favour of Ha ( since Z > 1.96)		
<b>Conclusion:</b>	There is a difference between the population means		

### 3.6 BAYESIAN ANALYSIS

In the Bayesian analysis we use the random variable  $X$  which represents the event of having a high or low sunshine.

$$\text{Thus } X = \begin{cases} 0 & \text{low sunshine} \\ 1 & \text{high sunshine} \end{cases}$$

The values of  $X$  are generated based on a given threshold below which we describe the sunshine level as being low and as high otherwise. The threshold value used in this analysis is  $120kWhm^{-2}$ . A random sample of 1500 sunshine hours were selected and based on the above mentioned threshold the monthly average number of high and low sunshine levels was computed.

The beta distribution was used as prior distribution and the Bernoulli probability distribution as a likelihood function. The posterior probability distribution is derived in section 2.46.

Bayesian analysis for the various months of the year was conducted to estimate the posterior expectation and variance of the beta distribution for number of iterations until these iterations are converging to a certain value given a tolerance level. The analysis started with initial parameter values of alpha and beta to compute the prior means and prior variance. The posterior parameters are then computed based on the sample size and the number of high sunshine hours. The posterior parameters were then used as the prior parameters to compute the mean and variance. The procedure is repeated for twenty iterations for all the months of the year.

### 3.6.1 Beta Distribution

As discussed in the section 2.4.7 a random variable  $X$  is said to have beta distribution with parameters  $\alpha > 0$   $\beta > 0$  if the density of  $X$  is

$$f(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \text{ for } 0 \leq x \leq 1 \text{ and } f(x) = 0 \text{ outside the interval } [0,1]$$

The mean and the variance of the beta distribution is given below

$$\text{Mean denoted } E(X) = \mu = \frac{\alpha}{\alpha + \beta} \text{ and variance denoted } \text{Var}(X) = \frac{\alpha\beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$$

### 3.6.2 Bayesian Methodology

The prior beta distribution is used to together with the likelihood function to obtain the posterior distribution. The beta distribution being a conjugate prior always has a posterior beta distribution with parameters

$$\alpha^* = \sum x + \alpha$$

$$\beta^* = n - \sum x + \beta$$

Where  $n$  the sample is size and  $\sum x$  is the total number of sunshine hours in the sampled data under the threshold of  $120\text{kwh/m}^2$ . The alpha and beta values were computed and used as the initial guess. The posterior parameters were computed using the formula above. The posterior alpha and beta values were used as the prior alpha and beta values to compute the posterior parameters for that iteration.

The posterior mean and variance for the posterior probability distribution was computed as below for the number of iterations.

$$E(X) = \frac{\alpha^*}{\alpha^* + \beta^*}$$

$$\text{Var}(X) = \frac{\alpha^* \beta^*}{(\alpha^* + \beta^*)^2 (\alpha^* + \beta^* + 1)}$$

### 3.6.3 Presentation of Results

A sample of size 1500 was selected from each month of the year. Based on the threshold value the total number of sunshine hours is calculated to help in the computation of posterior beta distribution parameters.

### 3.6.4 Bayesian Results for January

The sample selected from the month of January was analyzed using Bayesian approach.

Twenty iterations were performed and the results are shown on the table 3.34 below

The results in table 3.34 indicate that the prior and posterior means converges to 0.87000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 4.0E-06 at 0.000001 tolerance level. The analysis shows that the month of January has a high possibility (.87000) of sunshine.

Table 3.34 Bayesian analysis for January

ITERATION S	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	27	7	0.794117647	0.00467128	1332	202	0.868318123	7.44897E-05
2	1332	202	0.868318123	7.44897E-05	2637	397	0.869149637	3.74723E-05
3	2637	397	0.869149637	3.74723E-05	3942	592	0.869430966	2.50321E-05
4	3942	592	0.869430966	2.50321E-05	5247	787	0.869572423	1.87931E-05
5	5247	787	0.869572423	1.87931E-05	6552	982	0.869657552	1.50436E-05
6	6552	982	0.869657552	1.50436E-05	7857	1177	0.869714412	1.25414E-05
7	7857	1177	0.869714412	1.25414E-05	9162	1372	0.869755079	1.07528E-05
8	9162	1372	0.869755079	1.07528E-05	10467	1567	0.869785607	9.41077E-06
9	10467	1567	0.869785607	9.41077E-06	11772	1762	0.869809369	8.36653E-06
10	11772	1762	0.869809369	8.36653E-06	13077	1957	0.869828389	7.53089E-06
11	13077	1957	0.869828389	7.53089E-06	14382	2152	0.869843958	6.84702E-06
12	14382	2152	0.869843958	6.84702E-06	15687	2347	0.869856937	6.27701E-06
13	15687	2347	0.869856937	6.27701E-06	16992	2542	0.869867923	5.79461E-06
14	16992	2542	0.869867923	5.79461E-06	18297	2737	0.869877341	5.38107E-06
15	18297	2737	0.869877341	5.38107E-06	19602	2932	0.869885506	5.02262E-06
16	19602	2932	0.869885506	5.02262E-06	20907	3127	0.869892652	4.70894E-06
17	20907	3127	0.869892652	4.70894E-06	22212	3322	0.869898958	4.43214E-06
18	22212	3322	0.869898958	4.43214E-06	23517	3517	0.869904565	4.18608E-06
19	23517	3517	0.869904565	4.18608E-06	24822	3712	0.869909582	3.9659E-06
20	24822	3712	0.869909582	3.9659E-06	26127	3907	0.869914097	3.76772E-06

### 3.6.5 Bayesian Analysis for February

The sample selected from the month of February was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.35. The results in the table 3.35 below indicate that the prior and posterior means converge to 0.83000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of February has a high possibility (0.82000) of sunshine.

Table 3.35: Bayesian analysis for February

iterations	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	1	6	0.142857143	0.015306122	1046	461	0.694094227	0.000140801
2	1046	461	0.694094227	0.000140801	2091	671	0.757060101	6.65654E-05
3	2091	671	0.757060101	6.65654E-05	3136	881	0.780682101	4.26126E-05
4	3136	881	0.780682101	4.26126E-05	4181	1091	0.793057663	3.11241E-05
5	4181	1091	0.793057663	3.11241E-05	5226	1301	0.800674123	2.44478E-05
6	5226	1301	0.800674123	2.44478E-05	6271	1511	0.805833976	2.01035E-05
7	6271	1511	0.805833976	2.01035E-05	7316	1721	0.809560695	1.70582E-05
8	7316	1721	0.809560695	1.70582E-05	8361	1931	0.812378546	1.48081E-05
9	8361	1931	0.812378546	1.48081E-05	9406	2141	0.814583875	1.30791E-05
10	9406	2141	0.814583875	1.30791E-05	10451	2351	0.816356819	1.17096E-05
11	10451	2351	0.816356819	1.17096E-05	11496	2561	0.817813189	1.05986E-05
12	11496	2561	0.817813189	1.05986E-05	12541	2771	0.819030825	9.67931E-06
13	12541	2771	0.819030825	9.67931E-06	13586	2981	0.820063983	8.90627E-06
14	13586	2981	0.820063983	8.90627E-06	14631	3191	0.820951633	8.24721E-06
15	14631	3191	0.820951633	8.24721E-06	15676	3401	0.821722493	7.67872E-06
16	15676	3401	0.821722493	7.67872E-06	16721	3611	0.82239819	7.18337E-06
17	16721	3611	0.82239819	7.18337E-06	17766	3821	0.822995321	6.74792E-06
18	17766	3821	0.822995321	6.74792E-06	18811	4031	0.823526837	6.36214E-06
19	18811	4031	0.823526837	6.36214E-06	19856	4241	0.824002988	6.01801E-06
20	19856	4241	0.824002988	6.01801E-06	20901	4451	0.824431997	5.70914E-06

### 3.6.6 Bayesian Analysis for March

Table 3.36: Bayesian analysis for March

iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	13	1	0.928571429	0.004421769	1363	151	0.900264201	5.92664E-05
2	1363	151	0.900264201	5.92664E-05	2713	301	0.900132714	2.98155E-05
3	2713	301	0.900132714	2.98155E-05	4063	451	0.900088613	1.99179E-05
4	4063	451	0.900088613	1.99179E-05	5413	601	0.900066511	1.49537E-05
5	5413	601	0.900066511	1.49537E-05	6763	751	0.900053234	1.19704E-05
6	6763	751	0.900053234	1.19704E-05	8113	901	0.900044375	9.97942E-06
7	8113	901	0.900044375	9.97942E-06	9463	1051	0.900038045	8.55631E-06
8	9463	1051	0.900038045	8.55631E-06	10813	1201	0.900033294	7.48842E-06
9	10813	1201	0.900033294	7.48842E-06	12163	1351	0.900029599	6.65752E-06
10	12163	1351	0.900029599	6.65752E-06	13513	1501	0.900026642	5.99259E-06
11	13513	1501	0.900026642	5.99259E-06	14863	1651	0.900024222	5.44842E-06
12	14863	1651	0.900024222	5.44842E-06	16213	1801	0.900022205	4.99485E-06
13	16213	1801	0.900022205	4.99485E-06	17563	1951	0.900020398	4.6111E-06
14	17563	1951	0.900020498	4.6111E-06	18913	2101	0.900019035	4.28193E-06
15	18913	2101	0.900019035	4.28193E-06	20263	2251	0.900017767	3.99671E-06
16	20263	2251	0.900017767	3.99671E-06	21613	2401	0.900016657	3.74711E-06
17	21613	2401	0.900016657	3.74711E-06	22963	2551	0.900015678	3.52685E-06
18	22963	2551	0.900015678	3.52685E-06	24313	2701	0.900014807	3.33104E-06
19	24313	2701	0.900014807	3.33104E-06	25663	2851	0.900014028	3.15584E-06
20	25663	2851	0.900014028	3.15584E-06	27013	3001	0.900013327	2.99815E-06

The sample selected from the month of March was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.36. The results in the table 3.35 indicate that the prior and posterior means converges to 0.9000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 3.0E-06 at 0.000001 tolerance level. The analysis shows that the month of February has a high possibility (0.9000) of sunshine.

### 3.6.7 Bayesian Analysis for April

The sample selected from the month of April was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.37. The results in the table 3.37 indicate that the prior and posterior means converge to 0.86000 tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of April has a high possibility (0.86000) of sunshine.

Table 3.37: Bayesian analysis for April

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	26	16	0.619047619	0.005484364	1316	226	0.853437095	8.10643E-05
2	1316	226	0.853437095	8.10643E-05	2606	436	0.856673241	4.03497E-05
3	2606	436	0.856673241	4.03497E-05	3896	646	0.857771907	2.68543E-05
4	3896	646	0.857771907	2.68543E-05	5186	856	0.858325058	2.0123E-05
5	5186	856	0.858325058	2.0123E-05	6476	1066	0.858658161	1.60897E-05
6	6476	1066	0.858658161	1.60897E-05	7766	1276	0.858880779	1.34031E-05
7	7766	1276	0.858880779	1.34031E-05	9056	1486	0.85904003	1.14854E-05
8	9056	1486	0.85904003	1.14854E-05	10346	1696	0.859159608	1.00477E-05
9	10346	1696	0.859159608	1.00477E-05	11636	1906	0.859252695	8.92989E-06
10	11636	1906	0.859252695	8.92989E-06	12926	2116	0.859327217	8.03589E-06
11	12926	2116	0.859327217	8.03589E-06	14216	2326	0.859388224	7.30461E-06
12	14216	2326	0.859388224	7.30461E-06	15506	2536	0.859439087	6.69531E-06
13	15506	2536	0.859439087	6.69531E-06	16796	2746	0.859482141	6.17984E-06
14	16796	2746	0.859482141	6.17984E-06	18086	2956	0.859519057	5.73806E-06
15	18086	2956	0.859519057	5.73806E-06	19376	3166	0.85955106	5.35523E-06
16	19376	3166	0.85955106	5.35523E-06	20666	3376	0.85957907	5.02029E-06
17	20666	3376	0.85957907	5.02029E-06	21956	3586	0.85960379	4.72478E-06
18	21956	3586	0.85960379	4.72478E-06	23246	3796	0.859625767	4.46213E-06
19	23246	3796	0.859625767	4.46213E-06	24536	4006	0.859645435	4.22714E-06
20	24536	4006	0.859645435	4.22714E-06	25826	4216	0.859663138	4.01566E-06

### 3.6.8 Bayesian Analysis for May

The sample selected from the month of February was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.38. The results in the table 3.38 indicate that the prior and posterior means converge to 0.92000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 2.0E-06 at 0.000001 tolerance level. The analysis shows that the month of February has a very high possibility (0.92000) of sunshine.

Table 3.38: Bayesian analysis for May

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	15	1	0.9375	0.003446691	1400	116	0.92348285	4.65803E-05
2	1400	116	0.92348285	4.65803E-05	2785	231	0.923408488	2.34422E-05
3	2785	231	0.923408488	2.34422E-05	4170	346	0.923383525	1.56623E-05
4	4170	346	0.923383525	1.56623E-05	5555	461	0.923371011	1.17595E-05
5	5555	461	0.923371011	1.17595E-05	6940	576	0.923363491	9.41378E-06
6	6940	576	0.923363491	9.41378E-06	8325	691	0.923358474	7.84824E-06
7	8325	691	0.923358474	7.84824E-06	9710	806	0.923354888	6.72917E-06
8	9710	806	0.923354888	6.72917E-06	11095	921	0.923352197	5.8894E-06
9	11095	921	0.923352197	5.8894E-06	12480	1036	0.923350104	5.23598E-06
10	12480	1036	0.923350104	5.23598E-06	13865	1151	0.923348428	4.71307E-06
11	13865	1151	0.923348428	4.71307E-06	15250	1266	0.923347057	4.28512E-06
12	15250	1266	0.923347057	4.28512E-06	16635	1381	0.923345915	3.92841E-06
13	16635	1381	0.923345915	3.92841E-06	18020	1496	0.923344948	3.62653E-06
14	18020	1496	0.923344948	3.62653E-06	19405	1611	0.923344119	3.36774E-06
15	19405	1611	0.923344119	3.36774E-06	20790	1726	0.9233434	3.14342E-06
16	20790	1726	0.9233434	3.14342E-06	22175	1841	0.923342771	2.94712E-06
17	22175	1841	0.923342771	2.94712E-06	23560	1956	0.923342217	2.77389E-06
18	23560	1956	0.923342217	2.77389E-06	24945	2071	0.923341723	2.6199E-06
19	24945	2071	0.923341723	2.6199E-06	26330	2186	0.923341282	2.4821E-06
20	26330	2186	0.923341282	2.4821E-06	27715	2301	0.923340885	2.35808E-06

### 3.6.9 Bayesian Analysis for June

The sample selected from the month of June was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.39. The results in table 3.39 indicates that the prior and posterior means converge to 0.77000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of February has a high possibility (0.77000) of sunshine

Table 3.39: Bayesian analysis for June

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	21	18	0.538461538	0.006213018	1075	464	0.698505523	0.00013675
2	1075	464	0.698505523	0.00013675	2129	910	0.700559395	6.90052E-05
3	2129	910	0.700559395	6.90052E-05	3183	1356	0.701255783	4.61445E-05
4	3183	1356	0.701255783	4.61445E-05	4237	1802	0.701606226	3.46614E-05
5	4237	1802	0.701606226	3.46614E-05	5291	2248	0.701817217	2.77546E-05
6	5291	2248	0.701817217	2.77546E-05	6345	2694	0.701958181	2.3143E-05
7	6345	2694	0.701958181	2.3143E-05	7399	3140	0.702059019	1.98456E-05
8	7399	3140	0.702059019	1.98456E-05	8453	3586	0.702134729	1.73705E-05
9	10346	1696	0.859159608	1.00477E-05	11400	2142	0.841825432	9.83204E-06
10	11400	2142	0.841825432	9.83204E-06	12454	2588	0.827948411	9.46951E-06
11	12454	2588	0.827948411	9.46951E-06	13508	3034	0.816588079	9.0535E-06
12	13508	3034	0.816588079	9.0535E-06	14562	3480	0.807116728	8.62824E-06
13	14562	3480	0.807116728	8.62824E-06	15616	3926	0.799099376	8.21468E-06
14	15616	3926	0.799099376	8.21468E-06	16670	4372	0.792225074	7.82229E-06
15	16670	4372	0.792225074	7.82229E-06	17724	4818	0.786265637	7.45473E-06
16	17724	4818	0.786265637	7.45473E-06	18778	5264	0.781049829	7.11271E-06
17	18778	5264	0.781049829	7.11271E-06	19832	5710	0.776446637	6.79549E-06
18	19832	5710	0.776446637	6.79549E-06	20886	6156	0.772354116	6.50162E-06
19	20886	6156	0.772354116	6.50162E-06	21940	6602	0.768691753	6.22936E-06
20	21940	6602	0.768691753	6.22936E-06	22994	7048	0.765395114	5.97695E-06

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### 3.6.10 Bayesian Analysis for July

Table 3.40: Bayesian analysis for July

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	28	14	0.666666667	0.005167959	1082	427	0.717031146	0.000134369
2	1082	427	0.717031146	0.000134369	2136	840	0.717741935	6.80512E-05
3	2136	840	0.717741935	6.80512E-05	3190	1253	0.717983345	4.55633E-05
4	3190	1253	0.717983345	4.55633E-05	4244	1656	0.718104907	3.42464E-05
5	4244	1656	0.718104907	3.42464E-05	5298	2079	0.718178121	2.74327E-05
6	5298	2079	0.718178121	2.74327E-05	6352	2492	0.718227047	2.28804E-05
7	6352	2492	0.718227047	2.28804E-05	7406	2905	0.71826205	1.96239E-05
8	7406	2905	0.71826205	1.96239E-05	8460	3318	0.718288334	1.71789E-05
9	8460	3318	0.718288334	1.71789E-05	9514	3731	0.718308796	1.52757E-05
10	9514	3731	0.718308796	1.52757E-05	10568	4144	0.718325177	1.37521E-05
11	10568	4144	0.718325177	1.37521E-05	11622	4557	0.718338587	1.25048E-05
12	11622	4557	0.718338587	1.25048E-05	12676	4970	0.718349768	1.1465E-05
13	12676	4970	0.718349768	1.1465E-05	13730	5383	0.718359232	1.05849E-05
14	13730	5383	0.718359232	1.05849E-05	14784	5796	0.718367347	9.83022E-06
15	14784	5796	0.718367347	9.83022E-06	15838	6209	0.718374382	9.17601E-06
16	15838	6209	0.718374382	9.17601E-06	16892	6622	0.718380539	8.60344E-06
17	16892	6622	0.718380539	8.60344E-06	17946	7035	0.718385973	8.09813E-06
18	17946	7035	0.718385973	8.09813E-06	19000	7448	0.718390805	7.64889E-06
19	19000	7448	0.718390805	7.64889E-06	20054	7861	0.718395128	7.24687E-06
20	20054	7861	0.718395128	7.24687E-06	21108	8274	0.71839902	6.885E-06

The sample selected from the month of July was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.40. The results in the table 3.40 below indicates that the prior means and posterior mean converges to 0.72000 tolerance level of 0.0001 while the prior and posterior variances are converging to 7.0E-06 at 0.000001 tolerance level. The analysis shows that the month of July has a greater possibility (0.72000) of sunshine

### 3.6.11 Bayesian Analysis for August

Table 3.41: Bayesian analysis for August

iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	26	16	0.619047519	0.005484354	1277	265	0.828145266	9.22363E-05
2	1277	265	0.828145266	9.22363E-05	2528	514	0.831032216	4.61445E-05
3	2528	514	0.831032216	4.61445E-05	3779	763	0.832012329	3.07655E-05
4	3779	763	0.832012329	3.07655E-05	5030	1012	0.832505793	2.30746E-05
5	5030	1012	0.832505793	2.30746E-05	6281	1261	0.83280297	1.84598E-05
6	6281	1261	0.83280297	1.84598E-05	7532	1510	0.833001548	1.53832E-05
7	7532	1510	0.833001548	1.53832E-05	8783	1759	0.833143616	1.31856E-05
8	8783	1759	0.833143616	1.31856E-05	10034	2008	0.833250291	1.15373E-05
9	10034	2008	0.833250291	1.15373E-05	11285	2257	0.833333333	1.02554E-05
10	11636	1906	0.859252695	8.92989E-06	12887	2155	0.856734477	8.15931E-06
11	12887	2155	0.856734477	8.15931E-06	14138	2404	0.854672954	7.50814E-06
12	14138	2404	0.854672954	7.50814E-06	15389	2653	0.852954218	6.95136E-06
13	15389	2653	0.852954218	6.95136E-06	16640	2902	0.851499335	6.47026E-06
14	16640	2902	0.851499335	6.47026E-06	17891	3151	0.850251877	6.05064E-06
15	17891	3151	0.850251877	6.05064E-06	19142	3400	0.849170437	5.68159E-06
16	19142	3400	0.849170437	5.68159E-06	20393	3649	0.848223941	5.35458E-06
17	20666	20393	3649	6.08838E-06	21917	20642	0.514979205	5.86879E-06
18	21917	20642	0.514979205	5.86879E-06	23168	20891	0.52584035	5.65893E-06
19	23168	20891	0.52584035	5.65893E-06	24419	21140	0.535986303	5.45885E-06
20	24419	21140	0.535986303	5.45885E-06	25670	21389	0.545485454	5.2684E-06

The sample selected from the month of August was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown in table 3.41. The results in table 3.41 indicates that the prior and posterior means converge to 0.55000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of August has a low possibility (0.55000) of sunshine

### 3.6.12 Bayesian Analysis for September

The sample selected from the month of September was analyzed using Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.42. The results in the table 3.42 below indicate that the prior and posterior means converge to 0.79000 with tolerance level of 0.0001 while the prior and posterior variances are converge to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of September has a high possibility (0.79000) of sunshine

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Table 3.42: Bayesian analysis for September

iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	29	13	0.690476	0.004970205	1207	335	0.782749676	0.000110209
2	1207	335	0.78275	0.000110209	2385	657	0.784023669	5.56459E-05
3	2385	657	0.784024	5.56459E-05	3563	979	0.784456187	3.72187E-05
4	3563	979	0.784456	3.72187E-05	4741	1301	0.784673949	2.79597E-05
5	4741	1301	0.784674	2.79597E-05	5919	1623	0.784805091	2.23898E-05
6	5919	1623	0.784805	2.23898E-05	7097	1945	0.784892723	1.86704E-05
7	7097	1945	0.784893	1.86704E-05	8275	2267	0.784955416	1.60107E-05
8	8275	2267	0.784955	1.60107E-05	9453	2589	0.785002491	1.40142E-05
9	9453	2589	0.785002	1.40142E-05	10631	2911	0.785039137	1.24605E-05
10	10631	2911	0.785039	1.24605E-05	11809	3233	0.785068475	1.12169E-05
11	11809	3233	0.785068	1.12169E-05	12987	3555	0.785092492	1.0199E-05
12	12987	3555	0.785092	1.0199E-05	14165	3877	0.785112515	9.35049E-06
13	14165	3877	0.785113	9.35049E-06	15343	4199	0.785129465	8.63231E-06
14	15343	4199	0.785129	8.63231E-06	16521	4521	0.785143998	8.01658E-06
15	16521	4521	0.785144	8.01658E-06	17699	4843	0.785156597	7.48284E-06
16	17699	4843	0.785157	7.48284E-06	18877	5165	0.785167623	7.01574E-06
17	18877	5165	0.785168	7.01574E-06	20055	5487	0.785177355	6.60353E-06
18	20055	5487	0.785177	6.60353E-06	21233	5809	0.785186007	6.23706E-06
19	21233	5809	0.785186	6.23706E-06	22411	6131	0.78519375	5.90914E-06
20	22411	6131	0.785194	5.90914E-06	23589	6453	0.785200719	5.61397E-06

### 3.6.13 Bayesian Analysis for October

A sample selected from the data for October was analyzed using the Bayesian approach. Iterations performed and the results (see table 3.43) indicate the prior and posterior means converge to 0.7200 at tolerance level of 0.0001 while the prior and posterior variances converge to 7.0E-06 at 0.000001 tolerance level. Thus the month of October has a high possibility of sunshine.

Table 3.43: Bayesian analysis for October

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	29	9	0.763157895	0.004634562	1104	434	0.717815345	0.000131616
2	1104	434	0.717815345	0.000131616	2179	859	0.71724819	6.67335E-05
3	2179	859	0.71724819	6.67335E-05	3254	1284	0.717055972	4.46985E-05
4	3254	1284	0.717055972	4.46985E-05	4329	1709	0.716959258	3.3603E-05
5	4329	1709	0.716959258	3.3603E-05	5404	2134	0.716901035	2.69205E-05
6	5404	2134	0.716901035	2.69205E-05	6479	2559	0.716862138	2.2455E-05
7	6479	2559	0.716862138	2.2455E-05	7554	2984	0.716834314	1.92602E-05
8	7554	2984	0.716834314	1.92602E-05	8629	3409	0.716813424	1.68612E-05
9	8629	3409	0.716813424	1.68612E-05	9704	3834	0.716797164	1.49936E-05
10	9704	3834	0.716797164	1.49936E-05	10779	4259	0.716784147	1.34985E-05
11	10779	4259	0.716784147	1.34985E-05	11854	4684	0.716773491	1.22746E-05
12	11854	4684	0.716773491	1.22746E-05	12929	5109	0.716764608	1.12541E-05
13	12929	5109	0.716764608	1.12541E-05	14004	5534	0.716757089	1.03903E-05
14	14004	5534	0.716757089	1.03903E-05	15079	5959	0.716750642	9.64966E-06
15	15079	5959	0.716750642	9.64966E-06	16154	6384	0.716745053	9.00757E-06
16	16154	6384	0.716745053	9.00757E-06	17229	6809	0.716740161	8.4456E-06
17	17229	6809	0.716740161	8.4456E-06	18304	7234	0.716735845	7.94963E-06
18	18304	7234	0.716735845	7.94963E-06	19379	7659	0.716732007	7.50868E-06
19	19379	7659	0.716732007	7.50868E-06	20454	8084	0.716728572	7.11408E-06
20	20454	8084	0.716728572	7.11408E-06	21529	8509	0.716725481	6.75888E-06

### 3.6.14 Bayesian Analysis for November

A sample selected from data for November was analyzed using the Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.44. The results in the table 3.44 below indicate that the prior and posterior means converge to 0.9000 with tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of November has a high possibility (0.9000) of sunshine.

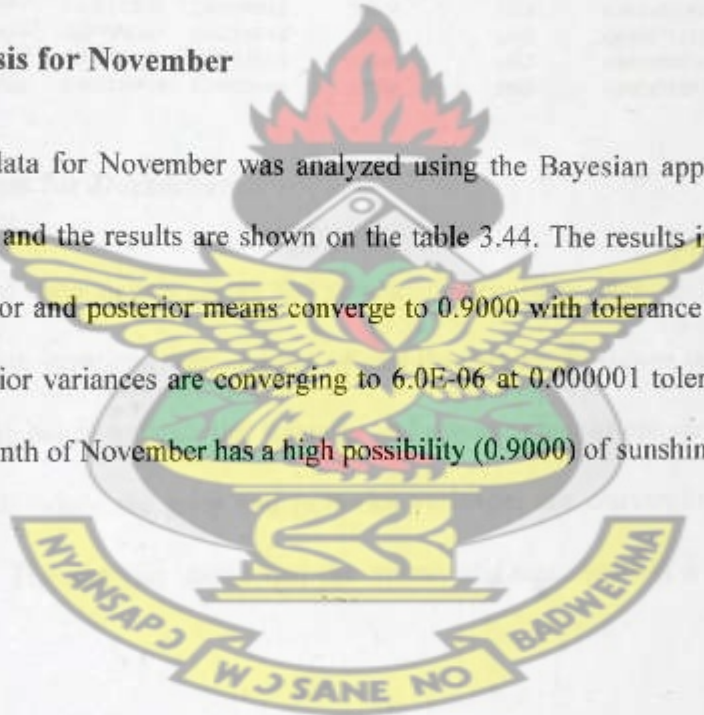


Table 3.44: Bayesian analysis for November

Iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	28	6	0.823529412	0.004152249	1406	128	0.916558018	4.98237E-05
2	1406	128	0.916558018	4.98237E-05	2784	250	0.917600527	2.49126E-05
3	2784	250	0.917600527	2.49126E-05	4162	372	0.917953242	1.66075E-05
4	4162	372	0.917953242	1.66075E-05	5540	494	0.918130593	1.24551E-05
5	5540	494	0.918130593	1.24551E-05	6918	616	0.918237324	9.96384E-06
6	6918	616	0.918237324	9.96384E-06	8296	738	0.918308612	8.30303E-06
7	7766	1276	0.858880779	1.34031E-05	9144	1398	0.867387592	1.09102E-05
8	9144	1398	0.867387592	1.09102E-05	10522	1520	0.87377512	9.1582E-06
9	10522	1520	0.87377512	9.1582E-06	11900	1642	0.8787476	7.86755E-06
10	11900	1642	0.8787476	7.86755E-06	13278	1764	0.882728361	6.88154E-06
11	13278	1764	0.882728361	6.88154E-06	14656	1886	0.885987184	6.10614E-06
12	14656	1886	0.885987184	6.10614E-06	16034	2008	0.888704135	5.48185E-06
13	16034	2008	0.888704135	5.48185E-06	17412	2130	0.891003991	4.96934E-06
14	17412	2130	0.891003991	4.96934E-06	18790	2252	0.892975953	4.54165E-06
15	18790	2252	0.892975953	4.54165E-06	20168	2374	0.894685476	4.17972E-06
16	20168	2374	0.894685476	4.17972E-06	21546	2496	0.896181682	3.86974E-06
17	21546	2496	0.896181682	3.86974E-06	22924	2618	0.897502153	3.60146E-06
18	22924	2618	0.897502153	3.60146E-06	24302	2740	0.898676133	3.36713E-06
19	24302	2740	0.898676133	3.36713E-06	25680	2862	0.899726719	3.16079E-06
20	25680	2862	0.899726719	3.16079E-06	27058	2984	0.900672392	2.97779E-06

### 3.6.15 Bayesian Analysis for December

The sample selected from data pertaining to the month of December was analyzed using the Bayesian approach. Twenty iterations were performed and the results are shown on the table 3.45. The results in the table 3.45 below indicate that the prior and posterior means converge to 0.87000 at a tolerance level of 0.0001 while the prior and posterior variances are converging to 6.0E-06 at 0.000001 tolerance level. The analysis shows that the month of December has a high possibility (.87000) of sunshine.

Table 3.45: Bayesian analysis for December

iteration	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	27	7	0.794117647	0.00467128	1332	202	0.868318123	7.44897E-05
2	1332	202	0.868318123	7.44897E-05	2637	397	0.869149637	3.74723E-05
3	2637	397	0.869149637	3.74723E-05	3942	592	0.869430966	2.50321E-05
4	3942	592	0.869430966	2.50321E-05	5247	787	0.869572423	1.87931E-05
5	5247	787	0.869572423	1.87931E-05	6552	982	0.869657552	1.50436E-05
6	6552	982	0.869657552	1.50436E-05	7857	1177	0.869714412	1.25414E-05
7	7857	1177	0.869714412	1.25414E-05	9162	1372	0.869755079	1.07528E-05
8	9162	1372	0.869755079	1.07528E-05	10467	1567	0.869785607	9.41077E-06
9	10467	1567	0.869785607	9.41077E-06	11772	1762	0.869809369	8.36653E-06
10	11772	1762	0.869809369	8.36653E-06	13077	1957	0.869828389	7.53089E-06
11	13077	1957	0.869828389	7.53089E-06	14382	2152	0.869843958	6.84702E-06
12	14382	2152	0.869843958	6.84702E-06	15687	2347	0.869856937	6.27701E-06
13	15687	2347	0.869856937	6.27701E-06	16992	2542	0.869867923	5.79461E-06
14	16992	2542	0.869867923	5.79461E-06	18297	2737	0.869877341	5.38107E-06
15	18297	2737	0.869877341	5.38107E-06	19602	2932	0.869885506	5.02262E-06
16	19602	2932	0.869885506	5.02262E-06	20907	3127	0.869892652	4.70894E-06
17	20907	3127	0.869892652	4.70894E-06	22212	3322	0.869898958	4.43214E-06
18	22212	3322	0.869898958	4.43214E-06	23517	3517	0.869904565	4.18608E-06
19	23517	3517	0.869904565	4.18608E-06	24822	3712	0.869909582	3.9659E-06
20	24822	3712	0.869909582	3.9659E-06	26127	3907	0.869914097	3.76772E-06

### 3.6.16 Bayesian Analysis for Clusters

Samples were selected in respect of months pertaining to the two clusters found and analyzed using the Bayesian approach. Twenty iterations were performed and the results are shown in tables 3.46 and 3.47 for cluster one and two respectively. The results in the table 3.46 and 3.47 indicates that the prior and posterior means converge respectively to 0.86000 and .76000 at the tolerance level of 0.0001 while the prior and posterior variances for cluster one and two converged to 6.0E-06 and 4.0E-06 respectively at 0.000001 tolerance level. The analysis shows that the clusters indicate a high potential (.86000 and 0.76000) of sunshine. This is also an indication that Kumasi generally has a high possibility of sunny days

Table 3.46: Bayesian analysis for cluster one

ITERATION S	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	23	9	0.71875	0.00612571	1103	429	0.71997389	0.000131514
2	1103	429	0.71997389	0.000131514	2183	624	0.777698611	6.15682E-05
3	2408	624	0.794195251	5.38903E-05	3488	819	0.809844439	3.57466E-05
4	3713	819	0.819285084	3.2662E-05	4793	1014	0.825383158	2.4815E-05
5	5018	1014	0.831896552	2.318E-05	6098	1209	0.83454222	1.88946E-05
6	6323	1209	0.839484865	1.7888E-05	7403	1404	0.840581356	1.52139E-05
7	7628	1404	0.844552702	1.45338E-05	8708	1599	0.844862715	1.27153E-05
8	8933	1599	0.848176984	1.22257E-05	10013	1794	0.848056238	1.09127E-05
9	10238	1794	0.850897606	1.05436E-05	11318	1989	0.850529796	9.55281E-06
10	11543	1989	0.853015075	9.26479E-06	12623	2184	0.852502195	8.4915E-06
11	12848	2184	0.854709952	8.26055E-06	13928	2379	0.854111731	7.64072E-06
12	14153	2379	0.856097266	7.45144E-06	15233	2574	0.855450104	6.9438E-06
13	15458	2574	0.857253771	6.78588E-06	16538	2769	0.856580515	6.36267E-06
14	16763	2769	0.858232644	6.22891E-06	17843	2964	0.857547941	5.87079E-06
15	18068	2964	0.85907189	5.75607E-06	19148	3159	0.85838526	5.44917E-06
16	19373	3159	0.859799396	5.34968E-06	20453	3354	0.859117066	5.08379E-06
17	20678	3354	0.860436085	4.99671E-06	21758	3549	0.859762121	4.76415E-06
18	21983	3549	0.860997963	4.68729E-06	23063	3744	0.860334987	4.4827E-06
19	23288	3744	0.861497484	4.41385E-06	24368	3939	0.86084714	4.23164E-06
20	24593	3939	0.861944483	4.17048E-06	25673	4134	0.861307747	4.00754E-06

Table 3.47: Bayesian analysis for cluster two

ITERATION S	alpha	beta	prior mean	prior variance	posterior alpha	posterior beta	posterior mean	posterior variance
1	24	9	0.727272727	0.005833738	1171	362	0.763861709	0.000117586
2	1171	362	0.763861709	0.000117586	2318	715	0.764259809	5.93826E-05
3	2318	715	0.764259809	5.93826E-05	3465	1068	0.764394441	3.97211E-05
4	3465	1068	0.764394441	3.97211E-05	4612	1421	0.764462125	2.98409E-05
5	4612	1421	0.764462125	2.98409E-05	5759	1774	0.764502854	2.38968E-05
6	5759	1774	0.764502854	2.38968E-05	6906	2127	0.764530056	1.99274E-05
7	6906	2127	0.764530056	1.99274E-05	8053	2480	0.764549511	1.70888E-05
8	8053	2480	0.764549511	1.70888E-05	9200	2833	0.764564115	1.49581E-05
9	9200	2833	0.764564115	1.49581E-05	10347	3186	0.764575482	1.32998E-05
10	10347	3186	0.764575482	1.32998E-05	11494	3539	0.764584581	1.19725E-05
11	11494	3539	0.764584581	1.19725E-05	12641	3892	0.764592028	1.08861E-05
12	12641	3892	0.764592028	1.08861E-05	13788	4245	0.764598237	9.98047E-06
13	13788	4245	0.764598237	9.98047E-06	14935	4598	0.764603492	9.21393E-06
14	14935	4598	0.764603492	9.21393E-06	16082	4951	0.764607997	8.55675E-06
15	16082	4951	0.764607997	8.55675E-06	17229	5304	0.764611903	7.98707E-06
16	17229	5304	0.764611903	7.98707E-06	18376	5657	0.764615321	7.48851E-06
17	18376	5657	0.764615321	7.48851E-06	19523	6010	0.764618337	7.04853E-06
18	19523	6010	0.764618337	7.04853E-06	20670	6363	0.764621019	6.65738E-06
19	20670	6363	0.764621019	6.65738E-06	21817	6716	0.764623418	6.30737E-06
20	21817	6716	0.764623418	6.30737E-06	22964	7069	0.764625579	5.99232E-06

## CHAPTER FOUR

### CONCLUSION AND RECOMMENDATIONS

#### 4.1 SUMMARY OF RESULTS

##### 4.1.1 Probability Distribution

It is well known that the solar irradiation reaching the surface of the earth is not normally distributed. There is the need to find the probability density function that best describes the sunshine hours in the various month of the year. In the present study, the method of mean square error was used to obtain the best distribution among a selection of standard distributions for each month of the year. In each case the standard probability distribution with the smallest mean square error relative to observed data was considered the best probability density function.

It was thus found that the solar irradiation for the month of April was Weibull distribution January, March and May is Lognormal distributed, and those for the months of June, August, September October and November follows Exponential, Weibull, Geometric and gamma distribution while February July and December were Lognormal, Exponential, Weibull, Geometric and gamma distributed.

Table 4.1 below summarizes the results for the various months indicating the relevant probability density function

#### 4.1.1 Classification of Distribution by Months

Table 4.1 summary of monthly probability distributions

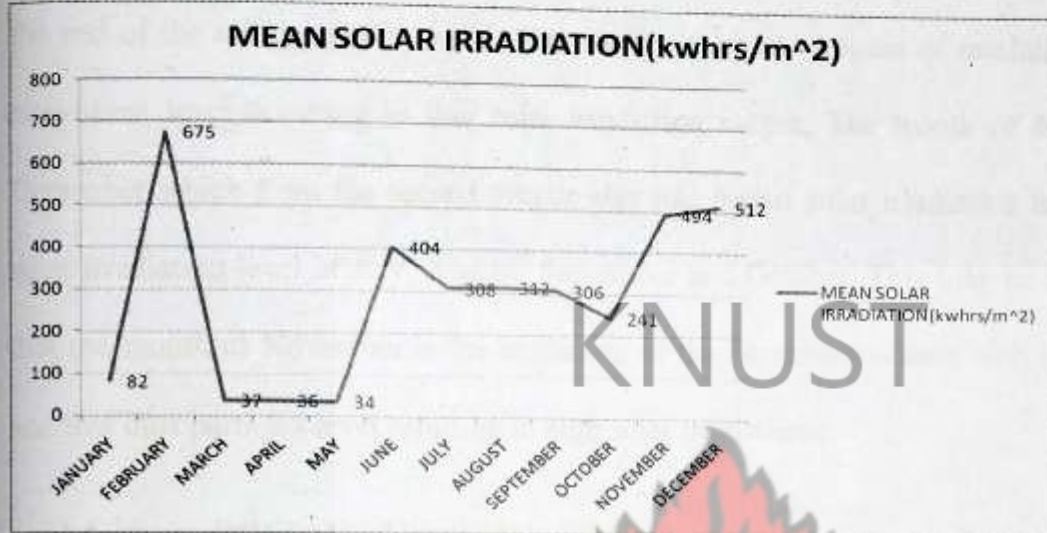
Month	Probability distributions
January	Lognormal
February	Lognormal, Exponential, Weibull, Geometric and gamma
March	Lognormal
April	Weibull, exponential, gamma and geometric
May	Lognormal
June	Exponential, Weibull, Geometric and gamma
July	Exponential, Weibull, Geometric, Lognormal and gamma
August	Exponential, Weibull, Geometric and gamma
September	Exponential, Weibull, Geometric and gamma
October	Exponential, Weibull, Geometric and gamma
November	Exponential, Weibull, Geometric and gamma
December	Exponential, Weibull, Geometric, Lognormal and gamma

Detailed charts and tables for the various probability density functions are displayed in the

Appendix (II)

#### 4.1.2 Mean hourly solar irradiation of the Months

The figure 4.1 Mean hourly solar irradiation of the Months



There are certain climatic factors that are accounting for the rise and fall of the hourly solar irradiation pattern in the year. The factors include cloudiness and the presence of dust particles in the atmosphere. Fig 4.1 above indicates that the maximum mean solar irradiation received occurred in February. This perhaps may be due to the fact that February marks the end of the harmattan and the beginning of the rainy season where both cloudiness and dust particles level in the atmosphere are low and hence the solar irradiation level reaching the earth is high. The months of March, April and May were found to be the months with the least mean solar irradiation. Perhaps this may be due to the fact by March, the rainy season would have begun for some time and the cloudiness level would be quite high resulting in a low level of solar irradiation. In the month of June there was a rise in solar irradiation even though the solar irradiation level falls in July, August, September and October which may

perhaps be due to the fact that in June there is intermittent cloudiness and sunshine level. In some cases we have a cloudy day throughout and also sunshine through certain day's amount of rainfall is not as in March April and May. The cloudiness level reduces and so the solar irradiation output is higher. The month of July, August, September and October marks the end of the raining season which also results in higher amount of rainfall hence higher cloudiness level resulting in low solar irradiation output. The month of November and December which form the second cluster also had higher solar irradiation level above the solar irradiation level of July, August, September and October. This may be due to the fact that the month of November is the beginning of the harmattan season with low cloudiness and low dust particles level resulting in high solar irradiation.

#### 4.1.3 Putting the Months into Clusters

Based on the theory of hypothesis testing the various months of the year with similar solar irradiation patterns were put into clusters. From the hypothesis testing carried out it was discovered that the months of the year can be put into two such clusters. The first cluster extends from the month of January to October and the second cluster is from November to December.

#### 4.1.4 Bayesian Estimation for Months

As indicated in table 4.2, the Bayesian analysis was conducted for the various months of the year. The summary of the results are given in table 4.2. The results presented are the means and the variance for twenty iterations for the various months of the year. The results in table 4.2 indicate that the beta distribution which was used as the prior distribution converged to

the mean and variance of the posterior distributions shown. The posterior means of various months indicates that despite the rain fall patterns in Kumasi there is a higher potential sunshine region

Table 4.2 Final iterative values of the Bayesian analysis

MONTH	PRIOR MEAN	PRIOR VARIANCE	POSTERIOR MEAN	POSTERIOR VARIANCE
JANUARY	0.869	3.97E-06	0.899	3.77E-06
FEBRUARY	0.824	6.02E-06	0.824	5.71E-06
MARCH	0.9	3.16E-06	0.9	2.99E-06
APRIL	0.859	4.22E-06	0.859	4.02E-06
MAY	0.923	2.48E-06	0.923	2.35E-06
JUNE	0.769	6.22E-06	0.769	5.96E-06
JULY	0.718	7.25E-06	0.718	6.88E-06
AUGUST	0.536	5.45E-06	0.536	5.27E-06
SEPTEMBER	0.785	5.91E-06	0.785	5.61E-06
OCTOBER	0.717	7.11E-06	0.717	6.76E-06
NOVEMBER	0.899	3.16E-06	0.899	2.98E-06
DECEMBER	0.869	3.97E-06	0.869	3.77E-06
BAYESIAN SUMMARY OF THE CLUSTERS				
CLUSTER ONE	0.86	4.17048E-06	0.86	4.00754E-06
CLUSTER TWO	0.765	6.30737E-06	0.764	5.99232E-06

## 4.2 CONCLUSION

The hourly solar irradiation data for Kumasi, Ghana, was analyzed in order to determine standard probability distribution models pertaining to the various months of the year, as well as for clusters of months of the year with similar patterns. From the analysis conducted it was found out that the month of February has the highest mean hourly solar irradiation of  $675Whm^{-2}$  and the month of May has least mean hourly solar irradiation of  $34Whm^{-2}$ . From the analysis conducted it was also found that hourly solar irradiation for January, March and May could be fitted to the lognormal probability distribution while the month of April could

be equally well fitted to the Weibull, exponential, gamma and the geometric distributions. The months of June, July, August, September, October, November and December could be equally well be fitted to the Exponential, Weibull, Geometric and Gamma distributions.

From hypothesis testing carried out, it was discovered that the months of the year could be put into two clusters. The first cluster is from the month of January to October and the second cluster comprises November and December. It was observed that there is a significant difference between the mean hourly solar irradiation pertaining to the two clusters. The first cluster had a higher mean hourly solar irradiation of  $428Whm^{-2}$  while the second cluster had a mean hourly solar irradiation of  $326Whm^{-2}$ .

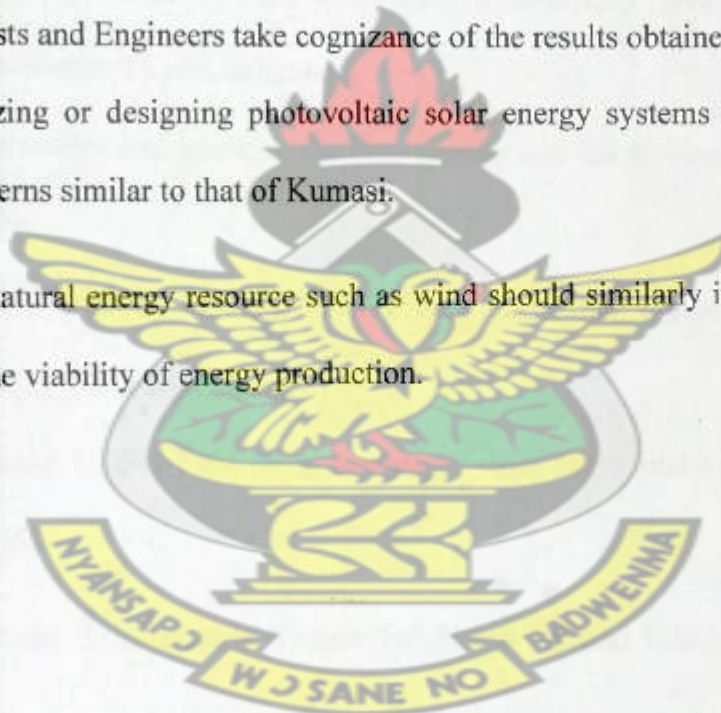
Bayesian analysis on high or low (i.e. respectively above or below a threshold of  $120kWhm^{-2}$  hourly solar irradiation) for each month with given prior beta distribution converged to posterior beta distribution after 20 iteration with average mean of 0.86. This shows that the on average the solar irradiation patterns in Kumasi tends to be high frequently. Also the prior variance of the various months of the year converged to the posterior tolerance level of 0.000001

Comparing the sunshine output of the two clusters indicated that the first cluster has a higher hourly solar irradiation output than the second cluster. The Bayesian analysis of the clusters confirms that this and we find that the first cluster has a higher possibility (0.8600) of high solar irradiation output as compared with the second cluster (0.7600).

## RECOMMENDATIONS

The following are the recommendations made after the study.

1. That a similar and more comprehensive study should be conducted in respect of other geographical locations to ascertain and compare hourly solar irradiation distributions.
2. That Scientists and Engineers take cognizance of the results obtained in this work when analyzing or designing photovoltaic solar energy systems in areas with weather patterns similar to that of Kumasi.
3. That other natural energy resource such as wind should similarly investigated to determine the viability of energy production.



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APENDIXSAMPLED DATA OF THE VARIOUS MONTHS

SAMPLED DATA FOR THE MONTH OF APRIL									
19.9	18.72	12.88	19.9	81.9	1.171	12.88	16.36	598.2	873
83.1	150.9	180.2	83.1	289	125.3	147.5	268.7	414.8	981
165	293.5	426.8	165	449	243.4	369.5	58.43	21.09	47.99
267.8	500.1	552.6	267.8	898	464.4	614.4	133.3	112.5	121.7
383.5	729	780	383.5	768	661.6	769	133.3	120.6	558
653.3	763	819	653.3	875	519.9	890	17.56	237.7	374.8
697.3	842	750	697.3	745	586.4	751	226.9	657.4	120.3
462.3	792	673.2	462.3	777	731	822	491.9	782	1.168
454.3	664.7	821	454.3	577.8	727	663.7	721	994	271.3
312	354.6	525.1	312	330.4	421.5	459.7	877	1006	268.9
355.2	206.5	301.1	355.2	105.1	191.5	290.6	600.4	986	332
210.3	50.18	126	8.17	590.1	33.88	117.9	657.6	578.9	260.2
4.673	210.3	9.37	22.25	861	76.2	24.57	47.99	481.6	428.4
169.5	4.673	165.1	133.5	722	364.7	157.9	121.7	608.8	19.85
26.94	169.5	345.2	498.1	737	4.685	16.39	177.9	356.9	57.38
148.7	406	53.84	531.8	451.9	113.6	17.56	197.7	241.5	172.1
265.6	589.1	147.5	262.1	303.6	300.9	30.42	271.3	60.68	88.6
53.84	791	839	181.3	68.91	361.6	132.2	268.9	168.3	862
147.5	440.9	974	366	308.1	590.5	288.8	332	260.7	848
262.1	543.5	934	358.9	821	687.2	509.4	168.3	187.1	861
181.3	555	665.2	448.8	842	599.3	630.4	260.7	79.5	606.9
366	858	535.7	337.7	848	634.2	789	187.1	14.03	552.1
358.9	867	403.9	308.4	621.1	345.7	878	79.5	9.37	148.4
448.8	741	142.4	199.8	364.3	181.4	650.8	14.03	105.4	5.844
337.7	542.8	4.671	0	133.2	438.3	439.8	404.8	432.6	325.1
308.4	294.2	137.9	565.2	15.19	644.5	289.4	680	698.7	319.8
199.8	53.71	452.9	93.7	11.72	412.1	4.679	93.6	764	199.7
137.9	1.168	731	188.5	105.4	157.6	127.5	159.1	773	135.4
1.168	602.2	773	280.8	746	11.68	230.3	255	16.39	687.2
532.8	608.9	497	248.1	171.5	5.859	476.8	88.6	121.7	941
446.3	680.9	285.9	271.4	771	118.3	713	333.5	318.7	345
508.3	608.7	91	662.9	785	338.5	705	701	143.6	341.4
116.8	584.4		70.2	884	624.7	751	810	17.51	475.8
0	495.5	0	76.1	566.3	821	797	719	76.1	455.7
116.8	280.5	4.682	154.4	379.5	1022	607.7	431.9	209.5	114.5
1.169	119.2	70.2	386.1	133.1	756	497	622.3	270.2	225.6
0	2.335	395.2	582.9	149.5	676.1	381.6	541.6	414.4	99.3
97.2	140.5	529.5	621.4	9.35	512.8	108.6	37.46	223	47.93
35.12	291.4	557.4	508	525.4	378.5	11.7	141.6	120.3	664.9
142.8	382.5	592.4	456.1	524	265.3	114.6	225.8	649	781
228.1	466.5	209.1	436.1	979	118.1	278.3	523.9	638.5	980
4.682	829	2.338	496.7	802	8.19	438.1	507.2	2.343	259.2
70.2	748	404.6	93.7	598.7	12.89	549.9	397.3	105.4	102.9
70.2	1051	514.3	325.3	266.1	141.7	734	310.8	15.22	28.09
76.1	853	523.6	378.9	112.2	394.4	520.3	273.4	139.3	270.2
154.4	655.6	231.2	620.8	100.7	608.2	714	415.9	368.5	318.9
386.1	445.7	72.4	980	184.9	698.9	467.7	262.9	520	467.1
404.6	197.2	5.844	928	453.8	752	260.1	148.4	438	444.8

SAMPLED DATA FOR THE MONTH OF AUGUST

23	967	245	96	268	269	20	725
117	763	428	33	355	439	138	734
279	612	475	27	131	491	246	675
414	458	330	172	47	615	182	64
390	269	373	213	22	733	268	24
450	40	315	378	86	505	551	7
551	46	183	684	181	491	395	28
341	98	61	902	292	266	489	142
305	178	16	837	546	211	498	375
360	307	87	640	719	20	333	621
153	504	236	536	705	22	202	700
55	666	277	432	697	102	61	653
15	686	532	238	518	267	23	775
61	680	670	45	346	422	82	643
181	579	398	44	180	690	134	516
370	445	611	189	73	563	161	165
361	222	430	315	30	650	261	215
134	31	437	472	137	511	306	17
177	46	109	660	310	294	368	25
517	150	30	594	267	215	384	113
634	161	13	891	195	124	280	250
277	344	57	769	187	32	244	313
187	568	99	598	280	43	182	658
46	572	153	327	219	243	44	682
15	680	206	61	147	441	6	828
93	490	169	8	146	608	47	763
149	451	415	48	109	758	98	603
370	301	335	99	52	839	312	387
505	179	356	431	21	934	489	200
432	32	233	679	98	729	412	23
216	19	261	797	108	565	359	5
515	135	51	895	324	67	127	32
444	350	7	877	584	19	188	61
545	627	51	842	440	2	61	208
225	758	87	655	666	47	193	163
94	870	199	96	296	139	93	167
11	887	172	88	332	235	7	238
70	719	196	40	257	406	42	248
196	545	204	60	86	676	133	251
216	426	282	199	40	693	155	300

SAMPLED DATA FOR THE MONTH OF DECEMBER

120.4	427	58.42	100.7	577.2	161.2	63.2	411.4	173.1
14.05	557.4	133.4	207.1	392.5	25.7	165	584.2	235.1
194.3	421.8	3.512	349.7	473.1	201.6	314.7	301.4	293.5
403.3	504.7	76.1	384.6	364.5	4.685	443.1	340	265.4
406.6	478.9	188.4	594.8	334.2	119.4	386.8	130.9	68.98
477.8	443.8	446.8	827	163.6	116.5	657.7	254.9	308.8
453.2	348.1	495.6	665.4	11.69	8.19	418.1	90.1	186
678.6	244.1	642.6	309.3	153.7	78.4	127.3	110	292.4
843	206.8	598	206.7	47.99	201.3	164.7	320.6	128.6
735	158.9	425.1	150.6	148.6	263.2	2.337	188.4	2.339
504.4	9.35	399.4	95.8	260.9	168.4	134.1	65.53	90.9
429.7	163.5	596.8	5.841	286.6	244.4	32.77	28.09	3.513
204.4	79.6	362.1	162.6	302.9	329.7	108.8	74.7	69.09
24.53	200.1	144.9	14.04	277.1	210.4	438.6	62.06	254.1
223.7	216.4	16.36	101.8	203.4	251.3	412.6	118.2	577.9
15.22	229.2	183.2	142.7	194.1	168.3	342.4	133.4	466.4
106.5	406.9	2.342	248	143.8	162.5	483.7	272.6	358.8
205.9	694.3	76.1	339.1	116.9	74.8	295.5	419.9	348.3
327.5	724	158	372.9	54.95	1.169	275.7	675.6	382.3
593.8	667	238.7	351.8	84.9	90.1	267.5	576	422
728	526.8	658.4	168.3	15.23	21.07	213.8	686.7	167.2
480	397.2	559.8	327.3	125.3	118.2	68.95	572.2	15.2
607.2	142.6	583	239.6	344	93.6	8.18	370.2	127.7
567.5	11.69	410	78.3	431.4	472.6	122.9	120.3	40.98
502.1	179	568.8	18.7	527	707	78.4	3.505	66.72
360.9	5.854	463.7	100.1	575.8	735	136.9	167.1	107.7
164.7	77.3	357.5	14.05	491.6	671.6	156.8	14.05	374.5
25.7	246.9	181.1	104.2	556.9	578.1	361.4	100.7	508.8
195.2	349.7	24.54	195.4	545.2	506.9	217.5	235.2	653.4
1.171	533.1	178.4	342.7	300	311.8	514.4	324	313.1
151	590	58.53	556.4	217.2	123.8	463.9	192.9	307.5
246.8	573.5	243.4	691.6	103.9	1.169	564.3	325.1	91.2
325.1	450.8	260.9	575.7	7.01	180.9	650.8	814	47.97
433.7	316.4	523.9	597.8	176.7	67.88	320.2	574.8	16.38
305	233.5	538.8	731	52.68	117	59.61	502.3	105.3
550.4	293.1	587.6	575.6	84.3	175.5	146.8	370.4	18.74
341.1	137.8	854	332.8	112.3	214.1	39.8	158.9	106.5
319	22.19	718	174	156.8	376.5	122.9	150.5	179.1
294.4	159.6	565.1	8.18	297.1	329.6	232.8	83.1	253.9
175.3	4.682	301.3	204.1	323.9	566.8	340.3	143.9	345.1

SAMPLED DATA FOR THE MONTH OF FEBRUARY

216.6	421	556.8	512	496	156	348	781	764
298.4	222	344.4	717	808	39	155	779	820
455.1	48	67.73	790	772	32	20	654	702
641.9	96	48	840	876	140	113	414	597
769	265	230.6	788	849	378	304	324	369
680	427	479.5	634	787	669	510	72	194
428.8	653	460.4	472	548.8	776	666	88	54
331.8	764	555.9	183	350.3	803	720	181	118
130.9	809	879	76	234.7	720	702	475	307
1.169	735	662.9	68	30.37	632	660	589	501
70.3	615	743	167	113.6	453	586	646	674
268	413	575.5	412	313.7	236	417	769	753
381.2	202	369	708	535.6	61	207	738	760
417.3	46	162.3	798	697.5	111	34	469	656
735	107	10.51	840	691.4	276	102	473	538
613.1	293	63.22	816	672.7	485	285	273	365
698.2	492	177.9	699	845	668	482	76	168
527.8	651	354.5	490	757	737	643	55	25
578	716	584.5	263	493.9	800	701	187	109
346.8	724	644.9	54	314.2	729	701	296	295
121.5	691	780	109	196.3	635	626	462	496
60.89	562	694.6	322	17.53	446	511	772	663
244.6	379	634	524	229	232	320	839	702
288.9	193	458.9	693	411	60	152	759	702
464.2	35	224.2	794	562	66	23	638	674
649.7	45	225.4	839	665	165	95	418	538
631.8	104	97.3	756	701	278	276	212	360
785	304	292.8	648	621	499	477	63	166
572	514	474.9	466	506	648	629	94	21
396.9	677	698.9	249	347	816	719	171	103
272	809	820	64	162	751	721	182	282
98.1	815	799	93	26	642	679	256	494
83.1	735	733	276	86	468	559	717	643
259.7	609	677.2	407	242	242	404	525	736
493.3	425	559.4	668	442	67	209	764	771
703	217	405.3	752	596	55	41	649	703
720	46	228.9	785	698	180	83	512	588
455.1	51	15.19	720	701	308	266	323	402
659.5	135	111.3	594	672	709	439	43	192
696.9	336	292.7	407	530	707	639	114	32

SAMPLED DATA FOR THE MONTH OF JANUARY

726	121	388	452	35	395	118	277.4	737
752	48	182	647	101	353	230	458.8	760
681	132	41	713	191	432	370	485	781
546	213	9	601	433	288	463	540.6	305
360	497	78	619	672	111	698	323.3	360
166	304	222	546	696	65	655	249.6	244
35	492	494	370	624	192	689	157.3	41
10	745	652	237	499	462	554	47.25	9
101	732	717	86	339	365	415	0.814	74
263	725	738	36	163	582	135	18	283
451	499	680	152	38	605	57	14	529
614	277	566	216	8	814	58	3.642	547
699	63	421	369	119	865	173	56.35	776
707	56	223	505	145	476	396	141.1	720
640	224	52	579	285	129	281	287.1	783
500	463	15	727	485	101	437	432.8	639
327	575	133	597	635	40	660	410.6	442
155	847	308	603	655	39	820	411.3	249
31	869	509	213	647	116	655	615.1	68
12	834	671	106	521	211	627	402.8	5
99	728	613	44	356	359	399	311	46
246	676	744	17	169	553	222	99.3	89
449	482	679	60	35	747	68	28.03	219
615	273	548	184	5	824	44	1.05	317
705	83	372	382	40	766	185	3.194	440
710	56	187	466	132	549	490	3.24	716
650	222	44	622	322	450	464	3.383	671
526	297	10	503	357	272	520	3.303	543
356	430	106	354	486	94	721	3.043	466
170	732	288	369	606	59	888	0.209	253
38	752	487	163	765	514	47	2.664	31.42
13	605	640	54	661	650	174	2.961	161
106	597	679	29	537	755	320	2.47	298.4
274	466	703	45	263	650	421	1.857	474
449	357	677	174	84	443	496	0.568	683
599	87	539	242	33	213	460	0.401	2.819
679	23	368	574	126	101	542	1.655	2.394
699	36	199	670	130	11	36.92	21.51	0.502
640	115	43	541	530	133	133.9	81.2	0.31
526	297	4	609	424	291	250.6	116.1	2.087

62	371	449	659	141	188	13	241	489
121	503	584	742	200	74	106	373	709
295	489	590	567	100	37	212	489	508
504	587	465	406	45	136	358	639	400
474	570	367	245	58	211	466	612	170
522	610	336	86	150	257	728	676	171
519	499	218	54	325	492	547	518	90
428	399	42	201	465	623	455	420	25
666	217	45	218	447	540	394	260	17
446	82	205	410	649	478	265	89	112
251	24	180	654	599	347	137	18	208
71	95	200	688	656	333	61	39	247
50	194	456	599	558	165	55	117	325
93	330	549	548	537	59	145	205	270
342	333	592	442	295	16	247	368	387
407	517	651	427	52	95	359	529	255
236	521	606	289	37	128	503	697	250
608	354	259	88	139	342	548	382	327
589	454	183	12	236	372	516	361	193
228	250	88	26	173	325	565	317	53
109	253	18	81	219	386	411	156	18
24	61	74	216	448	573	370	76	89
7	34	136	138	446	525	236	14	219
5	118	186	220	307	266	77	78	227
38	179	257	243	154	286	7	167	390
175	338	281	267	64	72	62	266	541
294	460	539	273	73	31	308	364	602
435	746	553	217	31	97	384	289	552
331	723	601	142	28	199	443	353	550
348	721	480	65	83	463	338	653	468
307	582	248	34	305	656	455	258	255
534	376	105	169	417	743	408	108	95
389	116	28	321	513	688	362	52	37
515	54	136	416	636	719	301	24	87
100	26	213	497	687	591	159	13	121
41	105	294	469	632	463	68	65	152
37	255	440	461	687	266	30	147	282
101	286	443	649	447	89	59	174	258
328	463	710	389	288	34	166	247	400
454	496	617	280	93	124	327	376	492

SAMPLED DATA FOR THE MONTH OF JUNE

37	31	436	618	382	124	42	221	512
88	104	733	851	368	193	74	424	680
146	242	815	663	169	87	239	581	735
325	278	732	570	58	39	386	627	404
338	246	612	444	63	120	539	674	446
405	417	388	360	165	195	627	509	279
537	582	252	266	459	347	832	231	229
472	557	77	89	629	485	726	156	50
408	324	19	29	748	736	634	91	39
298	275	77	161	797	701	363	35	100
149	177	170	229	740	782	179	38	276
62	41	232	385	676	675	62	131	418
75	49	454	396	677	414	41	202	433
234	181	577	733	433	249	162	201	513
318	208	620	752	248	75	233	318	533
371	283	609	748	80	36	414	604	729
518	691	616	633	30	165	471	692	504
697	737	444	270	100	205	453	330	241
905	657	209	128	314	457	909	444	117
700	743	74	32	499	572	609	259	75
420	708	71	107	557	720	315	71	41
455	498	115	262	623	657	251	60	123
203	253	217	359	846	726	164	49	298
86	47	372	409	676	674	57	175	265
39	82	738	643	622	492	35	243	400
169	224	678	825	470	261	95	406	428
322	208	862	775	249	33	156	690	540
418	285	608	450	96	93	277	636	439
66	407	663	512	55	135	523	456	438
45	576	398	261	133	331	459	818	398
106	869	228	87	258	585	754	515	178
143	824	81	25	459	440	841	296	46
151	621	66	87	674	596	463	193	61
217	448	182	260	758	529	588	35	209
159	186	207	354	811	640	93	43	406
51	53	1110	562	606	313	12	122	378
26	33	76	777	707	193	36	268	625
88	143	290	747	374	36	143	453	504
169	284	986	621	218	39	340	491	343
246	508	492	309	70	106	388	579	754

SAMPLED DATA FOR THE MONTH OF MARCH									
162	881	508	432	522	102	158	557	783	
405	883	713	283	258	283	276	562	865	
628	880	851	32	52	619	463	651	777	
715	680	682	20	9	707	697	360	588	
828	440	456	110	138	729	460	314	464	
840	248	191	231	357	829	600	204	282	
681	75	40	535	553	785	722	72	88	
242	24	6	747	687	703	643	21	22	
342	117	83	862	636	473	486	198	158	
176	299	188	876	441	246	279	418	387	
59	359	408	829	678	79	88	589	605	
10	531	539	672	724	44	14	756	727	
102	660	705	386	538	202	89	877	895	
373	791	679	270	232	417	218	884	835	
685	706	778	77	55	575	301	794	766	
740	647	507	22	18	780	507	673	631	
611	363	414	143	117	837	557	493	478	
575	256	304	344	303	646	500	261	274	
824	61	64	538	413	756	651	89	94	
666	46	9	745	523	651	662	17	12	
473	190	99	905	741	377	321	85	93	
194	177	290	885	687	128	121	265	265	
65	291	562	860	785	54	40	510	587	
20	477	501	729	478	33	26	771	745	
162	742	656	488	429	202	184	675	730	
383	578	573	297	247	325	408	744	767	
583	627	471	94	87	619	552	542	946	
713	689	565	20	6	522	565	612	752	
878	486	447	168	44	700	870	437	362	
944	126	236	407	123	635	774	285	227	
897	55	67	570	241	828	872	72	40	
542	25	14	841	276	448	761	169.3	17	
452	116	84	805	399	447	556	66.7	56	
276	202	204	767	672	244	306	180.2	167	
82	587	361	633	757	77	94	191.8	442	
7	845	704	531	535	34	24	236.2	569	
34	812	769	429	439	129	171	377.6	793	
55	885	895	181	230	197	367	527	688	
87	894	736	89	93	456	600	546.7	565	
190	704	667	15	19	524	784	437.9	617	

SAMPLED DATA FOR THE MONTH OF MAY								
20	270	740	126	211	645	351	16	111
74	57	721	301	503	123	457	8	46
220	79	618	544	485	244	207	36	39
542	211	663	429	547	256	146	168	31
665	347	691	561	784	113	104	532	133
578	443	526	582	503	9	72	616	310
329	466	329	659	724	32	26	805	446
435	358	26	534	603	99	56	916	743
398	396	36	275	432	194	194	868	879
395	391	154	62	289	355	459	756	682
219	401	327	38	145	437	559	510	639
82	214	328	87	51	431	705	470	443
24	105	490	232	173	462	700	271	252
131	15	582	344	351	791	575	92	111
148	59	661	701	511	477	756	88	65
467	236	600	761	619	393	552	298	53
646	236	403	507	696	257	525	516	170
784	312	329	404	758	91	301	607	443
837	468	343	383	686	27	99	597	506
801	570	52	72	592	182	37	633	655
611	466	252.9	32	434	321	218	595	646
442	774	0	28	234	234	362	440	733
251	698	39.78	48	75	644	530	373	828
92	556	203.5	91	39	253	898	288	686
39	286	407.8	196	189	755	743	224	445
265	104	589.5	394	459	786	599	79	80
312	50	794	530	613	535	627	21	16
445	140	19	780	796	107	721	109	28
554	285	148	746	821	16	485	159	99
626	420	362	581	752	3	243	326	319
826	674	443	459	571	23	55	557	588
546	629	721	458	396	39	56	388	649
710	708	568	132	203	147	193	336	851
430	318	648	18	110	205	324	351	651
220	518	647	33	46	475	523	480	715
43	342	654	202	38	478	697	311	492
44	98	419	422	165	560	477	205	411
280	24	293	605	213	445	472	111	167
513	31	59	496	252	493	595	24	28
635	68	29	601	438	204	254	83	323

SAMPLED DATA FOR THE MONTH OF NOVEMBER

29	134	651	655	631	253	56	262	529
149	226	717	636	401	47	29	359	640
304	239	853	503	150	82	96	409	313
712	544	728	326	27	224	178	643	193
747	698	596	140	23	208	296	825	36
815	667	376	20	80	285	570	775	39
819	492	172	34	152	407	532	450	106
735	289	28	150	172	576	546	512	221
437	129	21	284	150	869	530	261	424
390	22	108	475	197	824	358	87	581
175	25	190	538	489	621	295	25	627
29	198	388	546	681	448	105	87	674
53	393	517	570	513	186	9	260	509
208	549	770	580	365	53	28	354	231
333	661	502	479	160	33	140	562	156
363	787	710	248	25	143	238	777	91
399	800	439	120	17	284	307	747	35
597	699	287	15	80	508	484	621	38
775	511	179	27	170	618	660	309	131
601	329	38	113	436	851	652	124	202
447	141	50	221	526	663	690	193	201
360	22	157	347	558	570	526	87	318
174	36	285	465	749	444	320	39	604
28	147	293	584	672	360	158	120	692
40	352	582	581	482	266	22	195	330
92	490	582	545	239	89	30	347	444
200	725	573	325	102	29	72	485	259
235	789	715	186	22	161	657	736	71
352	730	546	56	17	229	726	701	60
423	690	324	7	78	385	674	782	49
654	533	153	49	236	396	492	675	175
672	332	24	181	442	733	261	414	243
452	141	25	208	418	752	9	249	406
372	23	87	283	613	748	33	75	690
170	32	168	691	566	633	93	36	636
269	132	249	737	536	270	135	165	456
14	274	366	657	541	128	331	205	818
115	398	656	743	273	9	585	457	515
268	574	666	708	99	32	440	572	296
521	607	560	498	16	107	596	720	193

SAMPLED DATA FOR THE MONTH OF OCTOBER

50.09	3.571	341.7	255	280.9	517.9	17.88	319.8	220.5
110.9	28.62	645.2	250.1	434.5	580.7	82.3	335.3	288.3
101.3	110.8	621.2	296.6	115.5	552	193.1	221.2	508.5
137.1	226.3	735	403.6	5.964	423.5	452.7	91.6	491.6
473.1	345.3	401	412.9	73.9	265.3	567.8	67.82	456.9
318	565.3	273.7	317.6	163.4	147.5	395	103.7	722
346.5	767	107.1	342.6	236	69.03	604.3	201.4	584.1
78.7	596.9	4.762	207	159.7	29.81	399.6	290.5	480.6
144.3	632.7	48.9	99.9	537.3	140.7	614.9	314	235.6
143.1	396.1	140.7	70.2	524	234.8	403.2	324.7	46.42
91.8	325.9	153.8	4.76	627.4	233.5	192.7	344.8	54.86
59.68	172.5	272.9	1.193	571.5	459.7	67.82	305.6	195.5
163.5	70.2	402.6	76.3	585.8	509.3	2.381	214.1	349.2
189.7	2.38	673.7	199.1	360.8	499.5	15.5	214.1	414.5
430.4	15.5	633.1	284.9	98.8	409.1	66.78	108.3	514.2
601.7	89.4	690.1	506.3	4.771	398.5	158.6	64.25	672.2
831	181.2	512.9	503.5	121.6	236.7	378.9	1.19	705
632.1	214.5	336.8	434.3	201.4	163	520.3	2.381	622
423.7	331.2	105.9	352.1	259.8	69.03	405.9	28.62	346.1
502.4	397.7	5.952	296.2	380	1.19	806	125.2	71.4
170.2	491.5	28.62	198.7	529.8	9.54	554.3	319.4	63.09
66.68	416.4	79.9	75	499.9	89.4	584	367.9	3.577
2.382	387.9	207.4	65.46	815	166.9	532.9	335.6	88.2
31.01	276.1	218.1	1.193	556.9	244.3	343.8	356.9	131.1
110.9	132.1	305	109.7	512.9	413.4	63.06	344.9	417
226.5	67.84	178.7	161	333.3	439.2	5.951	242.6	488.1
313.4	93	263.3	214.6	104.8	492.5	3.578	214.1	572.3
470.6	210.9	354.9	195.4	4.762	468.6	106.1	190.4	705
690.6	366.8	341.7	170.4	81.1	340.2	141.9	67.83	574.5
556.9	454.5	229.8	556.3	107.3	218.9	188.3	2.38	534.1
541.4	530.5	92.9	639.3	290.8	165.4	554	108.5	467.5
466.4	545.8	1.191	463	427.6	71.4	828	236	258.2
171.4	486.4	82.3	317.8	603.6	1.19	567.4	298.9	70.2
114.3	350.9	271.9	160.7	953	9.54	422.2	354.6	3.571
5.953	226	257.5	58.34	924	88.2	735	369.9	21.47
1.193	147.5	389.6	22.67	697.1	214.4	538.9	355.6	122.8
112.1	70.2	605	115.7	484.2	278.5	310.6	347.3	232.4
263.4	5.95	561.8	224.1	352.2	347.4	89.3	236.7	472.8
331.2	10.73	480.8	358.7	124.9	390	1.19	190.4	609.4
409.7	116.8	357.1	358.5	22.61	403	63.2	135.7	72.6

SAMPLED DATA FOR THE MONTH OF SEPTEMBER

193	140.1	16.36	163.5	85.5	35.12	56.2	64.37	177.9
19.9	185.3	133.3	66.71	184.9	145.1	144	70.2	197.7
83.1	36.28	47.99	73.7	273.8	153.3	172	108.8	196.5
165	84.3	121.7	193.1	236.2	133.4	118.2	256.2	417.5
267.8	117	177.9	326.3	304	168.4	324.1	251.5	424.3
383.5	157.9	197.7	290	392.8	257.3	315.8	276	492
653.3	237.4	271.3	233.8	169.5	437.2	281.8	598.6	400.7
697.3	353.1	268.9	275.8	90.1	531.7	76	696.5	384.3
462.3	624.1	332	439.4	169.6	523.4	174.3	463.8	234.8
454.3	565.4	168.3	407.8	60.84	241.8	253.8	364.5	73.6
312	327.1	260.7	319	1.17	122.7	65.48	98.1	8.18
355.2	244.2	187.1	147.2	83.9	1.169	82.6	135.4	127.8
210.3	130.9	79.5	22.2	1.171	114.6	37.46	166.2	49.16
4.673	16.36	14.03	116.5	69.08	2.341	132.3	236.3	88.9
169.5	120.6	88.6	9.36	199	67.88	380.3	283	329.9
53.84	32.79	37.46	79.6	280.8	218.8	373.1	289.9	460.6
147.5	70.3	141.6	424.6	410.5	499.2	566.9	497.8	528.1
262.1	302.1	225.8	327.2	588	580.6	470.9	542	412.3
181.3	885	523.9	586.4	512.9	550	467.3	886	262.8
366	868	507.2	676	620.3	740	495.3	537	670.5
358.9	586.5	397.3	626.7	567.6	487.9	371.5	319.9	520.9
448.8	808	310.8	863	211.4	692.1	219.6	175.1	280.3
337.7	658.7	273.4	729	145	298.8	92.3	129.6	100.5
308.4	419.3	415.9	512.3	2.338	434.3	8.18	5.854	154.3
199.8	185.7	262.9	234.6	150.3	190.3	150.6	81.9	8.19
137.9	52.58	148.4	8.17	7.03	3.503	10.54	263.3	80.8
1.168	202.9	5.844	211.6	72.6	198.6	72.6	478.3	166.2
116.8	5.852	135.4	30.44	89	8.19	100.7	534.1	224.6
4.682	69.05	56.19	166.2	167.4	83.1	100.6	358.7	507.4
70.2	180.2	307.8	271.4	382.6	345	232.8	504.7	363.4
70.2	336.8	345	374.2	391.7	317.9	651.2	608.5	460.3
76.1	365.9	341.4	340.2	427.9	446.3	337.7	594.5	488.2
154.4	544.5	475.8	785	499	200.9	357.5	552.4	464.7
386.1	467.2	455.7	869	260.6	431.1	442.7	283.8	286.1
404.6	380.7	114.5	932	294.5	606.1	281.5	107.5	106.3
514.3	745	225.6	650.1	278.1	637.5	278	21.07	127.4
523.6	566.3	99.3	300	146.1	134.3	86.5	217.7	8.18
527.1	175.1	47.93	213.7	7.01	42.04	1.169	259.8	137.1
255.9	77.1	102.9	122.6	126	1.169	123.1	205.9	21.07
195.1	10.51	44.5	210.6	1.171	135.6	59.7	166.1	70.2